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**glum**

**QuantCo Inc.**

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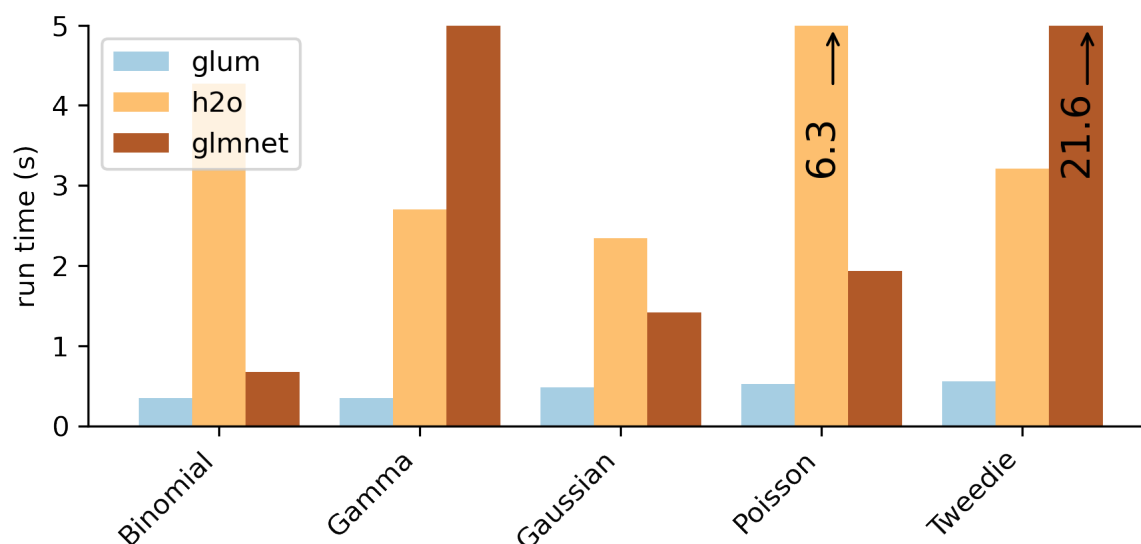
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glum is a fast, modern, Python-first GLM estimation library. Generalized linear modeling (GLM) is a core statistical tool that includes many common methods like least-squares regression, Poisson regression and logistic regression as special cases. In addition to fitting basic GLMs, glum supports a wide range of features. These include:

- Built-in cross validation for optimal regularization, efficiently exploiting a “regularization path”
- L1 and elastic net regularization, which produce sparse and easily interpretable solutions
- L2 regularization, including variable matrix-valued (Tikhonov) penalties, which are useful in modeling correlated effects
- Normal, Poisson, logistic, gamma, and Tweedie distributions, plus varied and customizable link functions
- Dispersion and standard errors
- Box and linear inequality constraints, sample weights, offsets.
- A scikit-learn-like API to fit smoothly into existing workflows.

glum was also built with performance in mind. The following figure shows the runtime of a realistic example using an insurance dataset. For more details and other benchmarks, see the [Benchmarks](#) section.



We suggest visiting the [Installation](#) and [Getting Started](#) sections first.



## INSTALLATION

You can install the package through conda:

```
conda install glum -c conda-forge
```

Head onwards to *Getting Started* to try it out!





## GETTING STARTED: FITTING A LASSO MODEL

The purpose of this tutorial is to show the basics of `glum`. It assumes a working knowledge of python, regularized linear models, and machine learning. The API is very similar to `scikit-learn`. After all, `glum` is based on a fork of `scikit-learn`.

If you have not done so already, please refer to our [installation instructions](#) for installing `glum`.

```
[1]: import pandas as pd
import sklearn
from sklearn.datasets import fetch_openml
from glum import GeneralizedLinearRegressor, GeneralizedLinearRegressorCV
```

### 2.1 Data

We start by loading the King County housing dataset from `openML` and splitting it into training and test sets. For simplicity, we don't go into any details regarding exploration or data cleaning.

```
[2]: house_data = fetch_openml(name="house_sales", version=3, as_frame=True)

# Use only select features
X = house_data.data[
    [
        "bedrooms",
        "bathrooms",
        "sqft_living",
        "floors",
        "waterfront",
        "view",
        "condition",
        "grade",
        "yr_built",
    ]
].copy()

# Targets
y = house_data.target
```

```
[3]: X_train, X_test, y_train, y_test = sklearn.model_selection.train_test_split(
    X, y, test_size = 0.3, random_state=5
)
```

## 2.2 GLM basics: fitting and predicting using the normal family

We'll use `glum.GeneralizedLinearRegressor` to predict the house prices using the available predictors.

We set three key parameters:

- **family**: the family parameter specifies the distributional assumption of the GLM and, as a consequence, the loss function to be minimized. Accepted strings are 'normal', 'poisson', 'gamma', 'inverse.gaussian', and 'binomial'. You can also pass in an instantiated `glum` distribution (e.g. `glum.TweedieDistribution(1.5)` )
- **alpha**: the constant multiplying the penalty term that determines regularization strength. (*Note*: `GeneralizedLinearRegressor` also has an alpha-search option. See the `GeneralizedLinearRegressorCV` example below for details on how alpha-search works).
- **l1\_ratio**: the elastic net mixing parameter ( $0 \leq \text{l1\_ratio} \leq 1$ ). For `l1_ratio = 0`, the penalty is the L2 penalty (ridge). For `l1_ratio = 1`, it is an L1 penalty (lasso). For  $0 < \text{l1\_ratio} < 1$ , the penalty is a combination of L1 and L2.

To be precise, we will be minimizing the function with respect to the parameters,  $\beta$ :

$$\frac{1}{N}(\mathbf{X}\beta - y)^2 + \alpha\|\beta\|_1 \quad (2.1)$$

```
[4]: glm = GeneralizedLinearRegressor(family="normal", alpha=0.1, l1_ratio=1)
```

The `GeneralizedLinearRegressor.fit()` method follows typical sklearn API style and accepts two primary inputs:

1. `X`: the design matrix with shape `(n_samples, n_features)`.
2. `y`: the `n_samples` length array of target data.

```
[5]: glm.fit(X_train, y_train)
```

```
[5]: GeneralizedLinearRegressor(alpha=0.1, l1_ratio=1)
```

Once the model has been estimated, we can retrieve useful information using an sklearn-style syntax.

```
[6]: # retrieve the coefficients and the intercept
      coefs = glm.coef_
      intercept = glm.intercept_
```

```
# use the model to predict on our test data
preds = glm.predict(X_test)

preds[0:5]
```

```
[6]: array([ 482648.22861066, 142902.68859995, 539452.61266391,
           569693.78048569, 1042446.90903451])
```

## 2.3 Fitting a GLM with cross validation

Now, we fit using automatic cross validation with `glum.GeneralizedLinearRegressorCV`. This mirrors the commonly used `cv.glmnet` function.

Some important parameters:

- `alphas`: for `GeneralizedLinearRegressorCV`, the best alpha will be found by searching along the regularization path. The regularization path is determined as follows:
  1. If `alpha` is an iterable, use it directly. All other parameters governing the regularization path are ignored.
  2. If `min_alpha` is set, create a path from `min_alpha` to the lowest alpha such that all coefficients are zero.
  3. If `min_alpha_ratio` is set, create a path where the ratio of `min_alpha` / `max_alpha` = `min_alpha_ratio`.
  4. If none of the above parameters are set, use a `min_alpha_ratio` of `1e-6`.
- `l1_ratio`: for `GeneralizedLinearRegressorCV`, if you pass `l1_ratio` as an array, the fit method will choose the best value of `l1_ratio` and store it as `self.l1_ratio_`.

```
[7]: glmcv = GeneralizedLinearRegressorCV(
    family="normal",
    alphas=None, # default
    min_alpha=None, # default
    min_alpha_ratio=None, # default
    l1_ratio=[0, 0.5, 1.0],
    fit_intercept=True,
    max_iter=150
)
glmcv.fit(X_train, y_train)
print(f"Chosen alpha: {glmcv.alpha_}")
print(f"Chosen l1 ratio: {glmcv.l1_ratio_}")
```

```
Chosen alpha:    0.0003274549162877732
Chosen l1 ratio: 0.0
```

Congratulations! You have finished our getting started tutorial. If you wish to learn more, please see our other tutorials for more advanced topics like Poisson, Gamma, and Tweedie regression, high dimensional fixed effects, and spatial smoothing using Tikhonov regularization.

```
[ ]:
```



## YET ANOTHER GLM PACKAGE?

`glum` was inspired by a desire to have a fast, maintainable, Python-first library for fitting GLMs with an extensive feature set.

At the beginning, we thoroughly examined all the existing contenders. The two mostly feature-complete options were `glmnet` and `H2O`. In many ways, the R package “`glmnet`” is the gold standard for regularized glm implementations. However, it is missing several useful features like built-in support for Tweedie and Gamma distributions. It also suffers from [impossible-to-maintain source](#) and thus has frequent bugs and segfaults. Although Python-to-`glmnet` interfaces exist, none is complete and well maintained. We also looked into the `H2O` implementation. It’s more feature-complete than `glmnet`, but there are serious integration issues with Python. As we discovered, there is also substantial room to improve performance beyond the level of `glmnet` or `H2O`.

So we decided to improve an existing package. Which one? To be a bit more precise, the question we wanted to answer was “Which library will be the least work to make feature-complete, high performance and correct?” To decide, we began by building a suite of benchmarks to compare the different libraries, and compared the libraries in terms of speed, the number of benchmarks that ran successfully, and code quality. In the end, we went with the code from [an sklearn pull request](#). We called it “sklearn-fork” and actually gave our code that name for quite a while too. `sklearn-fork` had decent, understandable code, converged in most situations, and included many of the features that we wanted. But it was slow. We figured it would be easier to speed up a functioning library than fix a broken but fast library. So we decided to start improving `sklearn-fork`. As a side note, a huge thank you to [Christian](#) for producing the baseline code for `glum`.

Ultimately, improving `sklearn-fork` seems to have been the right choice. We feel we have achieved our goals and `glum` is now *feature-complete*, *high-performance* and correct. However, over time, we uncovered more flaws in the optimizer than expected and, like most projects, building `sklearn-fork` into a feature-complete, fast, GLM library was a *harder task than we predicted*. When we started, `sklearn-fork` successfully converged for most problems. But, it was slow, taking hundreds or thousands of iteratively reweighted least squares (IRLS) iterations, many more than other similar libraries. Overall, the improvements we’ve made separate into three categories: algorithmic improvements, detailed software optimizations, and new features.

### 3.1 Algorithmic improvements

At the beginning, the lowest-hanging fruit came from debugging the implementation of IRLS and coordinate descent (CD) because those components were quite buggy and suboptimal. The algorithm we use is from [\[Yuan2012\]](#). We started by understanding the paper and relating it back to the code. This led to a circuitous chase around the code base, an effort that paid off when we noticed a hard-coded value in the optimizer was far too high. Fixing this was a one-line change that gave us 2-4X faster convergence!

Another large algorithmic improvement to the optimizer came from centering the predictor matrix to have mean zero. Coordinate descent cycles through one feature at a time, which is a strategy that works poorly with non-centered predictors because changing any coefficient changes the mean. In several cases, zero-centering reduced the total number of IRLS iterations by a factor of two, while leaving solutions unchanged. As we discuss below, centering is nontrivial

in the case of a sparse matrix because we don't want to modify the zero entries and destroy the sparsity. This was a major impetus for starting a tabular matrix handling library, [tabmat](#), as an extension of [glum](#).

Much later on, we made major improvements to the quality of the quadratic approximations for binomial, gamma, and Tweedie distributions, where the original Hessian approximations turned out to be suboptimal. For the first couple months, we took for granted that the quadratic log-likelihood approximations from [sklearn-fork](#) were correct. However, after substantial investigation, it turned out that we were using a Fisher information matrix-based approximation to the hessian rather than the true Hessian. This was done in [sklearn-fork](#) because the Fisher information matrix (FIM) is guaranteed to be positive definite for any link function or distribution, a necessary condition for guaranteed convergence. However, in cases where the true Hessian is also positive definite, using it will result in much faster convergence. It turned out that switching to using the true Hessian for these special cases (linear, Poisson, gamma, logistic regression and Tweedie regression for  $1 < p < 2$ ) gave huge reductions in the number of IRLS iterations. Some gamma regression problems dropped from taking 50-100 iterations to taking just 5-10 iterations.

Other important improvements:

- Using numerically stable log-likelihood, gradient and hessian formulas for the binomial distribution. In the naive version, we encounter floating point infinities for large parameter values in intermediate calculations.
- Exploring the use of an ADMM iterative L1 solver compared to our current CD solver. We ended up sticking with CD. This helped identify some crucial differences between [glum](#) and [H2O](#), which uses an ADMM solver.
- Active set iteration where we use heuristics to improve performance in L1-regularized problems by predicting, at the beginning of each iteration, which coefficients are likely to remain zero. This effectively reduces the set of predictors and significantly improves performance in severely L1-regularized problems.
- Making sure that we could correctly compare objective functions between libraries. The meaning of the regularization strength varies depending on the constant factors that multiply the log-likelihood.

## 3.2 Software optimizations

Substantial performance improvements came from many places.

- Removing redundant calculations and storing intermediate results to re-use later. The line search step had a particularly large number of such optimization opportunities.
- Cython-izing the coordinate descent implementation based on a version from [sklearn's](#) Lasso implementation. Several optimizations were possible even beyond the [sklearn](#) Lasso implementation and we hope to contribute some of these upstream.
- Hand-optimizing the formulas and Cython-izing the code for common distributions' log likelihood, gradients, and Hessians. We did this for normal, Poisson, gamma, Tweedie, binomial distributions.

The largest performance improvements have come from better tabular matrix handling. Initially, we were only handling uniformly dense or sparse matrices and using [numpy](#) and [scipy.sparse](#) to perform matrix operation. Now, we handle general "split" matrices that can be represented by a combination of dense, sparse, and categorical subcomponents. In addition, we built a `StandardizedMatrix` which handles the offsetting and multiplication needed to standardize a matrix to have mean zero and standard deviation one. We store the offsets and multipliers to perform this operation without modifying the underlying matrices.

We took our first step into developing custom matrix classes when we realized that even the pure dense and sparse matrix implementations were suboptimal. The default [scipy.sparse](#) matrix-multiply and matrix-vector product implementations are not parallel. Furthermore, many matrix-vector products only involve a small subset of rows or columns. As a result, we now have custom implementations of these operations that are parallelized and allow operating on a restricted set of rows and columns.

Before continuing, a quick summary of the only three matrix operations that we care about for GLM estimation:

- Matrix-vector products. `X.dot(v)` in [numpy](#) notation

- Transpose-matrix-vector products. `X.T.dot(v)`
- Sandwich products. `X.T @ diag(d) @ X`

As a matrix multiplication, the sandwich products are higher-dimensional operations than the matrix-vector products and, as such, are particularly expensive. Not only that, but the default implementation in `numpy` or `scipy.sparse` is going to be very inefficient. With dense `numpy` arrays, if we perform `X.T @ diag(d)`, that will allocate and create a whole new matrix that's just as large as the original `X` matrix. Then, we still need to perform a matrix multiply! As a result, we implemented a parallelized, cache-friendly, SIMD-optimized sandwich product operation that avoids the copy and performs the operation as a single matrix-multiply-like operation. We are in the process of contributing an implementation to the [BLIS library](#).

The next big matrix optimization came from realizing that most data matrices are neither fully dense nor fully sparse. Some columns will be very sparse (e.g. number of parrots owned), some columns will be one-hot encoded categoricals (e.g. preferred parrot species) while other columns will be dense (e.g. volume in liters of the most recently seen parrot). So we built a `SplitMatrix` class that splits a matrix into dense and sparse subcomponents. A threshold of around 90% sparsity seems to be about the level at which it is beneficial to use a simple CSR sparse matrix instead of a dense matrix. The benefit of this split matrix was large, improving performance across all the matrix operations by 2-5x.

Later on, we also added categorical matrix handling to the mix. Many categorical columns will be very sparse. If there are 100 evenly distributed categories, each column will have 99% sparse. However, simply treating them as a general sparse matrix is leaving a lot on the table. Beyond just being sparse, we know that every non-zero entry is a one and that every row has only a single non-zero column. This is particularly beneficial for sandwich products where the output ends up being diagonal. But, despite the clear gains, adding categorical matrices was quite a large undertaking. We needed to modify our data generation process to produce categoricals instead of one-hot-encoded columns, add and optimize each of our matrix operations for categoricals, and specify “sandwich” interactions between categorical matrices, sparse matrices, and dense matrices. The result was a large improvement in runtime, with some sandwich and matrix-transpose-dot operations sped up by more than an order of magnitude.

The end result of all these matrix optimizations is that we now have a fairly complete library for handling simple sandwich, dot and transpose-dot operations on a mix of dense, sparse and categorical matrices. This is perfect for most tabular data! So, we've split this component off into its own library, [tabmat](#).

### 3.3 New Features

In addition to the heavy focus on optimization and algorithmic correctness, we've also added a few important features to `glum` beyond what was already available in `sklearn-fork`.

- Automatic cross validation and regularization path handling similar in behavior to `glmnet`.
- Linear inequality constraints on coefficients.
- A step size convergence criterion in addition to the typical gradient-norm based criterion.
- The binomial distribution, and as a result, L1 and L2-regularized logistic regression.
- Standard errors.

## 3.4 References



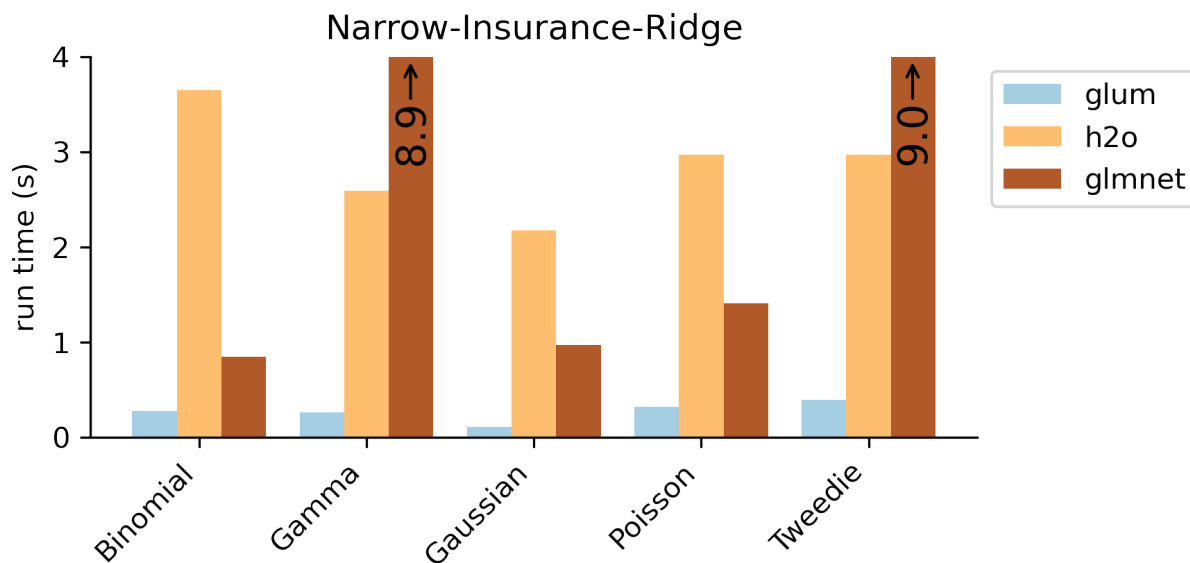
## BENCHMARKS AGAINST GLMNET AND H2O

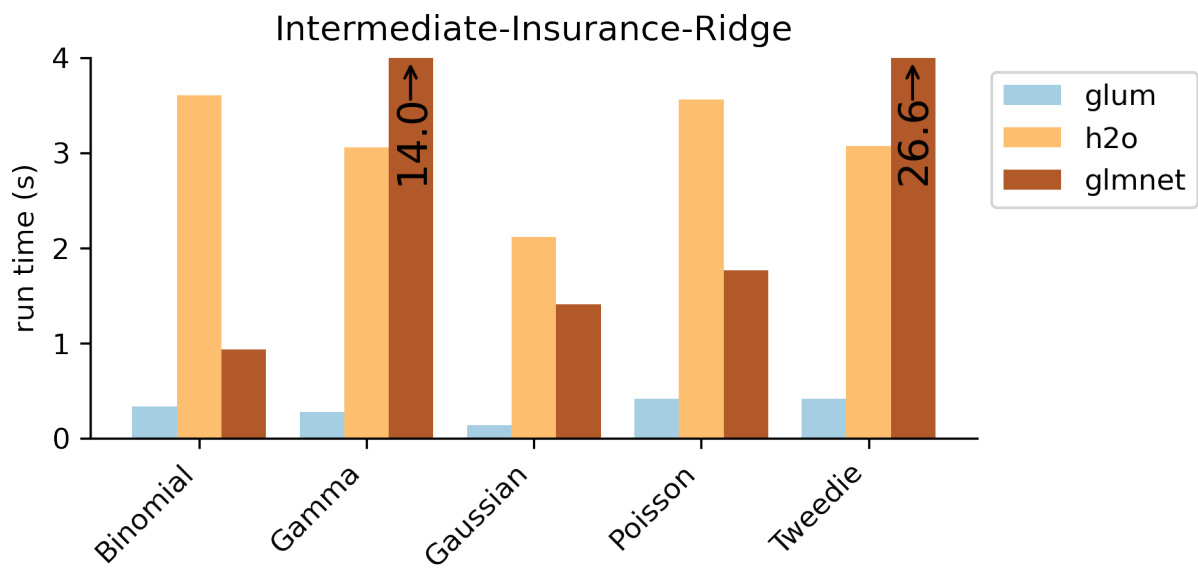
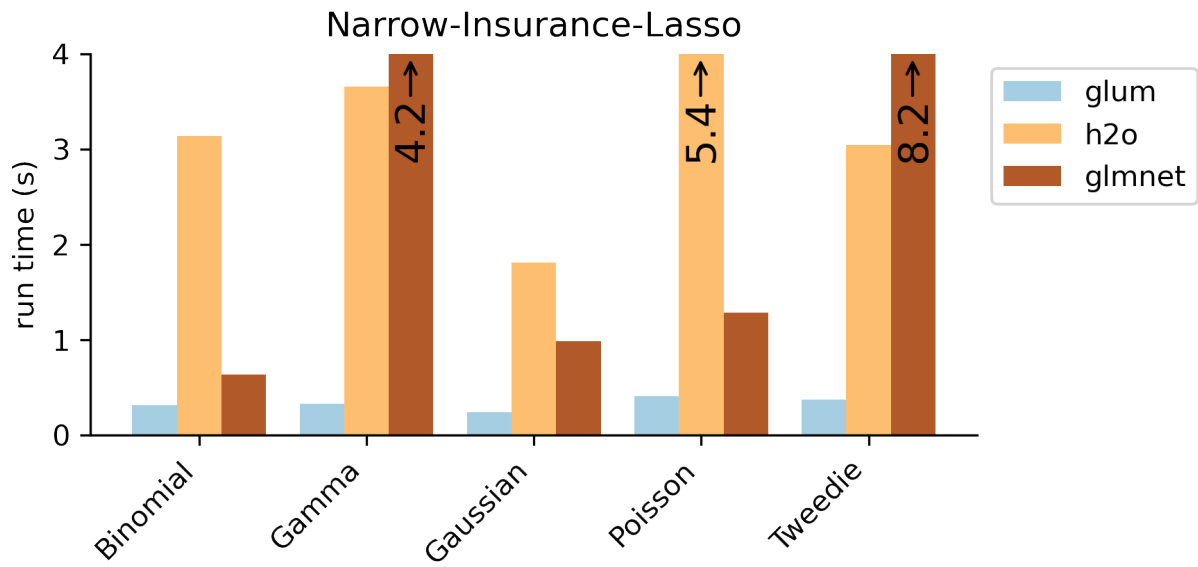
The following benchmarks were run on a MacBook Pro laptop with a quad-core Intel Core i5.

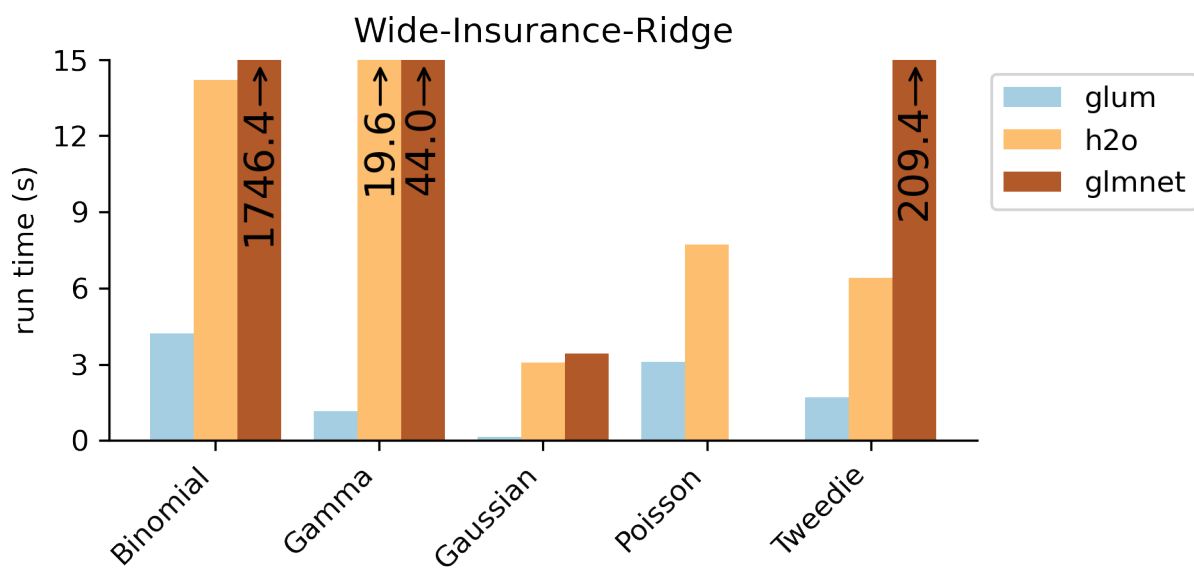
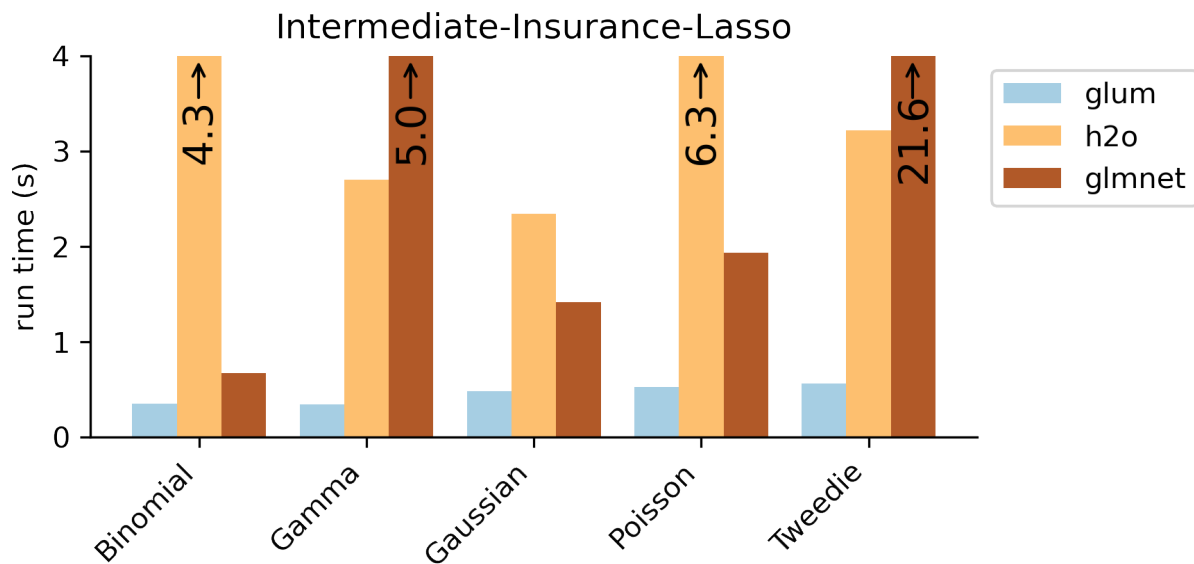
The title of each plot refers to both which dataset the benchmark was run on and whether a L2 ridge regression penalty or an L1 lasso penalty was included. For example “Narrow-Insurance-Ridge” was run on the `narrow-insurance` dataset with a ridge regression penalty. Each dataset/penalty pair is tested on five distributions that cover most of the common GLM types. The outcome variable is modified appropriately so that the behavior is similar to that expected for the distribution. For example, for the Poisson regression, we predict the number of claims per person. And for the binomial regression, we predict whether any given individual has ever had a claim. For the `housing` dataset, we only test three distributions because it does not contain count data that can be used as an outcome.

Note that `glum` was originally developed to solve problems where  $N \gg K$  (number of observations is larger than the number of predictors), which is the case for the following benchmarks.

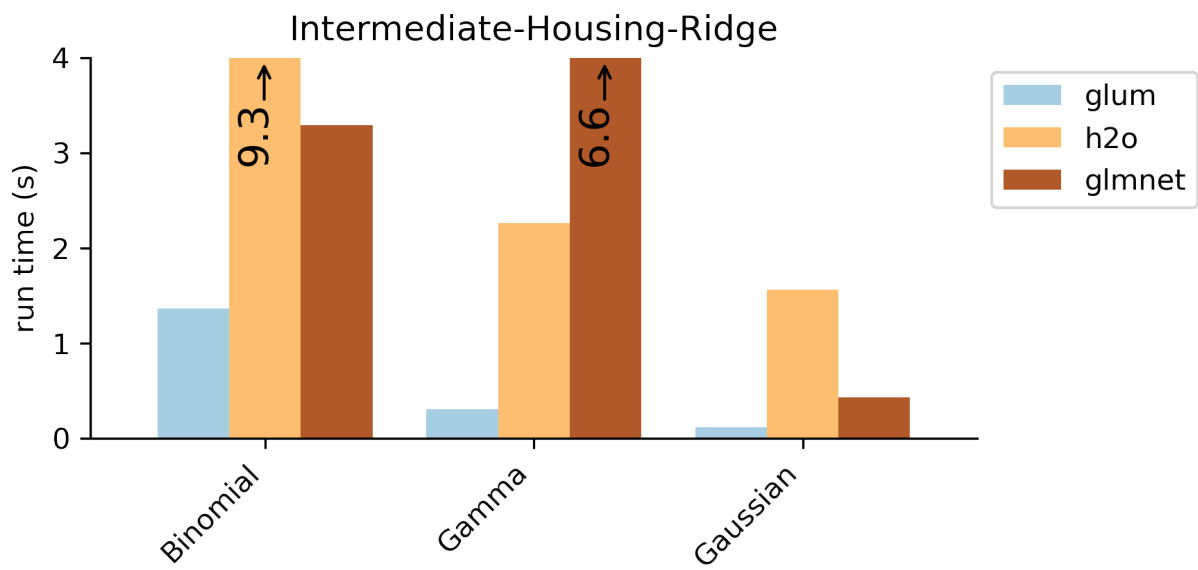
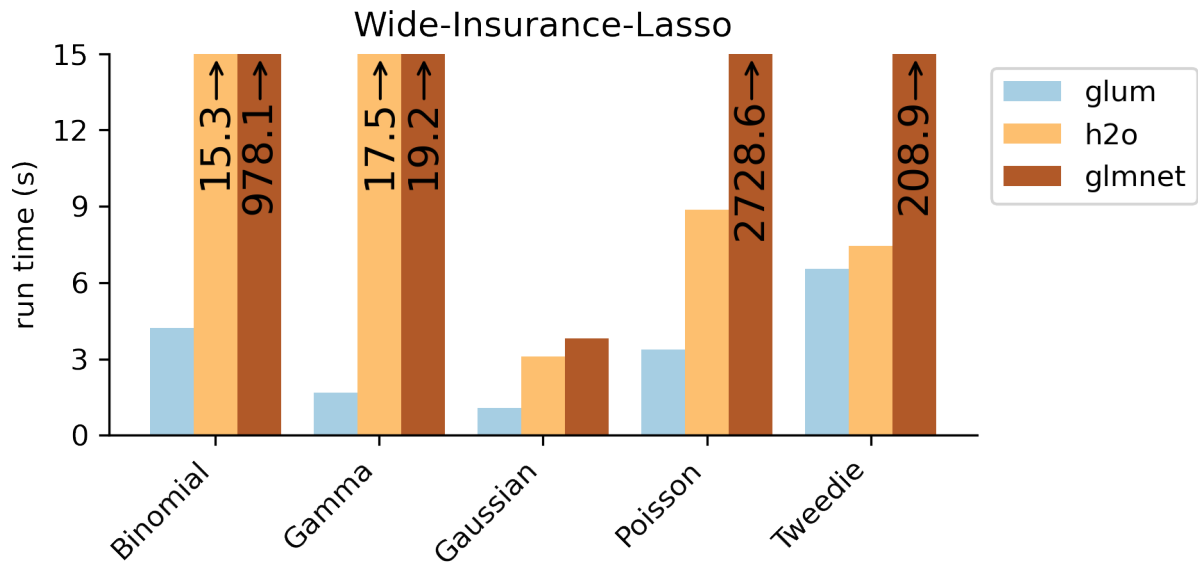
If a bar goes out of the range of the chart, the exact runtime is printed on the bar with an arrow indicating that the bar is truncated.

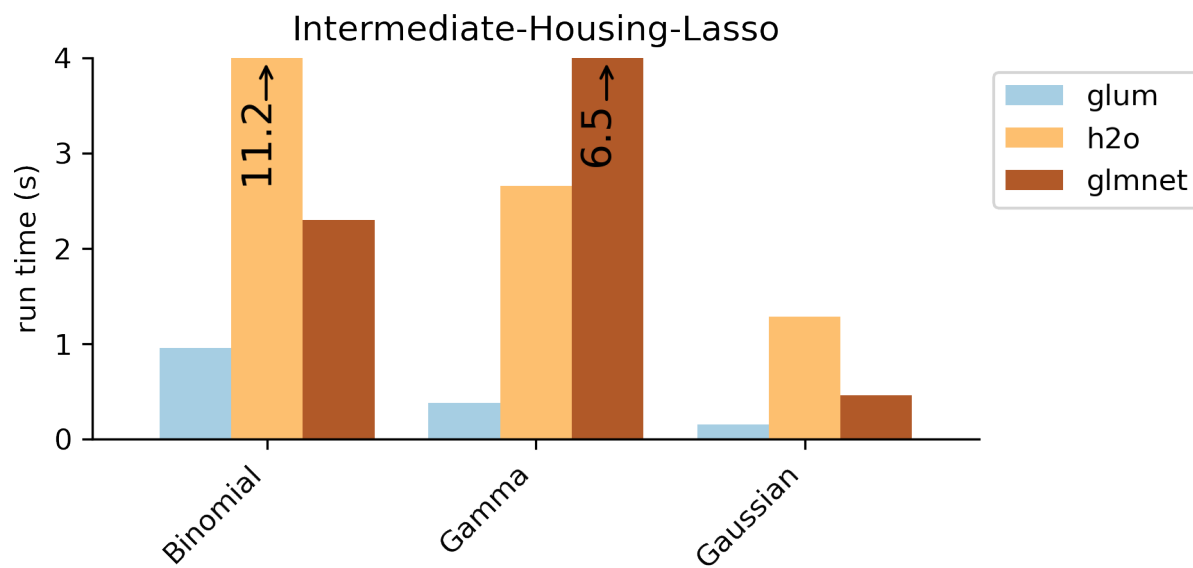






Note that the `r-glmnet` result for the wide-insurance-ridge Poisson benchmark is missing because `glmnet` did not converge after several hours of runtime.







## TUTORIALS

### 5.1 GLM Tutorial: Poisson, Gamma, and Tweedie with French Motor Third-Party Liability Claims

#### Intro

This tutorial shows why and how to use Poisson, Gamma, and Tweedie GLMs on an insurance claims dataset using `glum`. It was inspired by, and closely mirrors, two other GLM tutorials that used this dataset:

1. An sklearn-learn tutorial, [Tweedie regression on insurance claims](#), which was created for this (partially merged) [sklearn PR](#) that we based `glum` on
2. An R tutorial, [Case Study: French Motor Third-Party Liability Claims with R code](#).

#### Background

Insurance claims are requests made by a policy holder to an insurance company for compensation in the event of a covered loss. When modeling these claims, the goal is often to estimate, per policy, the total claim amount per exposure unit. (i.e. number of claims  $\times$  average amount per claim per year). This amount is also referred to as the pure premium.

Two approaches for modeling this value are:

1. Modeling the total claim amount per exposure directly
2. Modeling number of claims and claim amount separately with a frequency and a severity model

In this tutorial, we demonstrate both approaches. We start with the second option as it shows how to use two different families/distributions (Poisson and Gamma) within a GLM on a single dataset. We then show the first approach using a single poisson-gamma Tweedie regressor (i.e. a Tweedie with power  $p \in (1, 2)$ )

#### 5.1.1 Table of Contents

- 1. Load and Prepare Datasets from Openml.org
- 2. *Frequency GLM - Poisson Distribution*
- 3. *Severity GLM - Gamma Distribution*
- 4. *Combined GLM - Tweedie Distribution*

```
[1]: import matplotlib.pyplot as plt
import numpy as np
import pandas as pd
import scipy.optimize as optimize
```

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```

import scipy.stats
from dask_ml.preprocessing import Categorizer
from sklearn.metrics import mean_absolute_error
from sklearn.model_selection import ShuffleSplit
from glum import GeneralizedLinearRegressor
from glum import TweedieDistribution

from load_transform import load_transform

```

## 5.1.2 1. Load and prepare datasets from Openml

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First, we load in our *dataset from openML* and apply several transformations. In the interest of simplicity, we do not include the data loading and preparation code in this notebook. Below is a list of further resources if you wish to explore further:

1. If you want to run the same code yourself, please see the helper functions [here](#).
2. For a detailed description of the data, see [here](#).
3. For an excellent exploratory data analysis, see the case study paper linked above.

Some important notes about the dataset post-transformation:

- Total claim amounts are aggregated per policy
- For ClaimAmountCut, the claim amounts (pre-aggregation) were cut at 100,000 per single claim. We choose to use this amount rather than the raw ClaimAmount. (100,000 is the 0.9984 quantile but claims > 100,000 account for 25% of the overall claim amount)
- We aggregate the total claim amounts per policy
- ClaimNb is the total number of claims per policy with claim amount greater zero
- VehPower, VehAge, and DrivAge are clipped and/or digitized into bins so that they can be used as categoricals later on

```

[2]: df = load_transform()
with pd.option_context('display.max_rows', 10):
    display(df)

```

	ClaimNb	Exposure	Area	VehPower	VehAge	DrivAge	BonusMalus	\
IDpol								
1	0	0.10000	D	5	0	5	50	
3	0	0.77000	D	5	0	5	50	
5	0	0.75000	B	6	1	5	50	
10	0	0.09000	B	7	0	4	50	
11	0	0.84000	B	7	0	4	50	
...	...	...	...	...	...	...	...	
6114326	0	0.00274	E	4	0	5	50	
6114327	0	0.00274	E	4	0	4	95	
6114328	0	0.00274	D	6	1	4	50	
6114329	0	0.00274	B	4	0	5	50	
6114330	0	0.00274	B	7	1	2	54	

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IDpol	VehBrand	VehGas	Density	Region	ClaimAmount	ClaimAmountCut
1	B12	Regular	1217	R82	0.0	0.0
3	B12	Regular	1217	R82	0.0	0.0
5	B12	Diesel	54	R22	0.0	0.0
10	B12	Diesel	76	R72	0.0	0.0
11	B12	Diesel	76	R72	0.0	0.0
...	...	...	...	...	...	...
6114326	B12	Regular	3317	R93	0.0	0.0
6114327	B12	Regular	9850	R11	0.0	0.0
6114328	B12	Diesel	1323	R82	0.0	0.0
6114329	B12	Regular	95	R26	0.0	0.0
6114330	B12	Diesel	65	R72	0.0	0.0

[678013 rows x 13 columns]

## 5.1.3 2. Frequency GLM - Poisson distribution

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We start with the first part of our two part GLM - modeling the frequency of claims using a Poisson regression. Below, we give some background on why the Poisson family makes the most sense in this context.

### 2.1 Why Poisson distributions?

Poisson distributions are typically used to model the number of events occurring in a fixed period of time when the events occur independently at a constant rate. In our case, we can think of motor insurance claims as the events, and a unit of exposure (i.e. a year) as the fixed period of time.

To get more technical:

We define:

- $z$ : number of claims
- $w$ : exposure (time in years under risk)
- $y = \frac{z}{w}$ : claim frequency per year
- $X$ : feature matrix

The number of claims  $z$  is an integer,  $z \in [0, 1, 2, 3, \dots]$ . Theoretically, a policy could have an arbitrarily large number of claims—very unlikely but possible. The simplest distribution for this range is a Poisson distribution  $z \sim \text{Poisson}$ . However, instead of  $z$ , we will model the frequency  $y$ . Nonetheless, this is still (scaled) Poisson distributed with variance inverse proportional to  $w$ , cf. [wikipedia:Reproductive\\_EDM](#).

To verify our assumptions, we start by plotting the observed frequencies and a fitted Poisson distribution (Poisson regression with intercept only).

```
[3]: # plt.subplots(figsize=(10, 7))
df_plot = (
    df.loc[:, ['ClaimNb', 'Exposure']].groupby('ClaimNb').sum()
    .assign(Frequency_Observed = lambda x: x.Exposure / df['Exposure'].sum())
)
```

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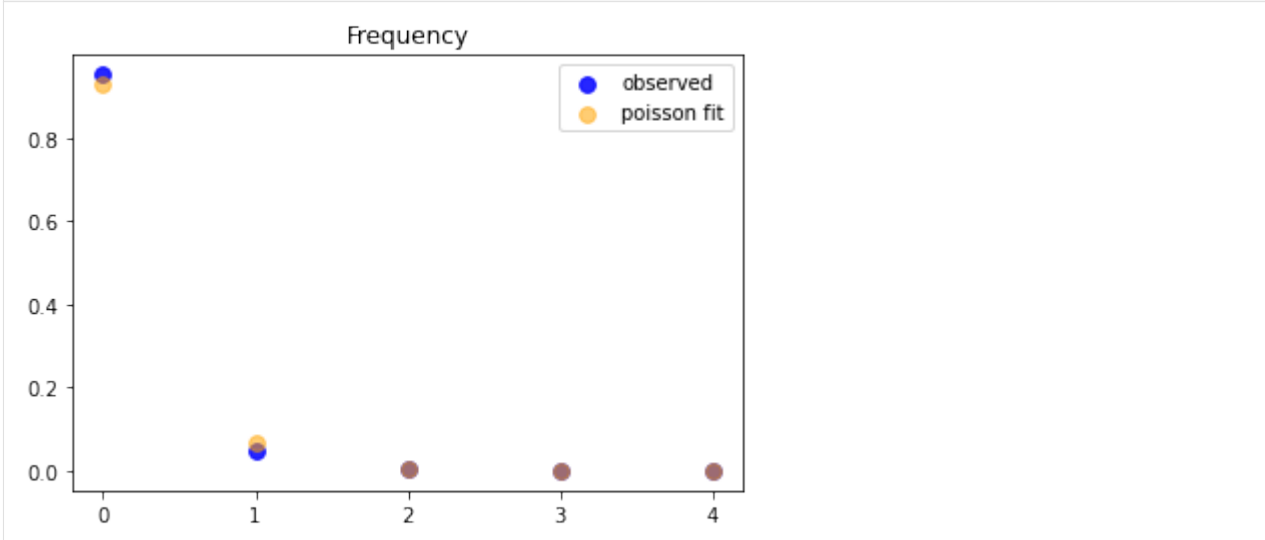
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```

mean = df['ClaimNb'].sum() / df['Exposure'].sum()

x = range(5)
plt.scatter(x, df_plot['Frequency_Observed'].values, color="blue", alpha=0.85, s=60, label='observed')
plt.scatter(x, scipy.stats.poisson.pmf(x, mean), color="orange", alpha=0.55, s=60, label='poisson fit')
plt.xticks(x)
plt.legend()
plt.title("Frequency");

```



This is a strong confirmation for the use of a Poisson when fitting!

## 2.2 Train and test frequency GLM

Now, we start fitting our model. We use claims frequency = claim number/exposure as our outcome variable. We then divide the dataset into training set and test set with a 9:1 random split.

Also, notice that we do not one hot encode our columns. Rather, we take advantage of `glum`'s integration with `tabmat`, which allows us to pass in categorical columns directly! `tabmat` will handle the encoding for us and even includes a handful of helpful matrix operation optimizations. We use the `Categorizer` from `dask_ml` to set our categorical columns as categorical dtypes and to ensure that the categories align in fitting and predicting.

```

[4]: z = df['ClaimNb'].values
weight = df['Exposure'].values
y = z / weight # claims frequency

ss = ShuffleSplit(n_splits=1, test_size=0.1, random_state=42)
train, test = next(ss.split(y))

categoricals = ["VehBrand", "VehGas", "Region", "Area", "DrivAge", "VehAge", "VehPower"]
predictors = categoricals + ["BonusMalus", "Density"]
glm_categorizer = Categorizer(columns=categoricals)

X_train_p = glm_categorizer.fit_transform(df[predictors].iloc[train])

```

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```
X_test_p = glm_categorizer.transform(df[predictors].iloc[test])
y_train_p, y_test_p = y[train], y[test]
w_train_p, w_test_p = weight[train], weight[test]
z_train_p, z_test_p = z[train], z[test]
```

Now, we define our GLM using the GeneralizedLinearRegressor class from glum.

- `family='poisson'`: creates a Poisson regressor
- `alpha_search=True`: tells the GLM to search along the regularization path for the best alpha
- `l1_ratio = 1` tells the GLM to only use l1 penalty (not l2). `l1_ratio` is the elastic net mixing parameter. For `l1_ratio = 0`, the penalty is an L2 penalty. For `l1_ratio = 1`, it is an L1 penalty. For  $0 < l1\_ratio < 1$ , the penalty is a combination of L1 and L2.

See the GeneralizedLinearRegressor class [API documentation](#) for more details.

*Note:* glum also supported a cross validation model GeneralizedLinearRegressorCV. However, because cross validation requires fitting many models, it is much slower and we don't demonstrate it in this tutorial.

```
[5]: f_glm1 = GeneralizedLinearRegressor(family='poisson', alpha_search=True, l1_ratio=1, fit_
    ↪ intercept=True)
```

```
f_glm1.fit(
    X_train_p,
    y_train_p,
    sample_weight=w_train_p
);

pd.DataFrame({'coefficient': np.concatenate([f_glm1.intercept_, f_glm1.coef_])},
    index=['intercept'] + f_glm1.feature_names_.T
```

```
[5]:
```

	intercept	VehBrand__B1	VehBrand__B10	VehBrand__B11	\	
coefficient	-4.269268	-0.003721	-0.010846	0.138466		
	VehBrand__B12	VehBrand__B13	VehBrand__B14	VehBrand__B2	\	
coefficient	-0.259298	0.0	-0.110712	-0.003604		
	VehBrand__B3	VehBrand__B4	...	VehAge__1	VehAge__2	\
coefficient	0.044075	0.0	...	0.045494	-0.139428	
	VehPower__4	VehPower__5	VehPower__6	VehPower__7	VehPower__8	\
coefficient	-0.070054	-0.028142	0.0	0.0	0.016531	
	VehPower__9	BonusMalus	Density			
coefficient	0.164711	0.026764	0.000004			

[1 rows x 60 columns]

To measure our model's test and train performance, we use the deviance function for the Poisson family. We can get the total deviance function directly from glum's distribution classes and divide it by the sum of our sample weight.

*Note:* a Poisson distribution is equivalent to a Tweedie distribution with power = 1.

```
[6]: PoissonDist = TweedieDistribution(1)
    print('training loss f_glm1: {}'.format(
```

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```

    PoissonDist.deviance(y_train_p, f_glm1.predict(X_train_p), sample_weight=w_train_p)/
    ↪np.sum(w_train_p)
))

print('test loss f_glm1: {}'.format(
    PoissonDist.deviance(y_test_p, f_glm1.predict(X_test_p), sample_weight=w_test_p)/
    ↪np.sum(w_test_p)))

training loss f_glm1: 0.45704947333555146
test loss f_glm1: 0.45793061314157685

```

A GLM with canonical link function (Normal - identity, Poisson - log, Gamma -  $1/x$ , Binomial - logit) with an intercept term has the so called **balance property**. Neglecting small deviations due to an imperfect fit, on the training sample the results satisfy the equality:

$$\sum_{i \in \text{training}} w_i y_i = \sum_{i \in \text{training}} w_i \hat{\mu}_i$$

As expected, this property holds in our real data:

```

[7]: # balance property of GLM with canonical link, like log-link for Poisson:
z_train_p.sum(), (f_glm1.predict(X_train_p) * w_train_p).sum()

[7]: (23785, 23785.198509368805)

```

### 5.1.4 3. Severity GLM - Gamma distribution

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Now, we fit a GLM for the severity with the same features as the frequency model. The severity  $y$  is the average claim size. We define:

- $z$ : total claim amount, single claims cut at 100,000
- $w$ : number of claims (with positive claim amount!)
- $y = \frac{z}{w}$ : severity

#### 3.1 Why Gamma distributions

The severity  $y$  is a positive, real number,  $y \in (0, \infty)$ . Theoretically, especially for liability claims, one could have arbitrary large numbers—very unlikely but possible. A very simple distribution for this range is an Exponential distribution, or its generalization, a Gamma distribution  $y \sim \text{Gamma}$ . In the insurance industry, it is well known that the severity might be skewed by a few very large losses. It's common to model these tail losses separately so here we cut out claims larger than 100,000 to focus on modeling small and moderate claims.

```

[8]: df_plot = (
    df.loc[:, ['ClaimAmountCut', 'ClaimNb']]
    .query('ClaimNb > 0')
    .assign(Severity_Observed = lambda x: x['ClaimAmountCut'] / df['ClaimNb'])
)

df_plot['Severity_Observed'].plot.hist(bins=400, density=True, label='Observed', )

```

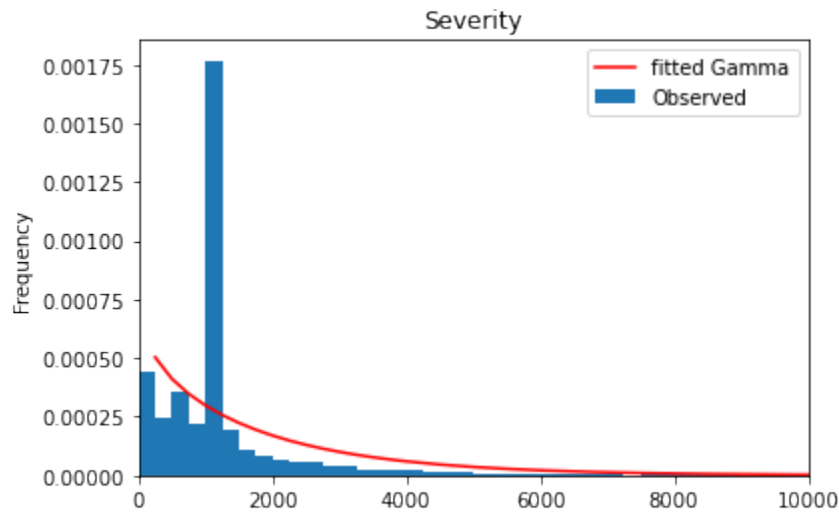
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```

x = np.linspace(0, 1e5, num=400)
plt.plot(x,
         scipy.stats.gamma.pdf(x, *scipy.stats.gamma.fit(df_plot['Severity_Observed'],
         floc=0)),
         'r-', label='fitted Gamma')
plt.legend()
plt.title("Severity");
plt.xlim(left=0, right = 1e4);
#plt.xticks(x);

```



```

[9]: # Check mean-variance relationship for Gamma:  $Var[Y] = E[Y]^2 / Exposure$ 
      # Estimate  $Var[Y]$  and  $E[Y]$ 
      # Plot estimates  $Var[Y]$  vs  $E[Y]^s/Exposure$ 
      # Note: We group by VehPower and BonusMalus in order to have different  $E[Y]$ .

def my_agg(x):
    """See https://stackoverflow.com/q/44635626"""
    x_sev = x['Sev']
    x_cnb = x['ClaimNb']
    n = x_sev.shape[0]
    names = {
        'Sev_mean': np.average(x_sev, sample_weight=x_cnb),
        'Sev_var': 1/(n-1) * np.sum((x_cnb/np.sum(x_cnb)) * (x_sev-np.average(x_sev,
        fsample_weight=x_cnb))**2),
        'ClaimNb_sum': x_cnb.sum()
    }
    return pd.Series(names, index=['Sev_mean', 'Sev_var', 'ClaimNb_sum'])

for col in ['VehPower', 'BonusMalus']:
    claims = df.groupby(col)['ClaimNb'].sum()
    df_plot = (df.loc[df[col].isin(claims[claims >= 4].index), :]
               .query('ClaimNb > 0')
               .assign(Sev = lambda x: x['ClaimAmountCut']/x['ClaimNb'])
               .groupby(col)

```

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```

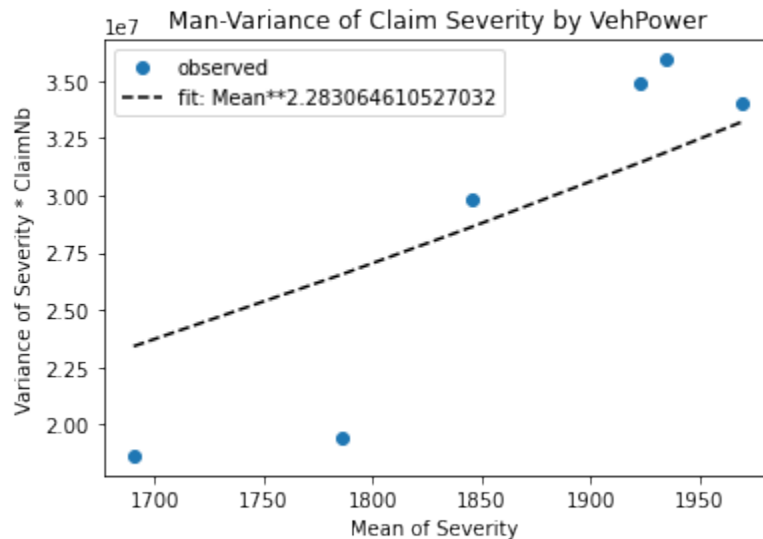
        .apply(my_agg)
    )

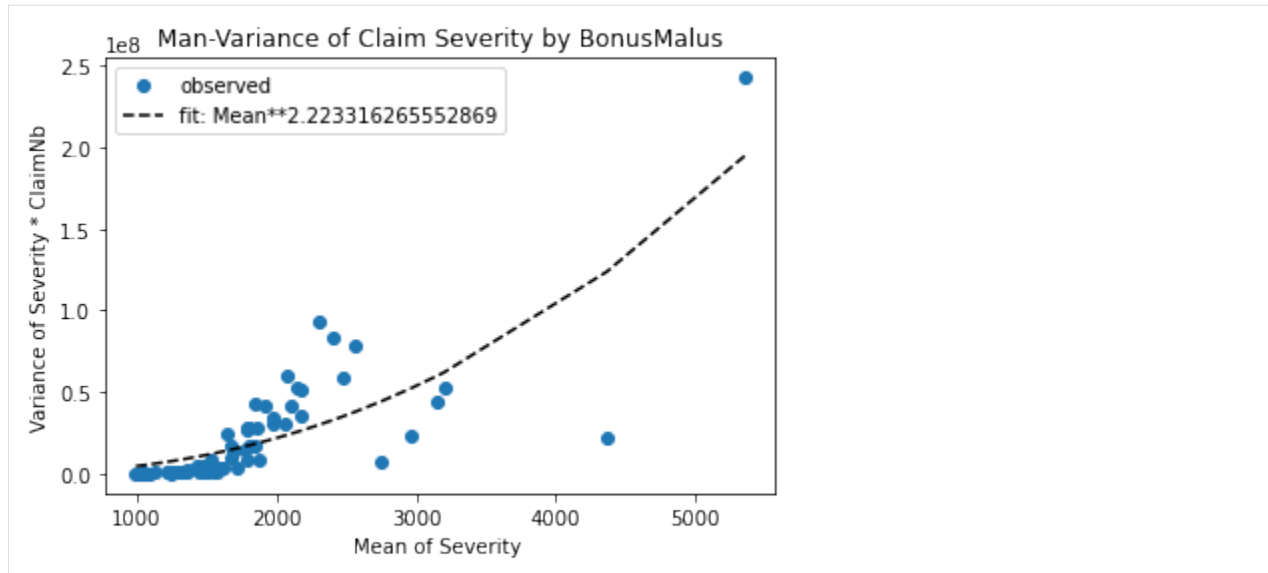
plt.plot(df_plot['Sev_mean'], df_plot['Sev_var'] * df_plot['ClaimNb_sum'], '.',
         markersize=12, label='observed')

# fit: mean**p/claims
p = optimize.curve_fit(lambda x, p: np.power(x, p),
                       df_plot['Sev_mean'].values,
                       df_plot['Sev_var'] * df_plot['ClaimNb_sum'],
                       p0 = [2])[0][0]
df_fit = pd.DataFrame({'x': df_plot['Sev_mean'],
                       'y': np.power(df_plot['Sev_mean'], p)})
df_fit = df_fit.sort_values('x')

plt.plot(df_fit.x, df_fit.y,
         'k--', label='fit: Mean**{}'.format(p))
plt.xlabel('Mean of Severity ')
plt.ylabel('Variance of Severity * ClaimNb')
plt.legend()
plt.title('Man-Variance of Claim Severity by {}'.format(col))
plt.show()

```





Great! A Gamma distribution seems to be an empirically reasonable assumption for this dataset.

*Hint:* If  $Y$  were normal distributed, one should see a horizontal line, because  $Var[Y] = constant/Exposure$  and the fit should give  $p \approx 0$ .

### 3.2 Severity GLM with train and test data

We fit a GLM for the severity with the same features as the frequency model. We use the same categorizer as before.

*Note:*

- We filter out  $ClaimAmount == 0$ . The severity problem is to model claim amounts conditional on a claim having already been submitted. It seems reasonable to treat a claim of zero as equivalent to no claim at all. Additionally, zero is not included in the open interval  $(0, \infty)$  support of the Gamma distribution.
- We use  $ClaimNb$  as sample weights.
- We use the same split in train and test data such that we can predict the final claim amount on the test set as the product of our Poisson claim number and Gamma claim severity GLMs.

```
[10]: idx = df['ClaimAmountCut'].values > 0

z = df['ClaimAmountCut'].values
weight = df['ClaimNb'].values
# y = claims severity
y = np.zeros_like(z) # zeros will never be used
y[idx] = z[idx] / weight[idx]

# we also need to represent train and test as boolean indices
itrain = np.zeros(y.shape, dtype='bool')
itest = np.zeros(y.shape, dtype='bool')
itrain[train] = True
itest[test] = True
# simplify life
itrain = idx & itrain
itest = idx & itest
```

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```

X_train_g = glm_categorizer.fit_transform(df[predictors].iloc[itrain])
X_test_g = glm_categorizer.transform(df[predictors].iloc[itest])
y_train_g, y_test_g = y[itrain], y[itest]
w_train_g, w_test_g = weight[itrain], weight[itest]
z_train_g, z_test_g = z[itrain], z[itest]

```

We fit our model with the same parameters before, but of course, this time we use `family=gamma`.

```

[11]: s_glm1 = GeneralizedLinearRegressor(family='gamma', alpha_search=True, l1_ratio=1, fit_
      ↪ intercept=True)
      s_glm1.fit(X_train_g, y_train_g, sample_weight=weight[itrain])

      pd.DataFrame({'coefficient': np.concatenate([s_glm1.intercept_, s_glm1.coef_])},
                    index=['intercept'] + s_glm1.feature_names_).T
[11]:
      intercept  VehBrand__B1  VehBrand__B10  VehBrand__B11  \
coefficient    7.3389      -0.034591      0.040528      0.13116

      VehBrand__B12  VehBrand__B13  VehBrand__B14  VehBrand__B2  \
coefficient    0.035838      0.100753     -0.073995     -0.033196

      VehBrand__B3  VehBrand__B4  ...  VehAge__1  VehAge__2  \
coefficient      0.0      0.049078  ...      0.0     -0.024827

      VehPower__4  VehPower__5  VehPower__6  VehPower__7  VehPower__8  \
coefficient    -0.009537     -0.089972      0.071376      0.009361     -0.042491

      VehPower__9  BonusMalus  Density
coefficient    0.051636      0.002365 -0.000001

[1 rows x 60 columns]

```

Again, we measure performance with the deviance of the distribution. We also compare against the simple arithmetic mean and include the mean absolute error to help understand the actual scale of our results.

*Note:* a Gamma distribution is equivalent to a Tweedie distribution with power = 2.

```

[26]: GammaDist = TweedieDistribution(2)
      print('training loss (deviance) s_glm1: {}'.format(
            GammaDist.deviance(
                y_train_g, s_glm1.predict(X_train_g), sample_weight=w_train_g
            )/np.sum(w_train_g)
        ))
      print('training mean absolute error s_glm1: {}'.format(
            mean_absolute_error(y_train_g, s_glm1.predict(X_train_g))
        ))

      print('\ntesting loss s_glm1 (deviance): {}'.format(
            GammaDist.deviance(
                y_test_g, s_glm1.predict(X_test_g), sample_weight=w_test_g
            )/np.sum(w_test_g)
        ))

```

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```

print('testing mean absolute error s_glm1: {}'.format(
    mean_absolute_error(y_test_g, s_glm1.predict(X_test_g))
))

print('\ntesting loss Mean (deviance):      {}'.format(
    GammaDist.deviance(
        y_test_g, np.average(z_train_g, sample_weight=w_train_g)*np.ones_like(z_test_g),
        ↪sample_weight=w_test_g
    )/np.sum(w_test_g)
))
print('testing mean absolute error Mean:    {}'.format(
    mean_absolute_error(y_test_g, np.average(z_train_g, sample_weight=w_train_g)*np.ones_
    ↪like(z_test_g))
))

```

```

training loss (deviance) s_glm1:      1.29010461534461
training mean absolute error s_glm1: 1566.1785138646032

testing loss s_glm1 (deviance):      1.2975718597070154
testing mean absolute error s_glm1: 1504.4458958597086

testing loss Mean (deviance):      1.3115309309577132
testing mean absolute error Mean:   1689.205530922944

```

Even though the deviance improvement seems small, the improvement in mean absolute error is not! (In the insurance world, this will make a significant difference when aggregated over all claims).

### 3.3 Combined frequency and severity results

We put together the prediction of frequency and severity to get the predictions of the total claim amount per policy.

```

[13]: #Put together freq * sev together
print("Total claim amount on train set, observed = {}, predicted = {}".
    format(df['ClaimAmountCut'].values[train].sum(),
        np.sum(df['Exposure'].values[train] * f_glm1.predict(X_train_p) * s_glm1.
        ↪predict(X_train_p)))
    )

print("Total claim amount on test set, observed = {}, predicted = {}".
    format(df['ClaimAmountCut'].values[test].sum(),
        np.sum(df['Exposure'].values[test] * f_glm1.predict(X_test_p) * s_glm1.
        ↪predict(X_test_p)))
    )

```

```

Total claim amount on train set, observed = 44594644.68, predicted = 44549152.42247057
Total claim amount on test set, observed = 4707551.37, predicted = 4946960.354743531

```

## 5.1.5 4. Combined GLM - Tweedie distribution

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Finally, to demonstrate an alternate approach to the combined frequency severity model, we show how we can model pure premium directly using a Tweedie regressor. Any Tweedie distribution with power  $p \in (1, 2)$  is known as [compound Poisson Gamma distribution](#)

```
[14]: weight = df['Exposure'].values
df["PurePremium"] = df["ClaimAmountCut"] / df["Exposure"]
y = df["PurePremium"]

X_train_t = glm_categorizer.fit_transform(df[predictors].iloc[train])
X_test_t = glm_categorizer.transform(df[predictors].iloc[test])
y_train_t, y_test_t = y.iloc[train], y.iloc[test]
w_train_t, w_test_t = weight[train], weight[test]
```

For now, we just arbitrarily select 1.5 as the power parameter for our Tweedie model. However for a better fit we could include the power parameter in the optimization/fitting process, possibly via a simple grid search.

*Note:* notice how we pass a TweedieDistribution object in directly for the family parameter. While glum supports strings for common families, it is also possible to pass in a glum distribution directly.

```
[15]: TweedieDist = TweedieDistribution(1.5)
t_glm1 = GeneralizedLinearRegressor(family=TweedieDist, alpha_search=True, l1_ratio=1,
    ↪ fit_intercept=True)
t_glm1.fit(X_train_t, y_train_t, sample_weight=w_train_t)

pd.DataFrame({'coefficient': np.concatenate([t_glm1.intercept_, t_glm1.coef_])},
    index=['intercept'] + t_glm1.feature_names_).T
```

```
[15]:
```

coefficient	intercept	VehBrand__B1	VehBrand__B10	VehBrand__B11	\
	2.88667	-0.064157	0.0	0.231868	
coefficient	VehBrand__B12	VehBrand__B13	VehBrand__B14	VehBrand__B2	\
	-0.211061	0.054979	-0.270346	-0.071453	
coefficient	VehBrand__B3	VehBrand__B4	...	VehAge__1	VehAge__2 \
	0.00291	0.059324	...	0.008117	-0.229906
coefficient	VehPower__4	VehPower__5	VehPower__6	VehPower__7	VehPower__8 \
	-0.111796	-0.123388	0.060757	0.005179	-0.021832
coefficient	VehPower__9	BonusMalus	Density		
	0.208158	0.032508	0.000002		

[1 rows x 60 columns]

Again, we use the distribution's deviance to measure model performance

```
[16]: print('training loss s_glm1: {}'.format(
    TweedieDist.deviance(y_train_t, t_glm1.predict(X_train_t), sample_weight=w_train_t)/
    ↪ np.sum(w_train_t)))
```

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```
print('testing loss s_glm1: {}'.format(
    TweedieDist.deviance(y_test_t, t_glm1.predict(X_test_t), sample_weight=w_test_t)/np.
    ↳sum(w_test_t)))
```

training loss s\_glm1: 73.91371104577475  
testing loss s\_glm1: 72.35318912371723

Finally, we again show the total predicted vs. true claim amount on the training and test set

```
[17]: #Put together freq * sev together
print("Total claim amount on train set, observed = {}, predicted = {}".
      format(df['ClaimAmountCut'].values[train].sum(),
             np.sum(df['Exposure'].values[train] * t_glm1.predict(X_train_p)))
      )

print("Total claim amount on test set, observed = {}, predicted = {}".
      format(df['ClaimAmountCut'].values[test].sum(),
             np.sum(df['Exposure'].values[test] * t_glm1.predict(X_test_p)))
      )
```

Total claim amount on train set, observed = 44594644.68, predicted = 45027861.66007367  
Total claim amount on test set, observed = 4707551.37, predicted = 4999381.03386664

In terms of the combined proximity to the true total claim amounts, the frequency severity model performed a bit better than Tweedie model. However, both approaches ultimately prove to be effective.

## 5.2 High Dimensional Fixed Effects with Rossman Sales Data

### Intro

This tutorial demonstrates how to create models with high dimensional fixed effects using glum. Using tabmat, we can pass categorical variables with a large range of values. glum and tabmat will handle the creation of the one-hot-encoded design matrix.

In some real-world problems, we have used millions of categories. This would be impossible with a dense matrix. General-purpose sparse matrices like compressed sparse row (CSR) matrices help but still leave a lot on the table. For a categorical matrix, we know that each row has only a single non-zero value and that value is 1. These optimizations are implemented in tabmat.CategoricalMatrix.

### Background

For this tutorial, we will be predicting sales for the European drug store chain Rossman. Specifically, we are tasked with predicting daily sales for future dates. Ideally, we want a model that can capture the many factors that influence stores sales – promotions, competition, school, holidays, seasonality, etc. As a baseline, we will start with a simple model that only uses a few basic predictors. Then, we will fit a model with a large number of fixed effects. For both models, we will use OLS with L2 regularization.

We will use a gamma distribution for our model. This choice is motivated by two main factors. First, our target variable, sales, is a positive real number, which matches the support of the gamma distribution. Second, it is expected that factors influencing sales are multiplicative rather than additive, which is better captured with a gamma regression than say, OLS.

*Note:* a few parts of this tutorial utilize local helper functions outside this notebook. If you wish to run the notebook on your own, you can find the rest of the code [here](#).

## 5.2.1 Table of Contents

- 1. Data Loading and Feature Engineering
- 2. Fit Baseline GLM
- 3. GLM with High Dimensional Fixed Effects
- 4. Plot Results

```
[1]: import os
from pathlib import Path

import altair as alt
import matplotlib.pyplot as plt
import numpy as np
import pandas as pd

from dask_ml.impute import SimpleImputer
from dask_ml.preprocessing import Categorizer
from glum import GeneralizedLinearRegressor
from sklearn.pipeline import Pipeline

from feature_engineering import apply_all_transformations
from process_data import load_test, load_train, process_data

import sys
sys.path.append("../")
from metrics import root_mean_squared_percentage_error

pd.set_option("display.float_format", lambda x: "%.3f" % x)
pd.set_option('display.max_columns', None)
alt.data_transformers.enable("json") # to allow for large plots

[1]: DataTransformerRegistry.enable('json')
```

## 5.2.2 1. Data loading and feature engineering

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We start by loading in the raw data. If you have not yet processed the raw data, it will be done below. (Initial processing consists of some basic cleaning and renaming of columns.

*Note:* if you wish to run this notebook on your own, and have not done so already, please download the data from the [Rossman Kaggle Challenge](#). This tutorial expects that it is in a folder named “raw\_data” under the same directory as the notebook.

## 1.1 Load

```
[2]: if not all(Path(p).exists() for p in ["raw_data/train.csv", "raw_data/test.csv", "raw_
      ↪data/store.csv"]):
      raise Exception("Please download raw data into 'raw_data' folder")

if not all(Path(p).exists() for p in ["processed_data/train.parquet", "processed_data/
      ↪test.parquet"]):
    "Processed data not found. Processing data from raw data..."
    process_data()
    "Done"

df = load_train().sort_values(["store", "date"])
df = df.iloc[:int(.1*len(df))]
df.head()
```

```
[2]:
```

	store	day_of_week	date	sales	customers	open	promo	\
1016095	1	2	2013-01-01	0	0	False	0	
1014980	1	3	2013-01-02	5530	668	True	0	
1013865	1	4	2013-01-03	4327	578	True	0	
1012750	1	5	2013-01-04	4486	619	True	0	
1011635	1	6	2013-01-05	4997	635	True	0	

	state_holiday	school_holiday	year	month	store_type	assortment	\
1016095	a	1	2013	1	c	a	
1014980	0	1	2013	1	c	a	
1013865	0	1	2013	1	c	a	
1012750	0	1	2013	1	c	a	
1011635	0	1	2013	1	c	a	

	competition_distance	competition_open_since_month	\
1016095	1270.000	9.000	
1014980	1270.000	9.000	
1013865	1270.000	9.000	
1012750	1270.000	9.000	
1011635	1270.000	9.000	

	competition_open_since_year	promo2	promo2_since_week	\
1016095	2008.000	0	NaN	
1014980	2008.000	0	NaN	
1013865	2008.000	0	NaN	
1012750	2008.000	0	NaN	
1011635	2008.000	0	NaN	

	promo2_since_year	promo_interval
1016095	NaN	None
1014980	NaN	None
1013865	NaN	None
1012750	NaN	None
1011635	NaN	None

## 1.2 Feature engineering

As mentioned earlier, we want our model to incorporate many factors that could influence store sales. We create a number of fixed effects to capture this information. These include fixed effects for:

- A certain number days before a school or state holiday
- A certain number days after a school or state holiday
- A certain number days before a promo
- A certain number days after a promo
- A certain number days before the store is open or closed
- A certain number days after the store is open or closed
- Each month for each store
- Each year for each store
- Each day of the week for each store

We also do several other transformations like computing the z score to eliminate outliers (in the next step)

```
[3]: df = apply_all_transformations(df)
df.head()
```

```
[3]:
```

	store	day_of_week	date	sales	customers	open	promo	\
1016095	1	2	2013-01-01	0	0	False	0	
1014980	1	3	2013-01-02	5530	668	True	0	
1013865	1	4	2013-01-03	4327	578	True	0	
1012750	1	5	2013-01-04	4486	619	True	0	
1011635	1	6	2013-01-05	4997	635	True	0	

	state_holiday	school_holiday	year	month	store_type	assortment	\
1016095	a	1	2013	1	c	a	
1014980	0	1	2013	1	c	a	
1013865	0	1	2013	1	c	a	
1012750	0	1	2013	1	c	a	
1011635	0	1	2013	1	c	a	

	competition_distance	competition_open_since_month	\
1016095	1270.000	9.000	
1014980	1270.000	9.000	
1013865	1270.000	9.000	
1012750	1270.000	9.000	
1011635	1270.000	9.000	

	competition_open_since_year	promo2	promo2_since_week	\
1016095	2008.000	0	NaN	
1014980	2008.000	0	NaN	
1013865	2008.000	0	NaN	
1012750	2008.000	0	NaN	
1011635	2008.000	0	NaN	

	promo2_since_year	promo_interval	age_quantile	competition_open	\
1016095	NaN	None	-1	1.000	
1014980	NaN	None	-1	1.000	

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1013865		NaN		None	-1	1.000	
1012750		NaN		None	-1	1.000	
1011635		NaN		None	-1	1.000	
	count	open_lag_1	open_lag_2	open_lag_3	open_lead_1	open_lead_2	\
1016095	0	1.0	1.0	1.0	True	True	
1014980	1	False	1.0	1.0	True	True	
1013865	2	True	False	1.0	True	True	
1012750	3	True	True	False	True	False	
1011635	4	True	True	True	False	True	
	open_lead_3	promo_lag_1	promo_lag_2	promo_lag_3	promo_lead_1		\
1016095	True	0.000	0.000	0.000	0.000		
1014980	True	0.000	0.000	0.000	0.000		
1013865	False	0.000	0.000	0.000	0.000		
1012750	True	0.000	0.000	0.000	0.000		
1011635	True	0.000	0.000	0.000	0.000		
	promo_lead_2	promo_lead_3	school_holiday_lag_1				\
1016095	0.000	0.000		0.000			
1014980	0.000	0.000		1.000			
1013865	0.000	0.000		1.000			
1012750	0.000	1.000		1.000			
1011635	1.000	1.000		1.000			
	school_holiday_lag_2	school_holiday_lag_3	school_holiday_lead_1				\
1016095	0.000		0.000		1.000		
1014980	0.000		0.000		1.000		
1013865	1.000		0.000		1.000		
1012750	1.000		1.000		1.000		
1011635	1.000		1.000		1.000		
	school_holiday_lead_2	school_holiday_lead_3	state_holiday_lag_1				\
1016095	1.000		1.000		0		
1014980	1.000		1.000		a		
1013865	1.000		1.000		0		
1012750	1.000		1.000		0		
1011635	1.000		1.000		0		
	state_holiday_lag_2	state_holiday_lag_3	state_holiday_lead_1				\
1016095	0		0		0		
1014980	0		0		0		
1013865	a		0		0		
1012750	0		a		0		
1011635	0		0		0		
	state_holiday_lead_2	state_holiday_lead_3	store_day_of_week				\
1016095	0		0		1_2		
1014980	0		0		1_3		
1013865	0		0		1_4		
1012750	0		0		1_5		
1011635	0		0		1_6		

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	store_month	store_school_holiday	store_state_holiday	store_year	\
1016095	1_1	1_1	1_True	1_2013	
1014980	1_1	1_1	1_False	1_2013	
1013865	1_1	1_1	1_False	1_2013	
1012750	1_1	1_1	1_False	1_2013	
1011635	1_1	1_1	1_False	1_2013	

	zscore
1016095	NaN
1014980	NaN
1013865	NaN
1012750	NaN
1011635	NaN

### 1.3 Train vs. validation selection

Lastly, we split our data into training and validation sets. Kaggle provides a test set for the Rossman challenge, but it does not directly include outcome data (sales), so we do not use it for our tutorial. Instead, we simulate predicting future sales by taking the last 5 months of our training data as our validation set.

```
[4]: validation_window = [pd.to_datetime("2015-03-15"), pd.to_datetime("2015-07-31")]
select_train = (df["sales"].gt(0) & df["date"].lt(validation_window[0]) & df["zscore"].
↳ abs().lt(5)).to_numpy()

select_val = (
    df["sales"].gt(0)
    & df["date"].ge(validation_window[0])
    & df["date"].lt(validation_window[1])
).to_numpy()

(select_train.sum(), select_val.sum())

[4]: (57876, 12502)
```

## 5.2.3 2. Fit baseline GLM

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We start with a simple model that uses only year, month, day of the week, and store as predictors. Even with these variables alone, we should still be able to capture a lot of valuable information. Year can capture overall sales trends, month can capture seasonality, week day can capture the variation in sales across the week, and store can capture locality. We will treat these all as categorical variables.

With the `GeneralizedLinearRegressor()` class, we can pass in `pandas.Categorical` variables directly without having to encode them ourselves. This is convenient, especially when we start adding more fixed effects. But it is very important that the categories are aligned between calls to `fit` and `predict`. One way of achieving this alignment is with a `dask_ml.preprocessing.Categorizer`. Note, however, that the `Categorizer` class fails to enforce category alignment if the input column is already a categorical data type.

You can reference the [pandas documentation on Categoricals](#) to learn more about how these data types work.



```
[5]: baseline_features = ["year", "month", "day_of_week", "store"]
baseline_categorizer = Categorizer(columns=baseline_features)
baseline_glm = GeneralizedLinearRegressor(
    family="gamma",
    scale_predictors=True,
    l1_ratio=0.0,
    alphas=1e-1,
)
```

Fit the model making sure to process the data frame with the Categorizer first and inspect the coefficients.

```
[6]: baseline_glm.fit(
    baseline_categorizer.fit_transform(df[select_train][baseline_features]),
    df.loc[select_train, "sales"]
)

pd.DataFrame(
    {'coefficient': np.concatenate([baseline_glm.intercept_, baseline_glm.coef_])},
    index=['intercept'] + baseline_glm.feature_names_
).T
```

```
[6]:
```

	intercept	year__2013	year__2014	year__2015	month__1	\	
coefficient	8.781	-0.010	0.008	0.003	-0.015		
	month__2	month__3	month__4	month__5	month__6	month__7	\
coefficient	-0.021	-0.019	0.020	0.008	-0.006	-0.002	
	month__8	month__9	month__10	month__11	month__12	\	
coefficient	-0.023	-0.034	-0.030	0.022	0.116		
	day_of_week__1	day_of_week__2	day_of_week__3	day_of_week__4	\		
coefficient	0.098	0.010	-0.014	-0.014			
	day_of_week__5	day_of_week__6	day_of_week__7	store__1	\		
coefficient	0.017	-0.100	0.285	-0.156			
	store__2	store__3	store__4	store__5	store__6	store__7	\
coefficient	-0.136	0.024	0.202	-0.161	-0.087	0.158	
	store__8	store__9	store__10	store__11	store__12	store__13	\
coefficient	-0.094	0.002	-0.081	0.110	0.074	-0.126	
	store__14	store__15	store__16	store__17	store__18	store__19	\
coefficient	-0.088	0.009	0.082	-0.020	0.001	-0.008	
	store__20	store__21	store__22	store__23	store__24	store__25	\
coefficient	0.085	-0.088	-0.178	-0.088	0.183	0.274	
	store__26	store__27	store__28	store__29	store__30	store__31	\
coefficient	0.008	0.192	-0.098	0.059	-0.101	-0.051	
	store__32	store__33	store__34	store__35	store__36	store__37	\
coefficient	-0.225	0.125	0.102	0.198	0.185	0.046	

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coefficient	store__38	store__39	store__40	store__41	store__42	store__43	\
	-0.048	-0.153	-0.128	-0.085	0.234	0.017	
coefficient	store__44	store__45	store__46	store__47	store__48	store__49	\
	-0.086	-0.101	-0.094	0.039	-0.243	0.058	
coefficient	store__50	store__51	store__52	store__53	store__54	store__55	\
	-0.208	0.012	0.079	-0.109	0.099	-0.180	
coefficient	store__56	store__57	store__58	store__59	store__60	store__61	\
	0.047	0.259	-0.017	-0.091	0.102	-0.149	
coefficient	store__62	store__63	store__64	store__65	store__66	store__67	\
	-0.009	0.034	0.251	-0.122	-0.035	0.086	
coefficient	store__68	store__69	store__70	store__71	store__72	store__73	\
	0.081	0.190	0.002	0.171	-0.189	-0.173	
coefficient	store__74	store__75	store__76	store__77	store__78	store__79	\
	-0.023	-0.038	0.172	0.077	-0.266	-0.091	
coefficient	store__80	store__81	store__82	store__83	store__84	store__85	\
	0.077	0.038	0.163	-0.248	0.369	0.019	
coefficient	store__86	store__87	store__88	store__89	store__90	store__91	\
	-0.155	-0.060	-0.048	-0.047	0.130	-0.059	
coefficient	store__92	store__93	store__94	store__95	store__96	store__97	\
	-0.047	-0.052	0.038	0.084	-0.121	0.007	
coefficient	store__98	store__99	store__100	store__101	store__102		
	-0.119	-0.111	0.108	0.069	0.038		
coefficient	store__103	store__104	store__105	store__106	store__107		
	-0.171	0.233	-0.175	0.136	0.058		
coefficient	store__108	store__109	store__110	store__111	store__112		
	0.284	-0.006	-0.174	0.018	-0.039		

And let's predict for our test set with the caveat that we will predict 0 for days when the stores are closed!

```
[7]: df.loc[lambda x: x["open"], "predicted_sales_baseline"] = baseline_glm.predict(
    baseline_categorizer.fit_transform(df.loc[lambda x: x["open"]][baseline_features])
)

df["predicted_sales_baseline"] = df["predicted_sales_baseline"].fillna(0)
df["predicted_sales_baseline"] = df["predicted_sales_baseline"]
```

We use root mean squared percentage error (RMSPE) as our performance metric. (Useful for thinking about error relative to total sales of each store).

```
[8]: train_err = root_mean_squared_percentage_error(
      df.loc[select_train, "sales"], df.loc[select_train, "predicted_sales_baseline"]
    )
    val_err = root_mean_squared_percentage_error(
      df.loc[select_val, "sales"], df.loc[select_val, "predicted_sales_baseline"]
    )
    print(f'Training Error: {round(train_err, 2)}%')
    print(f'Validation Error: {round(val_err, 2)}%')

Training Error: 27.96%
Validation Error: 29.57%
```

The results aren't bad for a start, but we can do better :)

### 5.2.4 3. GLM with high dimensional fixed effects

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Now, we repeat a similar process to above, but, this time, we take advantage of the full range of categoricals we created in our data transformation step. Since we will create a very large number of fixed effects, we may run into cases where our validation data has categorical values not seen in our training data. In these cases, Dask-ML's `Categorizer` will output null values when transforming the validation columns to the categoricals that were created on the training set. To fix this, we add Dask-ML's `SimpleImputer` to our pipeline.

```
[9]: highdim_features = [
      "age_quantile",
      "competition_open",
      "open_lag_1",
      "open_lag_2",
      "open_lag_3",
      "open_lead_1",
      "open_lead_2",
      "open_lead_3",
      "promo_lag_1",
      "promo_lag_2",
      "promo_lag_3",
      "promo_lead_1",
      "promo_lead_2",
      "promo_lead_3",
      "promo",
      "school_holiday_lag_1",
      "school_holiday_lag_2",
      "school_holiday_lag_3",
      "school_holiday_lead_1",
      "school_holiday_lead_2",
      "school_holiday_lead_3",
      "school_holiday",
      "state_holiday_lag_1",
      "state_holiday_lag_2",
      "state_holiday_lag_3",
      "state_holiday_lead_1",
      "state_holiday_lead_2",
      "state_holiday_lead_3",
    ]
```

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```

    "state_holiday",
    "store_day_of_week",
    "store_month",
    "store_school_holiday",
    "store_state_holiday",
    "store_year",
]
highdim_categorizer = Pipeline([
    ("categorize", Categorizer(columns=highdim_features)),
    ("impute", SimpleImputer(strategy="most_frequent"))
])
highdim_glm = GeneralizedLinearRegressor(
    family="gamma",
    scale_predictors=True,
    l1_ratio=0.0, # only ridge
    alpha=1e-1,
)

```

For reference, we output the total number of predictors after fitting the model. We can see that the number getting a bit larger, so we don't print out the coefficients this time.

```

[10]: highdim_glm.fit(
        highdim_categorizer.fit_transform(df[select_train][highdim_features]),
        df.loc[select_train, "sales"]
    )

print(f"Number of predictors: {len(highdim_glm.feature_names_)}")

```

Number of predictors: 2771

```

[11]: df.loc[lambda x: x["open"], "predicted_sales_highdim"] = highdim_glm.predict(
        highdim_categorizer.transform(df.loc[lambda x: x["open"]][highdim_features]),
    )

df["predicted_sales_highdim"] = df["predicted_sales_highdim"].fillna(0)
df["predicted_sales_highdim"] = df["predicted_sales_highdim"]

train_err = root_mean_squared_percentage_error(
    df.loc[select_train, "sales"], df.loc[select_train, "predicted_sales_highdim"]
)
val_err = root_mean_squared_percentage_error(
    df.loc[select_val, "sales"], df.loc[select_val, "predicted_sales_highdim"]
)
print(f'Training Error: {round(train_err, 2)}%')
print(f'Validation Error: {round(val_err, 2)}%')

```

Training Error: 12.54%  
Validation Error: 13.28%

From just the RMSPE, we can see a clear improvement from our baseline model.

### 5.2.5 4. Plot results

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Finally, to get a better look at our results, we make some plots.

```
[12]: sales_cols = ["sales", "predicted_sales_highdim", "predicted_sales_baseline"]
```

First, we plot true sales and the sales predictions from each model aggregated over month:

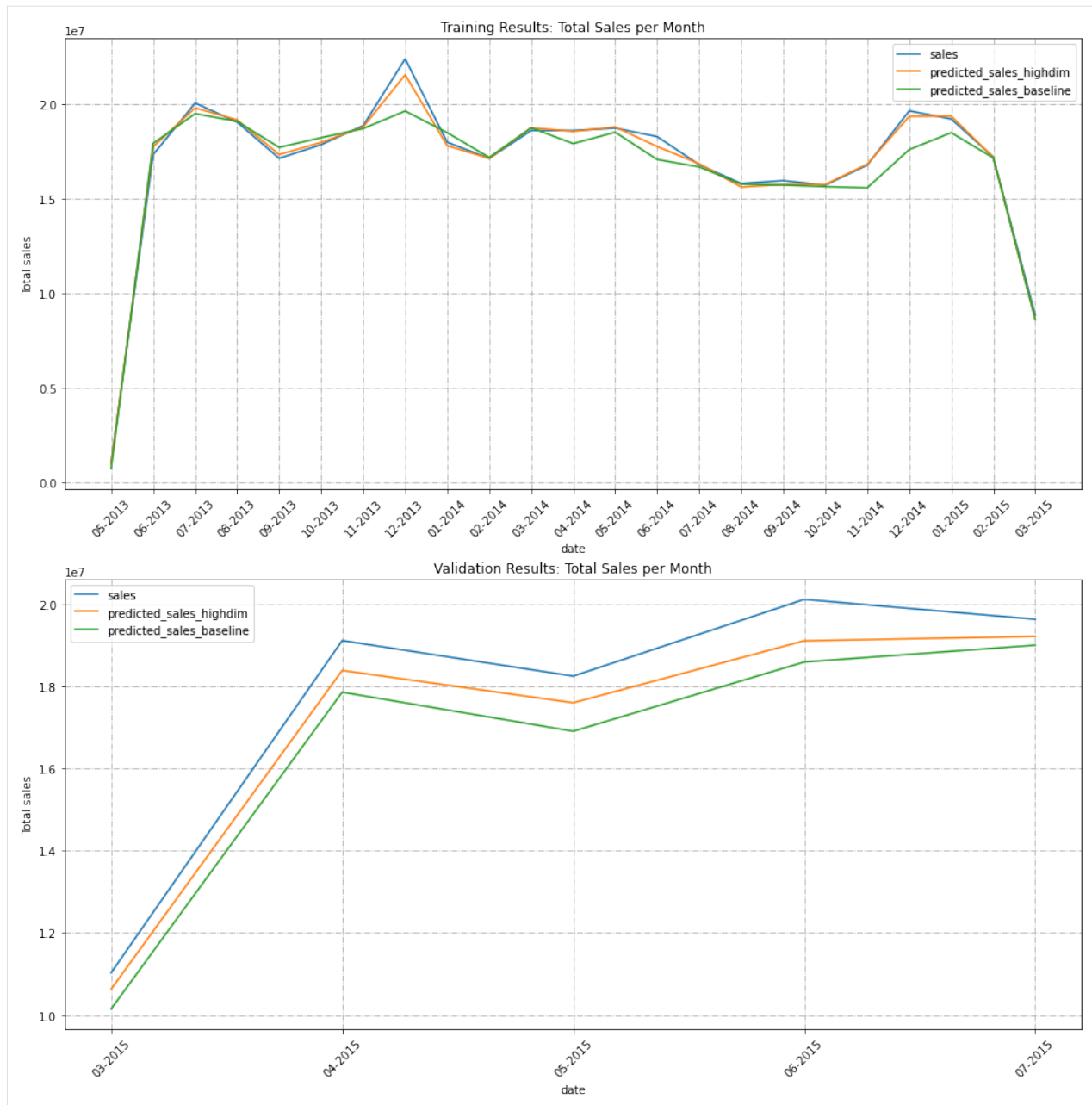
```
[13]: _, axs = plt.subplots(2, 1, figsize=(16, 16))

for i, select in enumerate([select_train, select_val]):
    ax = axs[i]
    df_plot_date = df[select].groupby(
        ["year", "month"]
    ).agg("sum")[sales_cols].reset_index()

    year_month_date = df_plot_date['month'].map(str) + '-' + df_plot_date['year'].map(str)
    df_plot_date['year_month'] = pd.to_datetime(year_month_date, format='%m-%Y').dt.
    ↪ strftime('%m-%Y')
    df_plot_date.drop(columns= ["year", "month"], inplace=True)

    df_plot_date.plot(x="year_month", ax=ax)
    ax.set_xticks(range(len(df_plot_date)))
    ax.set_xticklabels(df_plot_date.year_month, rotation=45)
    ax.set_xlabel("date")
    ax.set_ylabel("Total sales")
    ax.grid(True, linestyle='-.')

axs[0].set_title("Training Results: Total Sales per Month")
axs[1].set_title("Validation Results: Total Sales per Month")
plt.show()
```



We can also look at aggregate sales for a subset of stores. We select the first 20 stores and plot in order of increasing sales.

```
[14]: _, axs = plt.subplots(2, 1, figsize=(14, 12))
for i, select in enumerate([select_train, select_val]):
    ax = axs[i]
    df_plot_store = df[select].groupby(
        ["store"]
    ).agg("sum")[sales_cols].reset_index()[20].sort_values(by="sales")

    df_plot_store.plot.bar(x="store", ax=ax)
    ax.set_xlabel("Store")
```

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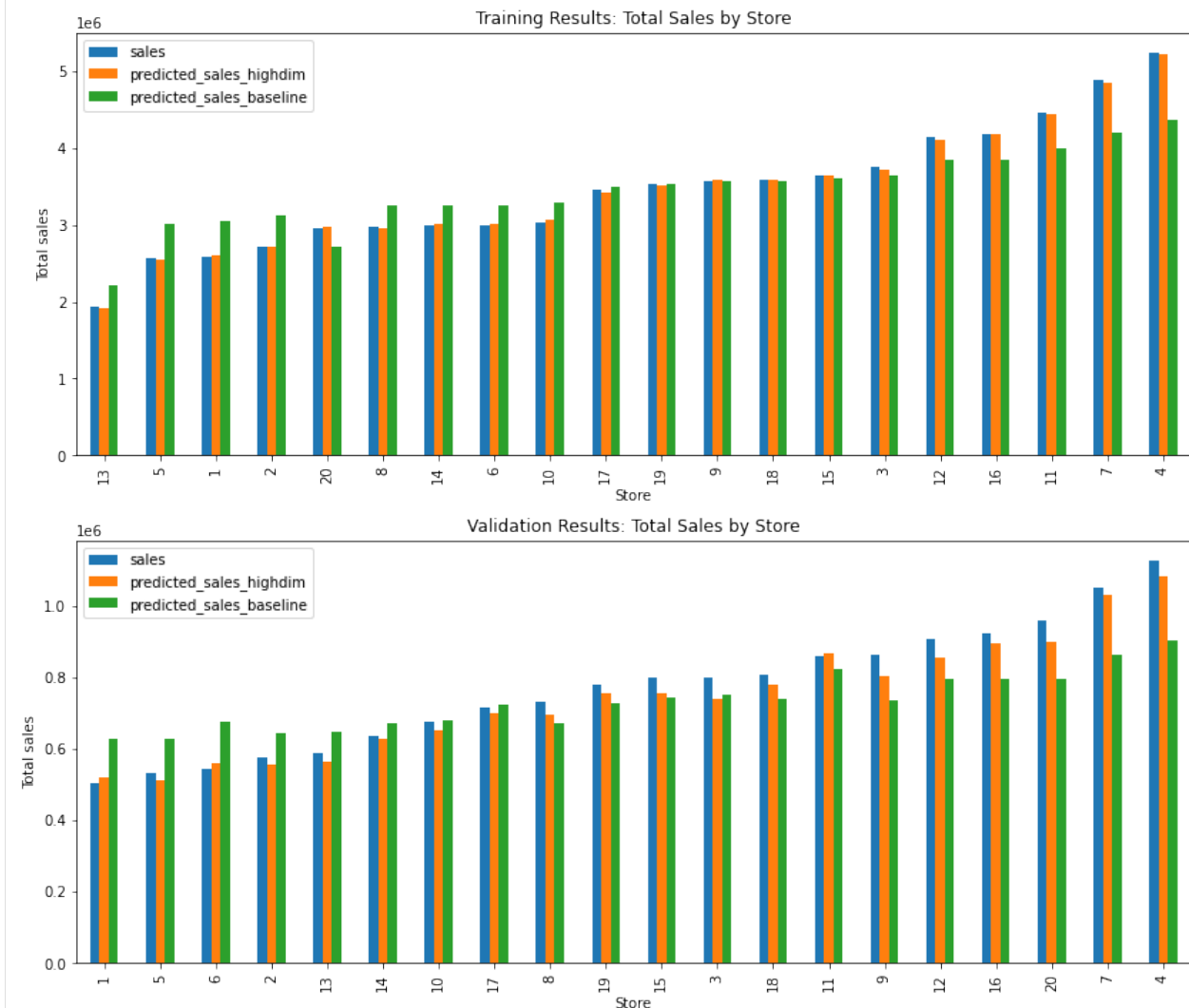
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```

ax.set_ylabel("Total sales")

axs[0].set_title("Training Results: Total Sales by Store")
axs[1].set_title("Validation Results: Total Sales by Store")
plt.show()

```



We can see that the high dimensional model is much better at capturing the variation between months and individual stores!

## 5.3 Tikhonov Regularization Tutorial: Seattle-Tacoma Housing Data

### Intro

This tutorial shows how to use variable  $L_2$  regularization with glum. The P2 parameter of the `GeneralizedLinearRegressor` class allows you to directly set the  $L_2$  penalty matrix  $w^T P_2 w$ . If a 2d array is passed for the P2 parameter, it is used directly, while if you pass a 1d array as P2 it will be interpreted as the diagonal of  $P_2$  and all other entries will be assumed to be zero.

*Note:* Variable  $L_1$  regularization is also available by passing an array with length `n_features` to the P1 parameter.

### Background

For this tutorial, we will model the selling price of homes in King's County, Washington (Seattle-Tacoma Metro area) between May 2014 and May 2015. However, in order to demonstrate a Tikhonov regularization-based spatial smoothing technique, we will focus on a small, skewed data sample from that region in our training data. Specifically, we will show that when we have (a) a fixed effect for each postal code region and (b) only a select number of training observations in a certain region, we can improve the predictive power of our model by regularizing the difference between the coefficients of neighboring regions. While we are constructing a somewhat artificial example here in order to demonstrate the spatial smoothing technique, we have found similar techniques to be applicable to real-world problems.

We will use a gamma distribution for our model. This choice is motivated by two main factors. First, our target variable, home price, is a positive real number, which matches the support of the gamma distribution. Second, it is expected that factors influencing housing prices are multiplicative rather than additive, which is better captured with a gamma regression than say, OLS.

*Note:* a few parts of this tutorial utilize local helper functions outside this notebook. If you wish to run the notebook on your own, you can find the rest of the code [here](#).

### 5.3.1 Table of Contents

- 1. Load and Prepare Datasets from Openml.org
- 2. Visualize Geographic Data with GIS Open Data
- 3. *Feature Selection and Transformation*
- 4. Create P matrix
- 5. *Fit Models*

```
[1]: import itertools

import geopandas as geopd
import libpysal
import matplotlib.pyplot as plt
import numpy as np
import openml
import pandas as pd

from sklearn.metrics import mean_squared_error
from sklearn.model_selection import train_test_split
from sklearn.preprocessing import StandardScaler
from glum import GeneralizedLinearRegressor

import sys
sys.path.append("../")
```

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```

from metrics import root_mean_squared_percentage_error

import warnings
warnings.filterwarnings("ignore", message="The weights matrix is not fully connected")

import data_prep
import maps

```

## 5.3.2 1. Load and prepare datasets from Openml

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### 1.1. Download and transform

The main dataset is downloaded from OpenML. You can find the main page for the dataset [here](#). It is also available through Kaggle [here](#).

As part of data preparation, we also do some transformations to the data:

- We remove some outliers (homes over 1.5 million and under 100k).
- Since we want to focus on geographic features, we also remove a handful of the other features.

Below, you can see some example rows from the dataset.

```
[2]: df = data_prep.download_and_transform()
df.head()
```

```
[2]:
```

	bedrooms	bathrooms	sqft_living	floors	waterfront	view	condition	\
0	3	1.00	1180	1.0	0	0	3	
1	3	2.25	2570	2.0	0	0	3	
2	2	1.00	770	1.0	0	0	3	
3	4	3.00	1960	1.0	0	0	5	
4	3	2.00	1680	1.0	0	0	3	

	sqft_basement	yr_built	zipcode	price
0	0	1955	98178	221900.0
1	400	1951	98125	538000.0
2	0	1933	98028	180000.0
3	910	1965	98136	604000.0
4	0	1987	98074	510000.0

## 5.3.3 2. Visualize geographic data with GIS open

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To help visualize the geographic data, we use geopandas and GIS Open Data to display price information on the King's county map. You can get the map data [here](#).

To show the relationship between home price and geography, we merge the map data with our sales data and use a heat map to plot mean home sale price for each postal code region.

```
[3]: maps.read_shapefile("Zip_Codes/Zip_Codes.shp")
```

```
[3]:
```

	OBJECTID	ZIP	ZIPCODE	COUNTY	SHAPE_Leng	SHAPE_Area	\
0	1	98031	98031	033	117508.211718	2.280129e+08	
1	2	98032	98032	033	166737.664791	4.826754e+08	
2	3	98033	98033	033	101363.840369	2.566747e+08	
3	4	98034	98034	033	98550.452509	2.725072e+08	
4	5	98030	98030	033	94351.264837	2.000954e+08	
..	...	...	...	...	...	...	
199	200	98402	98402	053	30734.178112	2.612224e+07	
200	201	98403	98403	053	23495.038425	2.890938e+07	
201	202	98404	98404	053	61572.154365	2.160645e+08	
202	203	98405	98405	053	50261.100559	1.193118e+08	
203	204	98406	98406	053	74118.972418	1.088373e+08	

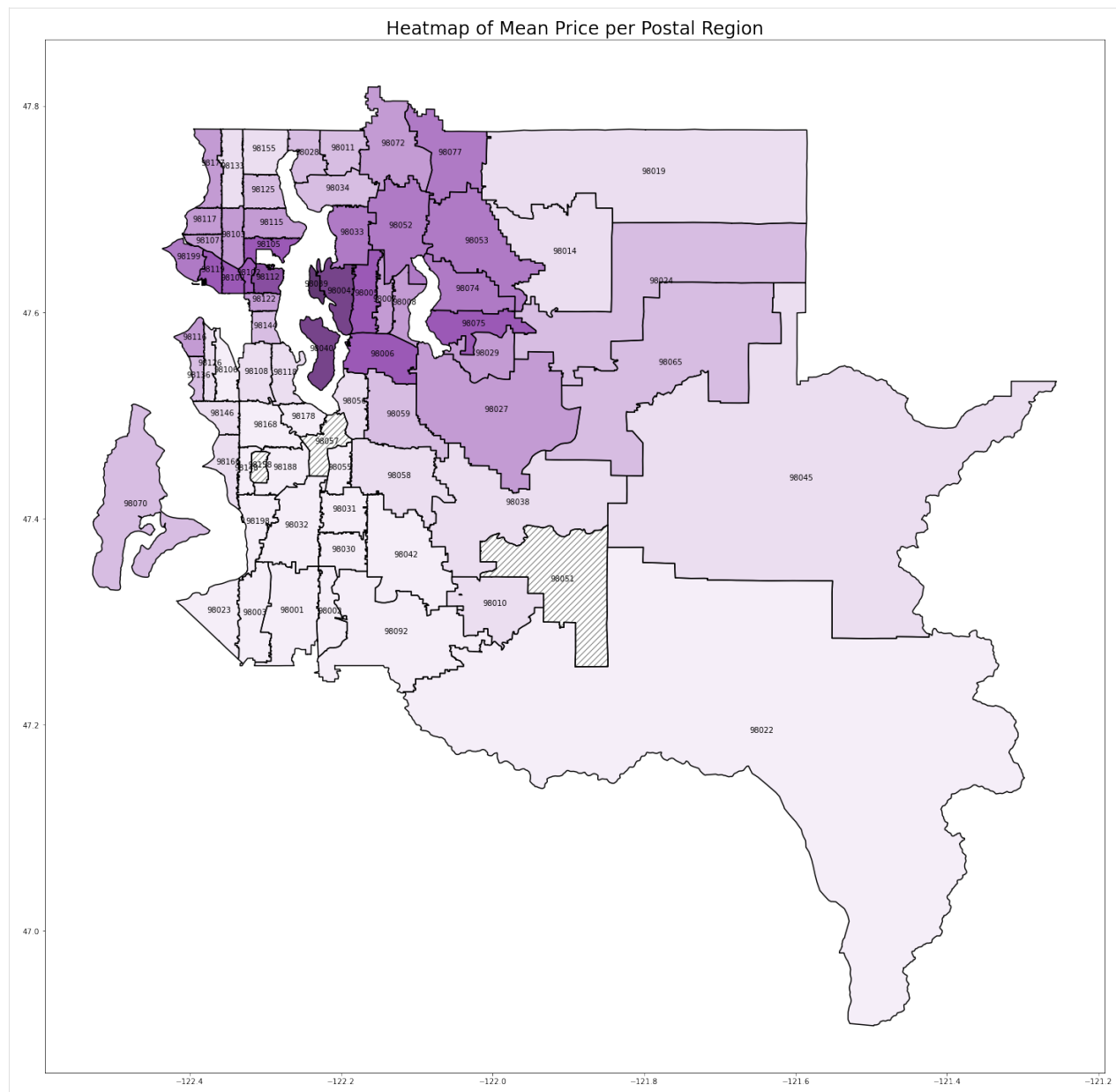
```

                                geometry
0  POLYGON ((-122.2184228967409 47.4375036485968,...
1  (POLYGON ((-122.2418694980486 47.4412158004961...
2  POLYGON ((-122.2057111926017 47.65169738162997...
3  POLYGON ((-122.1755100327681 47.73706057280546...
4  POLYGON ((-122.1674637459728 47.38548925033355...
..
199 POLYGON ((-122.4427945843513 47.2647926142345,...
200 POLYGON ((-122.4438167281511 47.26617469660845...
201 POLYGON ((-122.3889999141967 47.23495303304902...
202 POLYGON ((-122.4409198889526 47.23639133730699...
203 (POLYGON ((-122.5212509005256 47.2712095490982...

[204 rows x 7 columns]
```

```
[4]: df_shapefile = maps.read_shapefile("Zip_Codes/Zip_Codes.shp")
df_map = maps.create_kings_county_map_df(df, df_shapefile)

fix, ax = plt.subplots(figsize=(25, 25))
maps.plot_heatmap(df=df_map, label_col="ZIP", data_col="price", ax=ax)
ax.set_title("Heatmap of Mean Price per Postal Region", fontsize=24)
plt.show()
```



We can see a clear relationship between postal code and home price. Seattle (98112, 98102, etc.) and the Bellevue/Mercer/Medina suburbs (98039, 98004, 98040) have the highest prices. As you get further from the city, the prices start to drop.

### 5.3.4 3. Feature selection and transformation

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#### 3.1 Feature selection and one hot encoding

Since we want to focus on geographic data, we drop a number of columns below. We keep a handful of columns so that we can still create a reasonable model.

We then create a fixed effect for each of the postal code regions. We add the encoded postcode columns in numeric order to help us maintain the proper order of columns while building and training the model.

```
[5]: sorted_zips = sorted(list(df["zipcode"].unique()))
one_hot = pd.get_dummies(df["zipcode"], dtype=float)
one_hot = one_hot[sorted_zips]
df = df.drop('zipcode', axis=1)
df = one_hot.join(df)
df.head()
```

```
[5]:
```

	98001	98002	98003	98004	98005	98006	98007	98008	98010	98011	...	\
0	0	0	0	0	0	0	0	0	0	0	...	...
1	0	0	0	0	0	0	0	0	0	0	...	...
2	0	0	0	0	0	0	0	0	0	0	...	...
3	0	0	0	0	0	0	0	0	0	0	...	...
4	0	0	0	0	0	0	0	0	0	0	...	...

	bedrooms	bathrooms	sqft_living	floors	waterfront	view	condition	\
0	3	1.00	1180	1.0	0	0	3	
1	3	2.25	2570	2.0	0	0	3	
2	2	1.00	770	1.0	0	0	3	
3	4	3.00	1960	1.0	0	0	5	
4	3	2.00	1680	1.0	0	0	3	

	sqft_basement	yr_built	price
0	0	1955	221900.0
1	400	1951	538000.0
2	0	1933	180000.0
3	910	1965	604000.0
4	0	1987	510000.0

[5 rows x 80 columns]

#### 3.2 Test train split

As we mentioned in the introduction, we want to focus on modeling the selling price in a specific region while only using a very small, skewed data sample from that region in our training data. This scenario could arise if say, our task was to predict the sales prices for homes in Enumclaw (large region with zip code 98022 in the southeast corner of the map), but the only data we had from there was from a small luxury realtor.

To mimic this, instead creating a random split between our training and test data, we will intentionally create a highly skewed sample. For our test set, we will take all of the home sales in Enumclaw, *except* for the 15 highest priced homes.

Finally, we standardize our predictors.

```
[6]: predictors = [c for c in df.columns if c != "price"]

test_region = "98022"
df_train = df[df[test_region] == 0]
df_test = df[df[test_region] == 1].sort_values(by="price", ascending=False)

test_to_train = df_test[:15]

df_train = pd.concat([df_train, test_to_train])
df_test = df_test.drop(test_to_train.index)

X_train = df_train[predictors]
y_train = df_train["price"]
X_test = df_test[predictors]
y_test = df_test["price"]

scaler = StandardScaler().fit(X_train)
X_train = scaler.transform(X_train)
X_test = scaler.transform(X_test)
```

### 5.3.5 4. Creating the penalty matrix

*[back to table of contents](#)*

To smooth the coefficients for neighboring regions, we will create a penalty matrix  $P$  such that we penalize the squared difference in coefficient values for neighbouring regions, e.g. for 98022 and 98045. For example, if 98022 and 98045 were the only region in question, we would need a  $2 \times 2$  matrix  $P$  such that:

$$(\beta_{98022}, \beta_{98045}) P \begin{pmatrix} \beta_{98022} \\ \beta_{98045} \end{pmatrix} = (\beta_{98022} - \beta_{98045})^2$$

In this example, we would get this result with  $P = \begin{pmatrix} 1 & -1 \\ -1 & 1 \end{pmatrix}$ .

Since we have 72 postal code regions, it would be rather annoying to construct this matrix by hand. Luckily, there are libraries that exist for this. We use `pysal`'s `pysal.lib.weights.Queen` to retrieve a neighbor's matrix from our map data. The construction of the penalty matrix is rather straightforward once we have this information.

We leave the non-geographic features unregularized (all zeros in the  $P$  matrix).

```
[7]: # format is {zip1: {neighbor1: 1, neighbor2: 1, ...}}
neighbor_matrix = libpysal.weights.Queen.from_dataframe(df_map, ids="ZIP")

n_features = X_train.shape[1]
P2 = np.zeros((n_features, n_features))

zip2index = dict(zip(sorted_zips, range(len(sorted_zips))))
for zip1 in sorted_zips:
    for zip2 in neighbor_matrix[zip1].keys():
        if zip1 in zip2index and zip2 in zip2index: # ignore regions w/o data
            if zip2index[zip1] < zip2index[zip2]: # don't repeat if already saw neighbor
                ↪ pair in earlier iteration
                P2[zip2index[zip1], zip2index[zip2]] += 1
```

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```

P2[zip2index[zip2], zip2index[zip2]] += 1
P2[zip2index[zip1], zip2index[zip2]] -= 1
P2[zip2index[zip2], zip2index[zip1]] -= 1

```

P2

```

[7]: array([[ 3., -1., -1., ...,  0.,  0.,  0.],
          [-1.,  4.,  0., ...,  0.,  0.,  0.],
          [-1.,  0.,  4., ...,  0.,  0.,  0.],
          ...,
          [ 0.,  0.,  0., ...,  0.,  0.,  0.],
          [ 0.,  0.,  0., ...,  0.,  0.,  0.],
          [ 0.,  0.,  0., ...,  0.,  0.,  0.]])

```

### 5.3.6 5. Fit models

*[back to table of contents](#)*

Now, we will fit several L2 regularized OLS models using different levels of regularization. All will use the penalty matrix defined above, but the alpha parameter, the constant that multiplies the penalty terms and thus determines the regularization strength, will vary.

For each model, we will measure test performance using root mean squared percentage error (RMSPE), so that we can get a relative result. We will also plot a heatmap of the coefficient values over the regions.

*Note:* alpha=1e-12 is effectively no regularization. But we can't set alpha to zero because the unregularized problem has co-linear columns, resulting in a singular design matrix.

```

[8]: fig, axs = plt.subplots(nrows=2, ncols=2, figsize=(20, 20))
    for i, alpha in enumerate([1e-12, 1e-1, 1, 10]):

        glm = GeneralizedLinearRegressor(family='gamma', alpha=alpha, P2=P2, fit_
        ↪ intercept=True)
        glm.fit(X_train, y_train)
        y_test_hat = glm.predict(X_test)

        coeffs = pd.DataFrame({'coefficient': np.concatenate([glm.intercept_, glm.coef_])},
        ↪ ["intercept"]+predictors)

        print(f"alpha={alpha}")
        print(f"Test region coefficient: {coeffs.loc[test_region].values[0]}")
        print(f"Test RMSPE: {root_mean_squared_percentage_error(y_test_hat, y_test)}\n")

        df_map_coeffs = df_map.merge(
            coeffs.loc[sorted_zips],
            left_on="ZIP",
            right_index=True,
            how="outer"
        )

        ax = axs[i//2, i%2]
        df_map_coeffs["annotation"] = df_map_coeffs["ZIP"].apply(lambda x: "" if x!=test_
        ↪ region else x)
        maps.plot_heatmap(

```

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```
df=df_map_coeffs,
label_col="annotation",
data_col="coefficient",
ax=ax,
vmin=-0.015,
vmax=0.025
)
ax.set_title(f"alpha={alpha}")

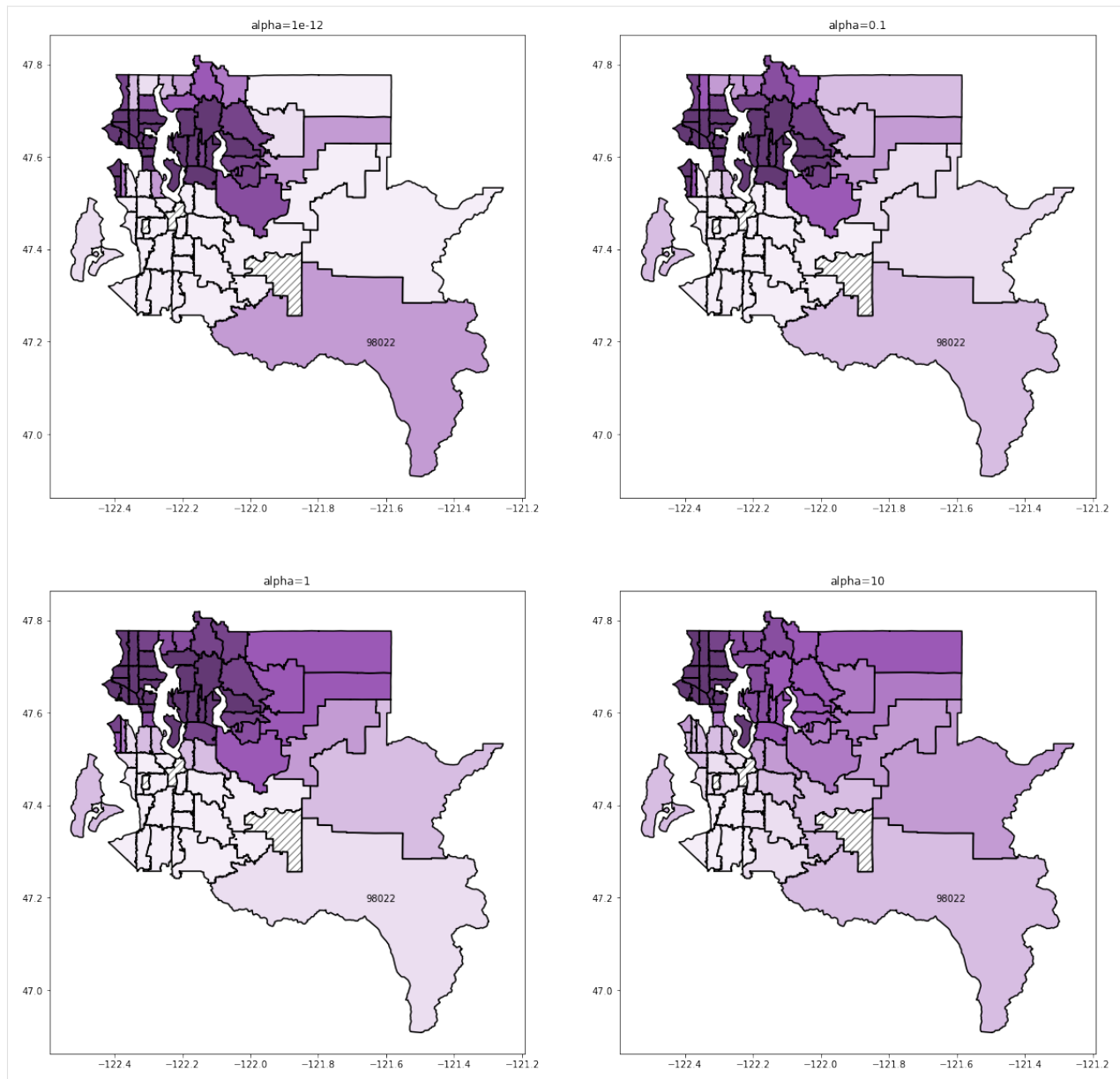
plt.show()
```

alpha=1e-12  
Test region coefficient: 0.0010920106922960072  
Test RMSPE: 72.65620542354644

alpha=0.1  
Test region coefficient: -0.0036087215513505183  
Test RMSPE: 43.926082004444204

alpha=1  
Test region coefficient: -0.01041392075707663  
Test RMSPE: 19.51113178158937

alpha=10  
Test region coefficient: -0.0033476740903954213  
Test RMSPE: 44.59786775358339



$\alpha=1$  seems to recover the best results. Remember that our test dataset is just a small subset of the data in region 98022 and that the training data is skewed towards high sales prices. For  $\alpha$  less than 1, we can see that the 98022 region coefficient is still much greater than its neighbors coefficients, which we can see is not accurate if we refer back to map we produced based on the raw data. For higher  $\alpha$  levels, we start to see poor predictions resulting from regional coefficients that are too smooth between adjacent regions.

A test RMSPE of 19.5% is a surprisingly good result considering that we only had 10 highly skewed observations from our test region in our training data and is far better than the RMSPE of 67.5% from the unregularized case.



---

## CONTRIBUTING AND DEVELOPMENT

Hello! And thanks for exploring glum more deeply. Please see the issue tracker and pull requests tabs on Github for information about what is currently happening. Feel free to post an issue if you'd like to get involved in development and don't really know where to start – we can give some advice.

We welcome contributions of any kind!

- New features
- Feature requests
- Bug reports
- Documentation
- Tests
- Questions

### 6.1 Pull request process

- Before working on a non-trivial PR, please first discuss the change you wish to make via issue, Slack, email or any other method with the owners of this repository. This is meant to prevent spending time on a feature that will not be merged.
- Please make sure that a new feature comes with adequate tests. If these require data, please check if any of our existing test data sets fits the bill.
- Please make sure that all functions come with proper docstrings. If you do extensive work on docstrings, please check if the Sphinx documentation renders them correctly. ReadTheDocs builds on every commit to an open pull request. You can see whether the documentation has successfully built in the “checks” section of the PR. Once the build finishes, your documentation should be accessible by clicking the “details” link next to the check in the GitHub interface and will appear at a URL like: <https://glum-###.org.readthedocs.build/en/###/> where ### is the number of your PR.
- Please make sure you have our pre-commit hooks installed.
- If you fix a bug, please consider first contributing a test that `_fails_` because of the bug and then adding the fix as a separate commit, so that the CI system picks it up.
- Please add an entry to the change log and increment the version number according to the type of change. We use semantic versioning. Update the major if you break the public API. Update the minor if you add new functionality. Update the patch if you fixed a bug. All changes that have not been released are collected under the date UNRELEASED.

## 6.2 Releases

- We make package releases infrequently, but usually any time a new non-trivial feature is contributed or a bug is fixed. To make a release, just open a PR that updates the change log with the current date. Once that PR is approved and merged, you can create a new release on [GitHub](<https://github.com/Quantco/glum/releases/new>). Use the version from the change log as tag and copy the change log entry into the release description.

## 6.3 Install for development

The first step is to set up a conda environment and install glum in editable mode. We strongly suggest to use mamba instead of conda as this provides the same functionality at much greater speed.

```
# First, make sure you have conda-forge as your primary conda channel:
conda config --add channels conda-forge

# Clone the repository
git clone git@github.com:Quantco/glum.git
cd glum

# Set up a conda environment with name "glum"
mamba env create

# If you want to install the dependencies necessary for benchmarking against other GLM_
↪ packages:
mamba env update -n glum --file environment-benchmark.yml

# If you want to work on the tutorial notebooks:
mamba env update -n glum --file environment-tutorial.yml

# Activate the previously created conda environment
conda activate glum

# Set up our pre-commit hooks for black, mypy, isort and flake8.
pre-commit install

# Install this package in editable mode.
pip install --no-use-pep517 --disable-pip-version-check -e .
```

## 6.4 Testing and continuous integration

The test suite is in `tests/`.

### 6.4.1 Golden master tests

We use golden master testing to preserve correctness. The results of many different GLM models have been saved. After an update, the tests will compare the new output to the saved models. Any significant deviation will result in a test failure. This doesn't strictly mean that the update was wrong. In case of a bug fix, it's possible that the new output will be more accurate than the old output. In that situation, the golden master results can be overwritten as explained below.

There are two sets of golden master tests, one with artificial data and one directly using the benchmarking problems from `glum_benchmarks`. For both sets of tests, creating the golden master and the tests definition are located in the same file. Calling the file with `pytest` will run the tests while calling the file as a python script will generate the golden master result. When creating the golden master results, both scripts accept the `--overwrite` command line flag. If set, the existing golden master results will be overwritten. Otherwise, only the new problems will be run.

### 6.4.2 Skipping the slow tests

If you want to skip the slow tests, add the `-m "not slow"` flag to any `pytest` command. The “wide” problems (all marked as slow tests) are especially poorly conditioned. This means that even for estimation with 10k observations, it might still be very slow. Furthermore, we also have golden master tests for the “narrow” and “intermediate” problems, so adding the “wide” problems do not add much coverage.

### 6.4.3 Storing and modifying

To store the golden master results:

```
python tests/glm/test_golden_master.py
python tests/glm/test_benchmark_golden_master.py
```

Add the `--overwrite` flag if you want to overwrite already existing golden master results

## 6.5 Building a conda package

To use the package in another project, we distribute it as a conda package. For building the package locally, you can use the following command:

```
conda build conda.recipe
```

This will build the recipe using the standard compiler flags set by the conda-forge activation scripts.

## 6.6 Developing the documentation

The documentation is built with a mix of Sphinx, autodoc, and nbsphinx. To develop the documentation:

```
cd docs
make html
python -m http.server --directory _build/html
```

Then, navigate to <http://localhost:8000> to view the documentation.

Alternatively, if you install [entr](#), then you can auto-rebuild the documentation any time a file changes with:

```
cd docs
./dev
```

---

**Note:** The tutorial notebooks are not executed as part of the documentation build. If you want to modify them, make sure to execute them manually and save the output before committing. Also don't forget to install the extra dependencies for the tutorial notebooks as described above.

---

If you are a newbie to Sphinx, the links below may help get you up to speed on some of the trickier aspects:

- [An idiot's guide to Sphinx](#)
- [Links between documents](#)
- [Cross-referencing python objects](#) using things like `:mod:` and `:meth:` and `:class:`.
- [autodoc](#) is used for automatically converting docstrings to docs
- [We follow the numpy docstring style guide](#)
- [To create links between ipynb files when using nbsphinx](#)

## 6.7 Where to start looking in the source?

The primary user interface of `glum` consists of the `GeneralizedLinearRegressor` and `GeneralizedLinearRegressorCV` classes via their constructors and the `fit()` and `predict()` functions. Those are the places to start looking if you plan to change the system in some way.

What follows is a high-level summary of the source code structure. For more details, please look in the documentation and docstrings of the relevant classes, functions and methods.

- `_glm.py` - This is the main entrypoint and implements the core logic of the GLM. Most of the code in this file handles input arguments and prepares the data for the GLM fitting algorithm.
- `_glm_cv.py` - This is the entrypoint for the cross validated GLM implementation. It depends on a lot of the code in `_glm.py` and only modifies the sections necessary for running training many models with different regularization parameters.
- `_solvers.py` - This contains the bulk of the IRLS and L-BFGS algorithms for training GLMs.
- `_cd_fast.pyx` - This is a Cython implementation of the coordinate descent algorithm used for fitting L1 penalty GLMs. Note the `.pyx` extension indicating that it is a Cython source file.
- `_distribution.py` - definitions of the distributions that can be used. Includes Normal, Poisson, Gamma, InverseGaussian, Tweedie, Binomial and GeneralizedHyperbolicSecant distributions.

- `_link.py` - definitions of the link functions that can be used. Includes identity, log, logit and Tweedie link functions.
- `_functions.pyx` - This is a Cython implementation of the log likelihoods, gradients and Hessians for several popular distributions.
- `_util.py` - This contains a few general purpose linear algebra routines that serve several other modules and don't fit well elsewhere.

## 6.8 The GLM benchmark suite

Before deciding to build a library custom built for our purposes, we did an thorough investigation of the various open source GLM implementations available. This resulted in an extensive suite of benchmarks for comparing the correctness, runtime and availability of features for these libraries.

The benchmark suite has two command line entrypoints:

- `glm_benchmarks_run`
- `glm_benchmarks_analyze`

Both of these CLI tools take a range of arguments that specify the details of the benchmark problems and which libraries to benchmark.

For more details on the benchmark suite, see the README in the source at `src/glum_benchmarks/README.md`.



## AN INTRODUCTION TO THE ALGORITHMS USED IN GLUM

Before continuing, please take a look at the [sklearn documentation](#) for a high-level intro to generalized linear models.

In addition, please take a look at [the tutorials](#).

For mathematical and algorithmic details, see below:

### 7.1 What is a GLM?

This package is intended to fit L1 and L2-norm penalized generalized linear models (GLMs). What is a GLM?

A GLM is a linear model ( $\eta = x^\top \beta$ ) wrapped in a transformation (link function) and equipped with a response distribution from an exponential family. The choice of link function and response distribution is very flexible. In a GLM, a predictive distribution for the response variable  $Y$  is associated with a vector of observed predictors  $x$ . The distribution has the form:

$$\begin{aligned} p(y|x) &= m(y, \phi) \exp\left(\frac{\theta T(y) - A(\theta)}{\phi}\right) && \text{random component / distribution family} \\ \theta &:= g(\eta) && \text{systematic component / link function} \\ \eta &:= x^\top \beta && \text{parametric link component / linear predictor} \end{aligned} \tag{7.1}$$

Here  $\beta$  are the parameters;  $\phi$  is a parameter representing dispersion (“variance”); and  $m$ ,  $T$ ,  $A$  are characterized by the user-specified model family; and  $g$  is the **link function**. For background on the exponential family, and how common distributions can be represented in the form of the “random component” above, see [Wikipedia | Exponential family](#).

The expected mean of  $Y$  depends on  $x$  by composition of **linear response**  $\eta$  and (inverse) link function, i.e.:

$$E[y|\eta] := \mu = g^{-1}(\eta).$$

Proving this involves use of [moment-generating functions](#) and [cumulants](#).

Defining an offset of  $\gamma$ , we can write the linear predictor as  $\eta = X^T \beta + \gamma$ .

### 7.2 Examples of GLMs

#### 7.2.1 Normal

Many of the “regression” models we commonly work with are GLMs. For example, the normal distribution with a linear link function is a simple GLM:  $y|x \sim N(x^T \beta, \sigma^2)$ . If we fit  $\beta$  with maximum likelihood, this is least-squares regression. If we replace the linear link function with an exponential link function, we get a simple multiplicative model that is also a GLM:  $y|x \sim N(e^{x^T \beta}, \sigma^2)$ .

## 7.2.2 Tweedie

In the insurance context, we usually work with a Tweedie distribution, which is a generalization of Poisson ( $p = 1$ ) and Gamma ( $p = 2$ ):

$$\theta = \begin{cases} \frac{\mu^{1-p}}{1-p} & \text{if } p \neq 1 \\ \log \mu & \text{if } p = 1 \end{cases},$$

$$T(y) = y,$$

$$A(\theta) = \begin{cases} \frac{\mu^{2-p}}{2-p} & \text{if } p \neq 2 \\ \log \mu & \text{if } p = 2. \end{cases},$$

With  $1 < p < 2$ , the Tweedie distribution is a [compound Poisson-gamma](#) distribution.  $y$  is distributed as if a number  $N$  is drawn from a Poisson distribution, and then  $N$  draws are taken IID from a gamma distribution and added. This distribution might model the total amount of a claim in an insurance context: There are  $N$  incidents, and each incident  $i$  has an amount  $X_i$ , and the total amount is the sum of  $X_i$ .

## 7.3 Objective function

To fit a GLM, we minimize the negative log-likelihood (or typically the unit deviance) subject to an elastic net constraint involving a mix of L1 and L2 penalty terms:

$$\min_{\beta} \mathcal{L} + \alpha \rho \|\beta\|_1 + \frac{\alpha(1-\rho)}{2} \|\beta\|_2^2$$

## 7.4 Optimizers/solvers

There are three solvers implemented in `glum`:

- `lbfgs` - This solver uses the `scipy fmin_l_bfgs_b` optimizer to minimize L2-penalized GLMs. The L-BFGS solver does not work with L1-penalties. Because L-BFGS does not store the full Hessian, it can be particularly effective for very high dimensional problems with several thousand or more columns. For more details, see [the scipy documentation](#).
- `irls-cd` and `irls-ls` - These solvers are both based on Iteratively Reweighted Least Squares (IRLS). IRLS proceeds by iteratively approximating the objective function with a quadratic, then solving that quadratic for the optimal update. For purely L2-penalized settings, the `irls-ls` uses a least squares inner solver for each quadratic subproblem. For problems that have any L1-penalty component, the `irls-cd` uses a coordinate descent inner solver for each quadratic subproblem. The IRLS-LS and IRLS-CD implementations largely follow the algorithm described in `newglmnet` (see references below).

### 7.4.1 Convexity

Our objective function will generally be convex, because the log-likelihoods of members of the exponential family are convex in their “natural parameterizations.” The natural parameterization may not be the most obvious one. Example: The log-likelihood of the normal distribution is convex in one over the variance, which is its natural parameterization. It is not convex in the variance. We can generally assume we are solving convex problems.

An exception is multicollinearity (rank deficiency) in the design matrix, without an L2 component to the penalty. In that case, the problem will be only weakly convex and will have no unique minimum. This is not an arcane consideration, since we frequently generate rank deficiency by constructing multiple sets of one-hot encoded categorical variables. This can make evaluating different optimizers tricky, since they could converge to different equally good optima. The



original `glmnet` paper suggests using at least a small L2 regularization component to remove “degeneracies and wild behavior.”

### 7.4.2 IRLS

We minimize the log likelihood using Iteratively Reweighted Least Squares (IRLS). IRLS can be justified by seeing it as taking a Newton step, but replacing the Hessian with the expected Hessian.

In the `irls-cd` and `irls-ls` solvers, the outer loop is an IRLS iteration that forms a quadratic approximation to the negative loglikelihood. That is, we find  $w$  and  $z$  so that the problem can be expressed as:

$$\min \sum_i w_i (z_i - x_i \beta)^2 + \text{penalty}$$

We exit when either the gradient is small (`gradient_tol`) or the step size is small (`step_size_tol`). Both of these tolerances are user configurable.

Once we have formed this quadratic approximation, an “inner solver” finds the minimum of the quadratic. In the `irls-ls` solver, the inner solver is simply a direct least squares solve.

See the `glmintro` reference for an excellent discussion of IRLS in the context of GLMs.

### 7.4.3 Coordinate Descent

With an L1 penalty, we use the `irls-cd` solver where the inner solver finds the minimum of the quadratic using coordinate descent. We exit the inner loop when the quadratic problem’s gradient is small. The classic reference here is the `glmnet` paper.

However, coordinate descent is older than the `glmnet` paper, and is a simple idea. In a problem with data  $y$  and  $x$  and parameters “params”, Coordinate Descent involves repeatedly optimizing one or several parameters while holding the rest fixed. Interestingly, Wikipedia [claims](#) that coordinate descent is often overlooked among researchers because it is simple to implement, and they would rather work on something more interesting.

```
[1]: """
Cyclical coordinate descent with Newton-like steps and a
twice-differentiable objective function.
"""
def coordinate_descent_inner(y, x, params, grad_func, hess_func):
    for i, elt in enumerate(params):
        step = -grad_func(y, x, params) / hess_func(y, x, params)
        params[i] += step
    return params

def coordinate_descent(y, x, params, grad_func, hess_func,
                      convergence_func):
    while not convergence_func():
        params = coordinate_descent_inner(y, x, params, grad_func,
                                         hess_func)
    return params
```

In practice, the implementation is a bit more complicated when the objective function is not differentiable. See also the `coordinate_descent` reference below.

### 7.4.4 Active set tracking

When penalizing with an L1-norm, it is common for many coefficients to be exactly zero. And, it is possible to predict during a given iteration which of those coefficients will stay zero. As a result, we track the “active set” consisting of all the coefficients that are either currently non-zero or likely to remain non-zero. We follow the outer loop active set tracking algorithm in the `newglmnet` reference. That paper refers to the same concept as “shrinkage”, whereas the `glmnet` reference calls this the “active set”. Currently, we have not yet implemented the inner loop active set tracking from the `newglmnet` reference.

## 7.5 Matrix Types

Along with the GLM solvers, this package supports dense, sparse, categorical matrix types and mixtures of these types. Using the most efficient matrix representations massively improves performances.

For more details, see the [README](#) for `tabmat`

- We support dense matrices via standard numpy arrays.
- We support sparse CSR and CSC matrices via standard `scipy.sparse` objects. However, we have extended these operations with custom matrix-vector and sandwich product routines that are optimized and parallelized. A user does not need to modify their code to take advantage of this optimization. If a `scipy.sparse.csc_matrix` object is passed in, it will be automatically converted to a `SparseMatrix` object. This operation is almost free because no data needs to be copied.
- We implement a `CategoricalMatrix` object that efficiently represents these matrices without nearly as much overhead as a normal CSC or CSR sparse matrix.
- Finally, `SplitMatrix` allows mixing different matrix types for different columns to minimize overhead.

## 7.6 Standardization

Internal to our solvers, all matrix types are wrapped in a `tabmat.StandardizedMatrix` which offsets columns to have mean zero and standard deviation one without modifying the matrix data itself. This avoids situations where modifying a matrix to have mean zero would result in losing the sparsity structure. It also avoids ever needing to copy or modify the input data matrix. As a result, excess memory usage is very low in `glum`.

## 7.7 References

`glmnet` - [Regularization Paths for Generalized Linear Models via Coordinate Descent](#)

`newglmnet` - [An Improved GLMNET for L1-regularized LogisticRegression](#)

`glmintro` - [Bryan Lewis on GLMs](#)

`coordinate_descent` - [Coordinate Descent Algorithms](#)

`glmbook` - [Generalized Linear Models](#), McCullagh and Nelder

[ ]:

## GLUM PACKAGE

The two main classes in *glum* are *GeneralizedLinearRegressor* and *GeneralizedLinearRegressorCV*. Most users will use *fit()* and *predict()*

**class** *glum.BinomialDistribution*

Bases: *ExponentialDispersionModel*

A class for the Binomial distribution.

The Binomial distribution is for targets *y* in  $[0, 1]$ .

**deviance**(*y*, *mu*, *sample\_weight=1*)

Compute the deviance.

The deviance is a weighted sum of the unit deviances,  $\sum_i s_i \times d(y_i, \mu_i)$ , where  $d(y, \mu)$  is the unit deviance and *s* are weights. In terms of the log likelihood, it is  $-2\phi \times [L(y, \mu, \phi/s) - L(y, y, \phi/s)]$ .

### Parameters

- **y** (array-like, shape (n\_samples,)) – Target values.
- **mu** (array-like, shape (n\_samples,)) – Predicted mean.
- **sample\_weight** (array-like, shape (n\_samples,)) optional (default=1) – Weights or exposure to which variance is inversely proportional.

### Return type

float

**deviance\_derivative**(*y*, *mu*, *sample\_weight=1*)

Compute the derivative of the deviance with respect to *mu*.

### Parameters

- **y** (array-like, shape (n\_samples,)) – Target values.
- **mu** (array-like, shape (n\_samples,)) – Predicted mean.
- **sample\_weight** (array-like, shape (n\_samples,)) (default=1) – Weights or exposure to which variance is inverse proportional.

### Return type

array-like, shape (n\_samples,)

**dispersion**(*y*, *mu*, *sample\_weight=None*, *ddof=1*, *method='pearson'*)

Estimate the dispersion parameter  $\phi$ .

### Parameters

- **y** (array-like, shape (n\_samples,)) – Target values.

- **mu** (*array-like, shape (n\_samples,)*) – Predicted mean.
- **sample\_weight** (*array-like, shape (n\_samples,)*, *optional (default=None)*) – Weights or exposure to which variance is inversely proportional.
- **ddof** (*int, optional (default=1)*) – Degrees of freedom consumed by the model for mu.
- **{'pearson' (method =)}** – Whether to base the estimate on the Pearson residuals or the deviance.
- **'deviance'** – Whether to base the estimate on the Pearson residuals or the deviance.
- **(default='pearson') (optional)** – Whether to base the estimate on the Pearson residuals or the deviance.

**Return type**

float

**eta\_mu\_deviance**(*link, factor, cur\_eta, X\_dot\_d, y, sample\_weight*)

Compute eta, mu and the deviance.

Compute: \* the linear predictor, eta, as `cur_eta + factor * X_dot_d`; \* the link-function-transformed prediction, mu; \* the deviance.**Returns**

- *numpy.ndarray, shape (X.shape[0],)* – The linear predictor, eta.
- *numpy.ndarray, shape (X.shape[0],)* – The link-function-transformed prediction, mu.
- *float* – The deviance.

**Parameters**

- **link** (*Link*) –
- **factor** (*float*) –
- **cur\_eta** (*ndarray*) –
- **X\_dot\_d** (*ndarray*) –
- **y** (*ndarray*) –
- **sample\_weight** (*ndarray*) –

**in\_y\_range**(*x*)

Return True if x is in the valid range of the EDM.

**Parameters****x** (*array-like, shape (n\_samples,)*) – Target values.**Return type**

np.ndarray

**log\_likelihood**(*y, mu, sample\_weight=None, dispersion=1*)

Compute the log likelihood.

**Parameters**

- **y** (*array-like, shape (n\_samples,)*) – Target values.
- **mu** (*array-like, shape (n\_samples,)*) – Predicted mean.
- **sample\_weight** (*array-like, shape (n\_samples,)*, *optional (default=1)*) – Sample weights.

- **dispersion** (*float, optional (default=1)*) – Ignored.

**Return type**

float

**rowwise\_gradient\_hessian**(*link, coef, dispersion, X, y, sample\_weight, eta, mu, offset=None*)

Compute the gradient and negative Hessian of the log likelihood row-wise.

**Returns**

- *numpy.ndarray, shape (X.shape[0],)* – The gradient of the log likelihood, row-wise.
- *numpy.ndarray, shape (X.shape[0],)* – The negative Hessian of the log likelihood, row-wise.

**Parameters**

- **link** ([Link](#)) –
- **coef** (*ndarray*) –
- **X** (*MatrixBase | StandardizedMatrix*) –
- **y** (*ndarray*) –
- **sample\_weight** (*ndarray*) –
- **eta** (*ndarray*) –
- **mu** (*ndarray*) –
- **offset** (*ndarray | None*) –

**unit\_deviance**(*y, mu*)

Get the unit-level deviance.

See superclass documentation.

**Parameters**

- **y** (*array-like*) –
- **mu** (*array-like*) –

**Return type**

array-like

**unit\_deviance\_derivative**(*y, mu*)Compute the derivative of the unit deviance with respect to **mu**.The derivative of the unit deviance is given by  $-2 \times (y - \mu)/v(\mu)$ , where  $v(\mu)$  is the unit variance.**Parameters**

- **y** (*array-like, shape (n\_samples,)*) – Target values.
- **mu** (*array-like, shape (n\_samples,)*) – Predicted mean.

**Return type**

array-like, shape (n\_samples,)

**unit\_variance**(*mu*)

Get the unit-level expected variance.

See superclass documentation.

**Parameters**

- **mu** (*array-like*) –

**Return type**

array-like

**unit\_variance\_derivative(mu)**

Get the derivative of the unit variance.

See superclass documentation.

**Parameters****mu** (array-like or float) –**Return type**

array-like

**variance(mu, dispersion=1, sample\_weight=1)**

Compute the variance function.

The variance of  $Y_i \sim \text{EDM}(\mu_i, \phi/s_i)$  is  $\text{var}(Y_i) = (\phi/s_i) * v(\mu_i)$ , with unit variance  $v(\mu)$  and weights  $s_i$ .**Parameters**

- **mu** (array-like, shape (n\_samples,)) – Predicted mean.
- **dispersion** (float, optional (default=1)) – Dispersion parameter  $\phi$ .
- **sample\_weight** (array-like, shape (n\_samples,)), optional (default=1))  
– Weights or exposure to which variance is inverse proportional.

**Return type**

array-like, shape (n\_samples,)

**variance\_derivative(mu, dispersion=1, sample\_weight=1)**

Compute the derivative of the variance with respect to mu.

The derivative of the variance is equal to  $(\phi/s_i) * v'(\mu_i)$ , where  $v(\mu)$  is the unit variance and  $s_i$  are weights.**Parameters**

- **mu** (array-like, shape (n\_samples,)) – Predicted mean.
- **dispersion** (float, optional (default=1)) – Dispersion parameter  $\phi$ .
- **sample\_weight** (array-like, shape (n\_samples,)), optional (default=1))  
– Weights or exposure to which variance is inverse proportional.

**Return type**

array-like, shape (n\_samples,)

**class glum.CloglogLink**Bases: [Link](#)The complementary log-log link function  $\log(-\log(-p))$ .**derivative(mu)**

Get the derivative of the cloglog link.

See superclass documentation.

**Parameters****mu** (array-like) –**Return type**

array-like

**inverse**(*lin\_pred*)

Get the inverse of the cloglog link.

See superclass documentation.

Note: since passing a very large value might result in an output of one, this function bounds the output to be between  $[50 \cdot \text{eps}, 1 - 50 \cdot \text{eps}]$ , where *eps* is floating point epsilon.

**Parameters**

**lin\_pred** (*array-like*) –

**Return type**

*array-like*

**inverse\_derivative**(*lin\_pred*)

Compute the derivative of the inverse link function  $h'(\text{lin\_pred})$ .

**Parameters**

**lin\_pred** (*array, shape (n\_samples,)*) – Usually the (fitted) linear predictor.

**inverse\_derivative2**(*lin\_pred*)

Compute second derivative of the inverse link function  $h''(\text{lin\_pred})$ .

**Parameters**

**lin\_pred** (*array, shape (n\_samples,)*) – Usually the (fitted) linear predictor.

**link**(*mu*)

Get the logit function of *mu*.

See superclass documentation.

**Parameters**

**mu** (*array-like*) –

**Return type**

*numpy.ndarray*

**class** `glum.ExponentialDispersionModel`

Bases: `object`

Base class for reproductive Exponential Dispersion Models (EDM).

The PDF of  $Y \sim \text{EDM}(\mu, \phi)$  is given by

$$\begin{aligned} p(y \mid \theta, \phi) &= c(y, \phi) \exp((\theta y - A(\theta))/\phi) \\ &= \tilde{c}(y, \phi) \exp(-d(y, \mu)/(2\phi)) \end{aligned}$$

with mean  $E(Y) = A'(\theta) = \mu$ , variance  $\text{var}(Y) = \phi \cdot v(\mu)$ , unit variance  $v(\mu)$  and unit deviance  $d(y, \mu)$ .

## 8.1 Properties

`lower_bound` `upper_bound` `include_lower_bound` `include_upper_bound`

**in\_y\_range**()

**Return type**

*ndarray*

**unit\_variance**()

**unit\_variance\_derivative()**

**variance()**

**Parameters**

**mu** (*ndarray*) –

**Return type**

*ndarray*

**variance\_derivative()**

**unit\_deviance()**

**unit\_deviance\_derivative()**

**deviance()**

**deviance\_derivative()**

**starting\_mu()**

**\_mu\_deviance\_derivative()**

**eta\_mu\_deviance()**

**Parameters**

- **link** (*Link*) –
- **factor** (*float*) –
- **cur\_eta** (*ndarray*) –
- **X\_dot\_d** (*ndarray*) –
- **y** (*ndarray*) –
- **sample\_weight** (*ndarray*) –

**gradient\_hessian()**

## References

[https://en.wikipedia.org/wiki/Exponential\\_dispersion\\_model](https://en.wikipedia.org/wiki/Exponential_dispersion_model).

**deviance**(*y*, *mu*, *sample\_weight=1*)

Compute the deviance.

The deviance is a weighted sum of the unit deviances,  $\sum_i s_i \times d(y_i, \mu_i)$ , where  $d(y, \mu)$  is the unit deviance and  $s$  are weights. In terms of the log likelihood, it is  $-2\phi \times [L(y, \mu, \phi/s) - L(y, y, \phi/s)]$ .

**Parameters**

- **y** (*array-like*, *shape (n\_samples,)*) – Target values.
- **mu** (*array-like*, *shape (n\_samples,)*) – Predicted mean.
- **sample\_weight** (*array-like*, *shape (n\_samples,)*, *optional (default=1)*) – Weights or exposure to which variance is inversely proportional.

**Return type**

*float*



**deviance\_derivative**(*y*, *mu*, *sample\_weight=1*)

Compute the derivative of the deviance with respect to  $\mu$ .

**Parameters**

- **y** (*array-like*, *shape* (*n\_samples*,)) – Target values.
- **mu** (*array-like*, *shape* (*n\_samples*,)) – Predicted mean.
- **sample\_weight** (*array-like*, *shape* (*n\_samples*,)) (*default=1*) – Weights or exposure to which variance is inverse proportional.

**Return type**

*array-like*, *shape* (*n\_samples*,)

**dispersion**(*y*, *mu*, *sample\_weight=None*, *ddof=1*, *method='pearson'*)

Estimate the dispersion parameter  $\phi$ .

**Parameters**

- **y** (*array-like*, *shape* (*n\_samples*,)) – Target values.
- **mu** (*array-like*, *shape* (*n\_samples*,)) – Predicted mean.
- **sample\_weight** (*array-like*, *shape* (*n\_samples*,)) (*optional* (*default=1*)) – Weights or exposure to which variance is inversely proportional.
- **ddof** (*int*, *optional* (*default=1*)) – Degrees of freedom consumed by the model for  $\mu$ .
- **{'pearson'}** (*method* =) – Whether to base the estimate on the Pearson residuals or the deviance.
- **'deviance'** – Whether to base the estimate on the Pearson residuals or the deviance.
- **(default='pearson')** (*optional*) – Whether to base the estimate on the Pearson residuals or the deviance.

**Return type**

*float*

**eta\_mu\_deviance**(*link*, *factor*, *cur\_eta*, *X\_dot\_d*, *y*, *sample\_weight*)

Compute  $\eta$ ,  $\mu$  and the deviance.

Compute: \* the linear predictor,  $\eta$ , as  $\text{cur\_eta} + \text{factor} * \text{X\_dot\_d}$ ; \* the link-function-transformed prediction,  $\mu$ ; \* the deviance.

**Returns**

- *numpy.ndarray*, *shape* (*X.shape*[0],) – The linear predictor,  $\eta$ .
- *numpy.ndarray*, *shape* (*X.shape*[0],) – The link-function-transformed prediction,  $\mu$ .
- *float* – The deviance.

**Parameters**

- **link** (*Link*) –
- **factor** (*float*) –
- **cur\_eta** (*ndarray*) –
- **X\_dot\_d** (*ndarray*) –
- **y** (*ndarray*) –

- **sample\_weight** (*ndarray*) –

**in\_y\_range**(*x*)

Return True if *x* is in the valid range of the EDM.

**Parameters**

**x** (*array-like*, *shape* (*n\_samples*,)) – Target values.

**Return type**

*np.ndarray*

**abstract property include\_lower\_bound: bool**

Return whether *lower\_bound* is allowed as a value of *y*.

**abstract property include\_upper\_bound: bool**

Return whether *upper\_bound* is allowed as a value of *y*.

**abstract property lower\_bound: float**

Get the lower bound of values for the EDM.

**rowwise\_gradient\_hessian**(*link*, *coef*, *dispersion*, *X*, *y*, *sample\_weight*, *eta*, *mu*, *offset=None*)

Compute the gradient and negative Hessian of the log likelihood row-wise.

**Returns**

- *numpy.ndarray*, *shape* (*X.shape*[0],) – The gradient of the log likelihood, row-wise.
- *numpy.ndarray*, *shape* (*X.shape*[0],) – The negative Hessian of the log likelihood, row-wise.

**Parameters**

- **link** (*Link*) –
- **coef** (*ndarray*) –
- **X** (*MatrixBase* | *StandardizedMatrix*) –
- **y** (*ndarray*) –
- **sample\_weight** (*ndarray*) –
- **eta** (*ndarray*) –
- **mu** (*ndarray*) –
- **offset** (*ndarray* | *None*) –

**abstract unit\_deviance**(*y*, *mu*)

Compute the unit deviance.

In terms of the log likelihood  $L$ , the unit deviance is  $-2\phi \times [L(y, \mu, \phi) - L(y, y, \phi)]$ .

**Parameters**

- **y** (*array-like*, *shape* (*n\_samples*,)) – Target values.
- **mu** (*array-like*, *shape* (*n\_samples*,)) – Predicted mean.

**unit\_deviance\_derivative**(*y*, *mu*)

Compute the derivative of the unit deviance with respect to *mu*.

The derivative of the unit deviance is given by  $-2 \times (y - \mu)/v(\mu)$ , where  $v(\mu)$  is the unit variance.

**Parameters**

- **y** (*array-like*, *shape* (*n\_samples*,)) – Target values.
- **mu** (*array-like*, *shape* (*n\_samples*,)) – Predicted mean.

**Return type**array-like, shape (*n\_samples*,)**abstract unit\_variance(mu)**

Compute the unit variance function.

The unit variance  $v(\mu)$  determines the variance as a function of the mean  $\mu$  by  $\text{var}(y_i) = (\phi/s_i) \times v(\mu_i)$ . It can also be derived from the unit deviance  $d(y, \mu)$  as

$$v(\mu) = \frac{2}{\frac{\partial^2 d(y, \mu)}{\partial \mu^2}} \Big|_{y=\mu}.$$

See also [variance\(\)](#).**Parameters****mu** (*array-like*, *shape* (*n\_samples*,)) – Predicted mean.**abstract unit\_variance\_derivative(mu)**

Compute the derivative of the unit variance with respect to mu.

Return  $v'(\mu)$ .**Parameters****mu** (*array-like*, *shape* (*n\_samples*,)) – Predicted mean.**abstract property upper\_bound: float**

Get the upper bound of values for the EDM.

**variance(mu, dispersion=1, sample\_weight=1)**

Compute the variance function.

The variance of  $Y_i \sim \text{EDM}(\mu_i, \phi/s_i)$  is  $\text{var}(Y_i) = (\phi/s_i) * v(\mu_i)$ , with unit variance  $v(\mu)$  and weights  $s_i$ .

**Parameters**

- **mu** (*array-like*, *shape* (*n\_samples*,)) – Predicted mean.
- **dispersion** (*float*, *optional* (*default=1*)) – Dispersion parameter  $\phi$ .
- **sample\_weight** (*array-like*, *shape* (*n\_samples*,), *optional* (*default=1*)) – Weights or exposure to which variance is inverse proportional.

**Return type**array-like, shape (*n\_samples*,)**variance\_derivative(mu, dispersion=1, sample\_weight=1)**

Compute the derivative of the variance with respect to mu.

The derivative of the variance is equal to  $(\phi/s_i) * v'(\mu_i)$ , where  $v(\mu)$  is the unit variance and  $s_i$  are weights.

**Parameters**

- **mu** (*array-like*, *shape* (*n\_samples*,)) – Predicted mean.
- **dispersion** (*float*, *optional* (*default=1*)) – Dispersion parameter  $\phi$ .
- **sample\_weight** (*array-like*, *shape* (*n\_samples*,), *optional* (*default=1*)) – Weights or exposure to which variance is inverse proportional.

**Return type**array-like, shape (*n\_samples*,)

**class glum.GammaDistribution**Bases: *TweedieDistribution*

Class for the Gamma distribution.

**deviance**(*y*, *mu*, *sample\_weight=None*)

Compute the deviance.

**Parameters**

- **y** (*array-like*, *shape* (*n\_samples*,)) – Target values.
- **mu** (*array-like*, *shape* (*n\_samples*,)) – Predicted mean.
- **sample\_weight** (*array-like*, *shape* (*n\_samples*,), *optional* (*default=1*)) – Sample weights.

**Return type**

float

**deviance\_derivative**(*y*, *mu*, *sample\_weight=1*)Compute the derivative of the deviance with respect to *mu*.**Parameters**

- **y** (*array-like*, *shape* (*n\_samples*,)) – Target values.
- **mu** (*array-like*, *shape* (*n\_samples*,)) – Predicted mean.
- **sample\_weight** (*array-like*, *shape* (*n\_samples*,) (*default=1*)) – Weights or exposure to which variance is inverse proportional.

**Return type***array-like*, *shape* (*n\_samples*,)**dispersion**(*y*, *mu*, *sample\_weight=None*, *ddof=1*, *method='pearson'*)Estimate the dispersion parameter  $\phi$ .**Parameters**

- **y** (*array-like*, *shape* (*n\_samples*,)) – Target values.
- **mu** (*array-like*, *shape* (*n\_samples*,)) – Predicted mean.
- **sample\_weight** (*array-like*, *shape* (*n\_samples*,), *optional* (*default=None*)) – Weights or exposure to which variance is inversely proportional.
- **ddof** (*int*, *optional* (*default=1*)) – Degrees of freedom consumed by the model for *mu*.
- **{'pearson'}** (*method* =) – Whether to base the estimate on the Pearson residuals or the deviance.
- **'deviance'** – Whether to base the estimate on the Pearson residuals or the deviance.
- **(default='pearson')** (*optional*) – Whether to base the estimate on the Pearson residuals or the deviance.

**Return type**

float

**eta\_mu\_deviance**(*link, factor, cur\_eta, X\_dot\_d, y, sample\_weight*)

Compute eta, mu and the deviance.

Compute: \* the linear predictor, eta, as `cur_eta + factor * X_dot_d`; \* the link-function-transformed prediction, mu; \* the deviance.

**Returns**

- *numpy.ndarray, shape (X.shape[0],)* – The linear predictor, eta.
- *numpy.ndarray, shape (X.shape[0],)* – The link-function-transformed prediction, mu.
- *float* – The deviance.

**Parameters**

- **link** ([Link](#)) –
- **factor** (*float*) –
- **cur\_eta** (*ndarray*) –
- **X\_dot\_d** (*ndarray*) –
- **y** (*ndarray*) –
- **sample\_weight** (*ndarray*) –

**in\_y\_range**(*x*)

Return True if *x* is in the valid range of the EDM.

**Parameters**

**x** (*array-like, shape (n\_samples,)*) – Target values.

**Return type**

*np.ndarray*

**property include\_lower\_bound: bool**

Return whether `lower_bound` is allowed as a value of *y*.

**log\_likelihood**(*y, mu, sample\_weight=None, dispersion=None*)

Compute the log likelihood.

For  $1 < \text{power} < 2$ , we use the series approximation by Dunn and Smyth (2005) to compute the normalization term.

**Parameters**

- **y** (*array-like, shape (n\_samples,)*) – Target values.
- **mu** (*array-like, shape (n\_samples,)*) – Predicted mean.
- **sample\_weight** (*array-like, shape (n\_samples,)*, *optional (default=1)*) – Sample weights.
- **dispersion** (*float, optional (default=None)*) – Dispersion parameter  $\phi$ . Estimated if None.

**Return type**

*float*

**property lower\_bound: float**

Return the lowest value of *y* allowed.

**property power: float**

Return the Tweedie power parameter.

**rowwise\_gradient\_hessian**(*link, coef, dispersion, X, y, sample\_weight, eta, mu, offset=None*)

Compute the gradient and negative Hessian of the log likelihood row-wise.

**Returns**

- *numpy.ndarray, shape (X.shape[0],)* – The gradient of the log likelihood, row-wise.
- *numpy.ndarray, shape (X.shape[0],)* – The negative Hessian of the log likelihood, row-wise.

**Parameters**

- **link** ([Link](#)) –
- **coef** (*ndarray*) –
- **X** (*MatrixBase | StandardizedMatrix*) –
- **y** (*ndarray*) –
- **sample\_weight** (*ndarray*) –
- **eta** (*ndarray*) –
- **mu** (*ndarray*) –
- **offset** (*ndarray | None*) –

**unit\_deviance**(*y, mu*)

Get the deviance of each observation.

**unit\_deviance\_derivative**(*y, mu*)

Compute the derivative of the unit deviance with respect to *mu*.

The derivative of the unit deviance is given by  $-2 \times (y - \mu)/v(\mu)$ , where  $v(\mu)$  is the unit variance.

**Parameters**

- **y** (*array-like, shape (n\_samples,)*) – Target values.
- **mu** (*array-like, shape (n\_samples,)*) – Predicted mean.

**Return type**

*array-like, shape (n\_samples,)*

**unit\_variance**(*mu*)

Compute the unit variance of a Tweedie distribution  $v(\mu) = \mu^{\text{power}}$ .

**Parameters**

**mu** (*array-like, shape (n\_samples,)*) – Predicted mean.

**Return type**

*numpy.ndarray, shape (n\_samples,)*

**unit\_variance\_derivative**(*mu*)

Compute the derivative of the unit variance of a Tweedie distribution.

Equation:  $v(\mu) = p \times \mu^{(p-1)}$ .

**Parameters**

**mu** (*array-like, shape (n\_samples,)*) – Predicted mean.

**Return type**

numpy.ndarray, shape (n\_samples,)

**variance**(*mu*, *dispersion=1*, *sample\_weight=1*)

Compute the variance function.

The variance of  $Y_i \sim \text{EDM}(\mu_i, \phi/s_i)$  is  $\text{var}(Y_i) = (\phi/s_i) * v(\mu_i)$ , with unit variance  $v(\mu)$  and weights  $s_i$ .

**Parameters**

- **mu** (array-like, shape (n\_samples,)) – Predicted mean.
- **dispersion** (float, optional (default=1)) – Dispersion parameter  $\phi$ .
- **sample\_weight** (array-like, shape (n\_samples,)), optional (default=1) – Weights or exposure to which variance is inverse proportional.

**Return type**

array-like, shape (n\_samples,)

**variance\_derivative**(*mu*, *dispersion=1*, *sample\_weight=1*)

Compute the derivative of the variance with respect to  $\mu$ .

The derivative of the variance is equal to  $(\phi/s_i) * v'(\mu_i)$ , where  $v(\mu)$  is the unit variance and  $s_i$  are weights.

**Parameters**

- **mu** (array-like, shape (n\_samples,)) – Predicted mean.
- **dispersion** (float, optional (default=1)) – Dispersion parameter  $\phi$ .
- **sample\_weight** (array-like, shape (n\_samples,)), optional (default=1) – Weights or exposure to which variance is inverse proportional.

**Return type**

array-like, shape (n\_samples,)

**class** glum.GeneralizedHyperbolicSecant

Bases: [ExponentialDispersionModel](#)

A class for the Generalized Hyperbolic Secant (GHS) distribution.

The GHS distribution is for targets  $y$  in  $(-\infty, +\infty)$ .

**deviance**(*y*, *mu*, *sample\_weight=1*)

Compute the deviance.

The deviance is a weighted sum of the unit deviances,  $\sum_i s_i \times d(y_i, \mu_i)$ , where  $d(y, \mu)$  is the unit deviance and  $s$  are weights. In terms of the log likelihood, it is  $-2\phi \times [L(y, \mu, \phi/s) - L(y, y, \phi/s)]$ .

**Parameters**

- **y** (array-like, shape (n\_samples,)) – Target values.
- **mu** (array-like, shape (n\_samples,)) – Predicted mean.
- **sample\_weight** (array-like, shape (n\_samples,)), optional (default=1) – Weights or exposure to which variance is inversely proportional.

**Return type**

float

**deviance\_derivative**(*y*, *mu*, *sample\_weight=1*)

Compute the derivative of the deviance with respect to  $\mu$ .

**Parameters**

- **y** (*array-like*, *shape* (*n\_samples*,)) – Target values.
- **mu** (*array-like*, *shape* (*n\_samples*,)) – Predicted mean.
- **sample\_weight** (*array-like*, *shape* (*n\_samples*,) (*default=1*)) – Weights or exposure to which variance is inverse proportional.

**Return type**

*array-like*, *shape* (*n\_samples*,)

**dispersion**(*y*, *mu*, *sample\_weight=None*, *ddof=1*, *method='pearson'*)

Estimate the dispersion parameter  $\phi$ .

**Parameters**

- **y** (*array-like*, *shape* (*n\_samples*,)) – Target values.
- **mu** (*array-like*, *shape* (*n\_samples*,)) – Predicted mean.
- **sample\_weight** (*array-like*, *shape* (*n\_samples*,), *optional* (*default=1*)) – Weights or exposure to which variance is inversely proportional.
- **ddof** (*int*, *optional* (*default=1*)) – Degrees of freedom consumed by the model for *mu*.
- **{'pearson'}** (*method* =) – Whether to base the estimate on the Pearson residuals or the deviance.
- **'deviance'** – Whether to base the estimate on the Pearson residuals or the deviance.
- **(default='pearson')** (*optional*) – Whether to base the estimate on the Pearson residuals or the deviance.

**Return type**

*float*

**eta\_mu\_deviance**(*link*, *factor*, *cur\_eta*, *X\_dot\_d*, *y*, *sample\_weight*)

Compute *eta*, *mu* and the deviance.

Compute: \* the linear predictor, *eta*, as *cur\_eta* + *factor* \* *X\_dot\_d*; \* the link-function-transformed prediction, *mu*; \* the deviance.

**Returns**

- *numpy.ndarray*, *shape* (*X.shape*[0],) – The linear predictor, *eta*.
- *numpy.ndarray*, *shape* (*X.shape*[0],) – The link-function-transformed prediction, *mu*.
- *float* – The deviance.

**Parameters**

- **link** ([Link](#)) –
- **factor** (*float*) –
- **cur\_eta** (*ndarray*) –
- **X\_dot\_d** (*ndarray*) –
- **y** (*ndarray*) –
- **sample\_weight** (*ndarray*) –



**in\_y\_range(x)**

Return True if *x* is in the valid range of the EDM.

**Parameters**

**x** (*array-like*, *shape* (*n\_samples*,)) – Target values.

**Return type**

np.ndarray

**rowwise\_gradient\_hessian(link, coef, dispersion, X, y, sample\_weight, eta, mu, offset=None)**

Compute the gradient and negative Hessian of the log likelihood row-wise.

**Returns**

- *numpy.ndarray*, *shape* (*X.shape*[0],) – The gradient of the log likelihood, row-wise.
- *numpy.ndarray*, *shape* (*X.shape*[0],) – The negative Hessian of the log likelihood, row-wise.

**Parameters**

- **link** ([Link](#)) –
- **coef** (*ndarray*) –
- **X** (*MatrixBase* | *StandardizedMatrix*) –
- **y** (*ndarray*) –
- **sample\_weight** (*ndarray*) –
- **eta** (*ndarray*) –
- **mu** (*ndarray*) –
- **offset** (*ndarray* | *None*) –

**unit\_deviance(y, mu)**

Get the unit-level deviance.

See superclass documentation.

**Parameters**

- **y** (*array-like*) –
- **mu** (*array-like*) –

**Return type**

array-like

**unit\_deviance\_derivative(y, mu)**

Compute the derivative of the unit deviance with respect to **mu**.

The derivative of the unit deviance is given by  $-2 \times (y - \mu)/v(\mu)$ , where  $v(\mu)$  is the unit variance.

**Parameters**

- **y** (*array-like*, *shape* (*n\_samples*,)) – Target values.
- **mu** (*array-like*, *shape* (*n\_samples*,)) – Predicted mean.

**Return type**

array-like, *shape* (*n\_samples*,)

**unit\_variance(mu)**

Get the unit-level expected variance.

See superclass documentation.

**Parameters**

**mu** (*array-like or float*) –

**Return type**

array-like

**unit\_variance\_derivative(mu)**

Get the derivative of the unit variance.

See superclass documentation.

**Parameters**

**mu** (*array-like or float*) –

**Return type**

array-like

**variance(mu, dispersion=1, sample\_weight=1)**

Compute the variance function.

The variance of  $Y_i \sim \text{EDM}(\mu_i, \phi/s_i)$  is  $\text{var}(Y_i) = (\phi/s_i) * v(\mu_i)$ , with unit variance  $v(\mu)$  and weights  $s_i$ .

**Parameters**

- **mu** (*array-like, shape (n\_samples,)*) – Predicted mean.
- **dispersion** (*float, optional (default=1)*) – Dispersion parameter  $\phi$ .
- **sample\_weight** (*array-like, shape (n\_samples,), optional (default=1)*) – Weights or exposure to which variance is inverse proportional.

**Return type**

array-like, shape (n\_samples,)

**variance\_derivative(mu, dispersion=1, sample\_weight=1)**

Compute the derivative of the variance with respect to mu.

The derivative of the variance is equal to  $(\phi/s_i) * v'(\mu_i)$ , where  $v(\mu)$  is the unit variance and  $s_i$  are weights.

**Parameters**

- **mu** (*array-like, shape (n\_samples,)*) – Predicted mean.
- **dispersion** (*float, optional (default=1)*) – Dispersion parameter  $\phi$ .
- **sample\_weight** (*array-like, shape (n\_samples,), optional (default=1)*) – Weights or exposure to which variance is inverse proportional.

**Return type**

array-like, shape (n\_samples,)

```
class glum.GeneralizedLinearRegressor(alpha=None, l1_ratio=0, P1='identity', P2='identity',
                                     fit_intercept=True, family='normal', link='auto', solver='auto',
                                     max_iter=100, gradient_tol=None, step_size_tol=None,
                                     hessian_approx=0.0, warm_start=False, alpha_search=False,
                                     alphas=None, n_alphas=100, min_alpha_ratio=None,
                                     min_alpha=None, start_params=None, selection='cyclic',
                                     random_state=None, copy_X=None, check_input=True,
                                     verbose=0, scale_predictors=False, lower_bounds=None,
                                     upper_bounds=None, A_ineq=None, b_ineq=None,
                                     force_all_finite=True, drop_first=False, robust=True,
                                     expected_information=False)
```

Bases: GeneralizedLinearRegressorBase

Regression via a Generalized Linear Model (GLM) with penalties.

GLMs based on a reproductive Exponential Dispersion Model (EDM) aimed at fitting and predicting the mean of the target  $y$  as  $\mu=h(X*w)$ . Therefore, the fit minimizes the following objective function with combined L1 and L2 priors as regularizer:

```
1/(2*sum(s)) * deviance(y, h(X*w); s)
+ alpha * l1_ratio * ||P1*w||_1
+ 1/2 * alpha * (1 - l1_ratio) * w*P2*w
```

with inverse link function  $h$  and  $s$ =sample\_weight. Note that, for  $\alpha=0$  the unregularized GLM is recovered. This is not the default behavior (see `alpha` parameter description for details). Additionally, for `sample_weight=None`, one has  $s_i=1$  and  $\text{sum}(s)=n_{\text{samples}}$ . For  $P1=P2='identity'$ , the penalty is the elastic net:

```
alpha * l1_ratio * ||w||_1 + 1/2 * alpha * (1 - l1_ratio) * ||w||_2^2.
```

If you are interested in controlling the L1 and L2 penalties separately, keep in mind that this is equivalent to:

```
a * L1 + b * L2,
```

where:

```
alpha = a + b and l1_ratio = a / (a + b).
```

The parameter `l1_ratio` corresponds to `alpha` in the R package `glmnet`, while `alpha` corresponds to the `lambda` parameter in `glmnet`. Specifically, `l1_ratio = 1` is the lasso penalty.

Read more in [background](#).

### Parameters

- **alpha** (*{float, array-like}, optional (default=None)*) – Constant that multiplies the penalty terms and thus determines the regularization strength. If `alpha_search` is `False` (the default), then `alpha` must be a scalar or `None` (equivalent to `alpha=1.0`). If `alpha_search` is `True`, then `alpha` must be an iterable or `None`. See `alpha_search` to find how the regularization path is set if `alpha` is `None`. See the notes for the exact mathematical meaning of this parameter. `alpha = 0` is equivalent to unpenalized GLMs. In this case, the design matrix  $X$  must have full column rank (no collinearities).
- **l1\_ratio** (*float, optional (default=0)*) – The elastic net mixing parameter, with  $0 \leq \text{l1\_ratio} \leq 1$ . For `l1_ratio = 0`, the penalty is an L2 penalty. For `l1_ratio = 1`, it is an L1 penalty. For  $0 < \text{l1\_ratio} < 1$ , the penalty is a combination of L1 and L2.

- **P1** (`{'identity', array-like, None}`, `shape (n_features,)`, optional (`default='identity'`)) – This array controls the strength of the regularization for each coefficient independently. A high value will lead to higher regularization while a value of zero will remove the regularization on this parameter. Note that `n_features = X.shape[1]`. If `X` is a pandas DataFrame with a categorical dtype and `P1` has the same size as the number of columns, the penalty of the categorical column will be applied to all the levels of the categorical.
- **P2** (`{'identity', array-like, sparse matrix, None}`, `shape (n_features,)` or `(n_features, n_features)`, optional (`default='identity'`)) – With this option, you can set the `P2` matrix in the L2 penalty  $w * P2 * w$ . This gives a fine control over this penalty (Tikhonov regularization). A 2d array is directly used as the square matrix `P2`. A 1d array is interpreted as diagonal (square) matrix. The default `'identity'` and `None` set the identity matrix, which gives the usual squared L2-norm. If you just want to exclude certain coefficients, pass a 1d array filled with 1 and 0 for the coefficients to be excluded. Note that `P2` must be positive semi-definite. If `X` is a pandas DataFrame with a categorical dtype and `P2` has the same size as the number of columns, the penalty of the categorical column will be applied to all the levels of the categorical. Note that if `P2` is two-dimensional, its size needs to be of the same length as the expanded `X` matrix.
- **fit\_intercept** (`bool`, optional (`default=True`)) – Specifies if a constant (a.k.a. bias or intercept) should be added to the linear predictor ( $X * \text{coef} + \text{intercept}$ ).
- **family** (`str` or `ExponentialDispersionModel`, optional (`default='normal'`)) – The distributional assumption of the GLM, i.e. the loss function to minimize. If a string, one of: `'binomial'`, `'gamma'`, `'gaussian'`, `'inverse.gaussian'`, `'normal'`, `'poisson'`, `'tweedie'` or `'negative.binomial'`. Note that `'tweedie'` sets the power of the Tweedie distribution to 1.5; to use another value, specify it in parentheses (e.g., `'tweedie (1.5)'`). The same applies for `'negative.binomial'` and theta parameter.
- **link** (`{'auto', 'identity', 'log', 'logit', 'cloglog'}` or `Link`, optional (`default='auto'`)) – The link function of the GLM, i.e. mapping from linear predictor ( $X * \text{coef}$ ) to expectation ( $\mu$ ). Option `'auto'` sets the link depending on the chosen family as follows:
  - `'identity'` for family `'normal'`
  - `'log'` for families `'poisson'`, `'gamma'`, `'inverse.gaussian'` and `'negative.binomial'`.
  - `'logit'` for family `'binomial'`
- **solver** (`{'auto', 'irls-cd', 'irls-ls', 'lbfgs', 'trust-constr'}`, optional (`default='auto'`)) – Algorithm to use in the optimization problem:
  - `'auto'`: `'irls-ls'` if `l1_ratio` is zero and `'irls-cd'` otherwise.
  - `'irls-cd'`: Iteratively reweighted least squares with a coordinate descent inner solver. This can deal with L1 as well as L2 penalties. Note that in order to avoid unnecessary memory duplication of `X` in the `fit` method, `X` should be directly passed as a Fortran-contiguous Numpy array or sparse CSC matrix.
  - `'irls-ls'`: Iteratively reweighted least squares with a least squares inner solver. This algorithm cannot deal with L1 penalties.
  - `'lbfgs'`: Scipy's L-BFGS-B optimizer. It cannot deal with L1 penalties.
  - `'trust-constr'`: Calls `scipy.optimize.minimize(method='trust-constr')`. It cannot deal with L1 penalties. This solver can optimize problems with inequality constraints, passed via `A_ineq` and `b_ineq`. It will be selected automatically when inequality

constraints are set and `solver='auto'`. Note that using this method can lead to significantly increased runtimes by a factor of ten or higher.

- **max\_iter** (*int, optional (default=100)*) – The maximal number of iterations for solver algorithms.
- **gradient\_tol** (*float, optional (default=None)*) – Stopping criterion. If `None`, solver-specific defaults will be used. The default value for most solvers is `1e-4`, except for `'trust-constr'`, which requires more conservative convergence settings and has a default value of `1e-8`.

For the IRLS-LS, L-BFGS and trust-constr solvers, the iteration will stop when  $\max\{|g_i|, i = 1, \dots, n\} \leq \text{tol}$ , where  $g_i$  is the  $i$ -th component of the gradient (derivative) of the objective function. For the CD solver, convergence is reached when  $\text{sum}_i(|\text{minimum norm of } g_i|)$ , where  $g_i$  is the subgradient of the objective and the minimum norm of  $g_i$  is the element of the subgradient with the smallest L2 norm.

If you wish to only use a step-size tolerance, set `gradient_tol` to a very small number.

- **step\_size\_tol** (*float, optional (default=None)*) – Alternative stopping criterion. For the IRLS-LS and IRLS-CD solvers, the iteration will stop when the L2 norm of the step size is less than `step_size_tol`. This stopping criterion is disabled when `step_size_tol` is `None`.
- **hessian\_approx** (*float, optional (default=0.0)*) – The threshold below which data matrix rows will be ignored for updating the Hessian. See the algorithm documentation for the IRLS algorithm for further details.
- **warm\_start** (*bool, optional (default=False)*) – Whether to reuse the solution of the previous call to `fit` as initialization for `coef_` and `intercept_` (supersedes `start_params`). If `False` or if the attribute `coef_` does not exist (first call to `fit`), `start_params` sets the start values for `coef_` and `intercept_`.
- **alpha\_search** (*bool, optional (default=False)*) – Whether to search along the regularization path for the best alpha. When set to `True`, `alpha` should either be `None` or an iterable. To determine the regularization path, the following sequence is used:
  1. If `alpha` is an iterable, use it directly. All other parameters governing the regularization path are ignored.
  2. If `min_alpha` is set, create a path from `min_alpha` to the lowest alpha such that all coefficients are zero.
  3. If `min_alpha_ratio` is set, create a path where the ratio of `min_alpha` / `max_alpha` = `min_alpha_ratio`.
  4. If none of the above parameters are set, use a `min_alpha_ratio` of `1e-6`.
- **alphas** (DEPRECATED. Use `alpha` instead.) –
- **n\_alphas** (*int, optional (default=100)*) – Number of alphas along the regularization path
- **min\_alpha\_ratio** (*float, optional (default=None)*) – Length of the path. `min_alpha_ratio=1e-6` means that `min_alpha` / `max_alpha` = `1e-6`. If `None`, `1e-6` is used.
- **min\_alpha** (*float, optional (default=None)*) – Minimum alpha to estimate the model with. The grid will then be created over `[max_alpha, min_alpha]`.
- **start\_params** (*array-like, shape (n\_features\*,), optional (default=None)*) – Relevant only if `warm_start` is `False` or if `fit` is called for

the first time (so that `self.coef_` does not exist yet). If `None`, all coefficients are set to zero and the start value for the intercept is the weighted average of `y` (If `fit_intercept` is `True`). If an array, used directly as start values; if `fit_intercept` is `True`, its first element is assumed to be the start value for the `intercept_`. Note that `n_features* = X.shape[1] + fit_intercept`, i.e. it includes the intercept.

- **selection** (*str, optional (default='cyclic')*) – For the CD solver 'cd', the coordinates (features) can be updated in either cyclic or random order. If set to 'random', a random coefficient is updated every iteration rather than looping over features sequentially in the same order, which often leads to significantly faster convergence, especially when `gradient_tol` is higher than `1e-4`.
- **random\_state** (*int or RandomState, optional (default=None)*) – The seed of the pseudo random number generator that selects a random feature to be updated for the CD solver. If an integer, `random_state` is the seed used by the random number generator; if a `RandomState` instance, `random_state` is the random number generator; if `None`, the random number generator is the `RandomState` instance used by `np.random`. Used when `selection` is 'random'.
- **copy\_X** (*bool, optional (default=None)*) – Whether to copy `X`. Since `X` is never modified by `GeneralizedLinearRegressor`, this is unlikely to be needed; this option exists mainly for compatibility with other scikit-learn estimators. If `False`, `X` will not be copied and there will be an error if you pass an `X` in the wrong format, such as providing integer `X` and float `y`. If `None`, `X` will not be copied unless it is in the wrong format.
- **check\_input** (*bool, optional (default=True)*) – Whether to bypass several checks on input: `y` values in range of family, `sample_weight` non-negative, P2 positive semi-definite. Don't use this parameter unless you know what you are doing.
- **verbose** (*int, optional (default=0)*) – For the IRLS solver, any positive number will result in a pretty progress bar showing convergence. This features requires having the `tqdm` package installed. For the L-BFGS and 'trust-constr' solvers, set `verbose` to any positive number for verbosity.
- **scale\_predictors** (*bool, optional (default=False)*) – If `True`, estimate a scaled model where all predictors have a standard deviation of 1. This can result in better estimates if predictors are on very different scales (for example, centimeters and kilometers).

Advanced developer note: Internally, predictors are always rescaled for computational reasons, but this only affects results if `scale_predictors` is `True`.

- **lower\_bounds** (*array-like, shape (n\_features,), optional (default=None)*) – Set a lower bound for the coefficients. Setting bounds forces the use of the coordinate descent solver ('irls-cd').
- **upper\_bounds** (*array-like, shape=(n\_features,), optional (default=None)*) – See `lower_bounds`.
- **A\_ineq** (*array-like, shape=(n\_constraints, n\_features), optional (default=None)*) – Constraint matrix for linear inequality constraints of the form  $A_{\text{ineq}} w \leq b_{\text{ineq}}$ . Setting inequality constraints forces the use of the local gradient-based solver 'trust-constr', which may increase runtime significantly. Note that the constraints only apply to coefficients related to features in `X`. If you want to constrain the intercept, add it to the feature matrix `X` manually and set `fit_intercept=False`.
- **b\_ineq** (*array-like, shape=(n\_constraints,), optional (default=None)*) – Constraint vector for linear inequality constraints of the form  $A_{\text{ineq}} w \leq b_{\text{ineq}}$ . Refer to the documentation of `A_ineq` for details.

- **drop\_first** (*bool*, *optional* (*default* = *False*)) – If True, drop the first column when encoding categorical variables. Set this to True when  $\alpha=0$  and `solver='auto'` to prevent an error due to a singular feature matrix.
- **robust** (*bool*, *optional* (*default* = *False*)) – If true, then robust standard errors are computed by default.
- **expected\_information** (*bool*, *optional* (*default* = *False*)) – If true, then the expected information matrix is computed by default. Only relevant when computing robust standard errors.
- **force\_all\_finite** (*bool*) –

**coef\_**

Estimated coefficients for the linear predictor ( $X \cdot \text{coef\_} + \text{intercept\_}$ ) in the GLM.

**Type**

numpy.array, shape (n\_features,)

**intercept\_**

Intercept (a.k.a. bias) added to linear predictor.

**Type**

float

**n\_iter\_**

Actual number of iterations used in solver.

**Type**

int

**Notes**

The fit itself does not need outcomes to be from an EDM, but only assumes the first two moments to be  $\mu_i \equiv E(y_i) = h(x_i'w)$  and  $\text{var}(y_i) = (\phi/s_i)v(\mu_i)$ . The unit variance function  $v(\mu_i)$  is a property of and given by the specific EDM; see [background](#).

The parameters  $w$  (`coef_` and `intercept_`) are estimated by minimizing the deviance plus penalty term, which is equivalent to (penalized) maximum likelihood estimation.

For  $\alpha > 0$ , the feature matrix  $X$  should be standardized in order to penalize features equally strong. Call `sklearn.preprocessing.StandardScaler` before calling `fit`.

If the target  $y$  is a ratio, appropriate sample weights  $s$  should be provided. As an example, consider Poisson distributed counts  $z$  (integers) and weights  $s = \text{exposure}$  (time, money, persons years, ...). Then you fit  $y = z/s$ , i.e. `GeneralizedLinearModel(family='poisson').fit(X, y, sample_weight=s)`. The weights are necessary for the right (finite sample) mean. Consider  $\bar{y} = \sum_i s_i y_i / \sum_i s_i$ : in this case, one might say that  $y$  follows a 'scaled' Poisson distribution. The same holds for other distributions.

## References

### For the coordinate descent implementation:

- Guo-Xun Yuan, Chia-Hua Ho, Chih-Jen Lin An Improved GLMNET for L1-regularized Logistic Regression, Journal of Machine Learning Research 13 (2012) 1999-2030 [https://www.csie.ntu.edu.tw/~cjlin/papers/l1\\_glmnet/long-glmnet.pdf](https://www.csie.ntu.edu.tw/~cjlin/papers/l1_glmnet/long-glmnet.pdf)

### `aic(X, y, sample_weight=None)`

Akaike's information criteria. Computed as:  $-2 \log \hat{\mathcal{L}} + 2\hat{k}$  where  $\hat{\mathcal{L}}$  is the maximum likelihood estimate of the model, and  $\hat{k}$  is the effective number of parameters. See `_compute_information_criteria` for more information on the computation of  $\hat{k}$ .

#### Parameters

- **X** (*array-like, sparse matrix*}, *shape (n\_samples, n\_features)*) – Same data as used in 'fit'
- **y** (*array-like*, *shape (n\_samples,)*) – Same data as used in 'fit'
- **sample\_weight** (*array-like, shape (n\_samples,)*, *optional (default=None)*) – Same data as used in 'fit'

### `aicc(X, y, sample_weight=None)`

Second-order Akaike's information criteria (or small sample AIC). Computed as:  $-2 \log \hat{\mathcal{L}} + 2\hat{k} + \frac{2k(k+1)}{n-k-1}$  where  $\hat{\mathcal{L}}$  is the maximum likelihood estimate of the model,  $n$  is the number of training instances, and  $\hat{k}$  is the effective number of parameters. See `_compute_information_criteria` for more information on the computation of  $\hat{k}$ .

#### Parameters

- **X** (*array-like, sparse matrix*}, *shape (n\_samples, n\_features)*) – Same data as used in 'fit'
- **y** (*array-like*, *shape (n\_samples,)*) – Same data as used in 'fit'
- **sample\_weight** (*array-like, shape (n\_samples,)*, *optional (default=None)*) – Same data as used in 'fit'

### `bic(X, y, sample_weight=None)`

Bayesian information criterion. Computed as:  $-2 \log \hat{\mathcal{L}} + k \log(n)$  where  $\hat{\mathcal{L}}$  is the maximum likelihood estimate of the model,  $n$  is the number of training instances, and  $\hat{k}$  is the effective number of parameters. See `_compute_information_criteria` for more information on the computation of  $\hat{k}$ .

#### Parameters

- **X** (*array-like, sparse matrix*}, *shape (n\_samples, n\_features)*) – Same data as used in 'fit'
- **y** (*array-like*, *shape (n\_samples,)*) – Same data as used in 'fit'
- **sample\_weight** (*array-like, shape (n\_samples,)*, *optional (default=None)*) – Same data as used in 'fit'

### `coef_table(confidence_level=0.95, X=None, y=None, mu=None, offset=None, sample_weight=None, dispersion=None, robust=None, clusters=None, expected_information=None)`

Get a table of of the regression coefficients.

Includes coefficient estimates, standard errors, t-values, p-values and confidence intervals.

#### Parameters



- **confidence\_level** (*float, optional, default=0.95*) – The confidence level for the confidence intervals.
- **X** (*{array-like, sparse matrix}, shape (n\_samples, n\_features), optional*) – Training data. Can be omitted if a covariance matrix has already been computed or if standard errors, etc. are not desired.
- **y** (*array-like, shape (n\_samples,), optional*) – Target values. Can be omitted if a covariance matrix has already been computed.
- **mu** (*array-like, optional, default=None*) – Array with predictions. Estimated if absent.
- **offset** (*array-like, optional, default=None*) – Array with additive offsets.
- **sample\_weight** (*array-like, shape (n\_samples,), optional, default=None*) – Individual weights for each sample.
- **dispersion** (*float, optional, default=None*) – The dispersion parameter. Estimated if absent.
- **robust** (*boolean, optional, default=None*) – Whether to compute robust standard errors instead of normal ones. If not specified, the model's `robust` attribute is used.
- **clusters** (*array-like, optional, default=None*) – Array with cluster membership. Clustered standard errors are computed if `clusters` is not `None`.
- **expected\_information** (*boolean, optional, default=None*) – Whether to use the expected or observed information matrix. Only relevant when computing robust standard errors. If not specified, the model's `expected_information` attribute is used.

#### Returns

A table of the regression results.

#### Return type

`pandas.DataFrame`

**covariance\_matrix**(*X=None, y=None, mu=None, offset=None, sample\_weight=None, dispersion=None, robust=None, clusters=None, expected\_information=None, store\_covariance\_matrix=False, skip\_checks=False*)

Calculate the covariance matrix for generalized linear models.

#### Parameters

- **X** (*{array-like, sparse matrix}, shape (n\_samples, n\_features), optional*) – Training data. Can be omitted if a covariance matrix has already been computed.
- **y** (*array-like, shape (n\_samples,), optional*) – Target values. Can be omitted if a covariance matrix has already been computed.
- **mu** (*array-like, optional, default=None*) – Array with predictions. Estimated if absent.
- **offset** (*array-like, optional, default=None*) – Array with additive offsets.
- **sample\_weight** (*array-like, shape (n\_samples,), optional, default=None*) – Individual weights for each sample.
- **dispersion** (*float, optional, default=None*) – The dispersion parameter. Estimated if absent.
- **robust** (*boolean, optional, default=None*) – Whether to compute robust standard errors instead of normal ones. If not specified, the model's `robust` attribute is used.

- **clusters** (*array-like, optional, default=None*) – Array with cluster membership. Clustered standard errors are computed if clusters is not None.
- **expected\_information** (*boolean, optional, default=None*) – Whether to use the expected or observed information matrix. Only relevant when computing robust standard errors. If not specified, the model's `expected_information` attribute is used.
- **store\_covariance\_matrix** (*boolean, optional, default=False*) – Whether to store the covariance matrix in the model instance. If a covariance matrix has already been stored, it will be overwritten.
- **skip\_checks** (*boolean, optional, default=False*) – Whether to skip input validation. For internal use only.

## Notes

We support three types of covariance matrices:

- non-robust
- robust (HC-1)
- clustered

For maximum-likelihood estimator, the covariance matrix takes the form  $\mathcal{H}^{-1}(\theta_0)\mathcal{I}(\theta_0)\mathcal{H}^{-1}(\theta_0)$  where  $\mathcal{H}^{-1}$  is the inverse Hessian and  $\mathcal{I}$  is the Information matrix. The different types of covariance matrices use different approximation of these quantities.

The non-robust covariance matrix is computed as the inverse of the Fisher information matrix. This assumes that the information matrix equality holds.

The robust (HC-1) covariance matrix takes the form  $\mathbf{H}^1(\hat{\theta})\mathbf{G}^T(\hat{\theta})\mathbf{G}(\hat{\theta})\mathbf{H}^1(\hat{\theta})$  where  $\mathbf{H}$  is the empirical Hessian and  $\mathbf{G}$  is the gradient. We apply a finite-sample correction of  $\frac{N}{N-p}$ .

The clustered covariance matrix uses a similar approach to the robust (HC-1) covariance matrix. However, instead of using  $\mathbf{G}^T(\hat{\theta})\mathbf{G}(\hat{\theta})$  directly, we first sum over all the groups first. The finite-sample correction is affected as well, becoming  $\frac{M}{M-1} \frac{N}{N-p}$  where  $M$  is the number of groups.

## References

**property family\_instance:** [\*ExponentialDispersionModel\*](#)

Return an [\*ExponentialDispersionModel\*](#).

**fit**(*X, y, sample\_weight=None, offset=None, store\_covariance\_matrix=False, clusters=None, weights\_sum=None*)

Fit a Generalized Linear Model.

### Parameters

- **X** (*{array-like, sparse matrix}, shape (n\_samples, n\_features)*) – Training data. Note that a float32 matrix is acceptable and will result in the entire algorithm being run in 32-bit precision. However, for problems that are poorly conditioned, this might result in poor convergence or flawed parameter estimates. If a Pandas data frame is provided, it may contain categorical columns. In that case, a separate coefficient will be estimated for each category. No category is omitted. This means that some regularization is required to fit models with an intercept or models with several categorical columns.
- **y** (*array-like, shape (n\_samples,)*) – Target values.

- **sample\_weight** (*array-like, shape (n\_samples,)*, *optional (default=None)*) – Individual weights  $w_i$  for each sample. Note that, for an Exponential Dispersion Model (EDM), one has  $\text{var}(y_i) = \phi \times v(\mu)/w_i$ . If  $y_i \sim EDM(\mu, \phi/w_i)$ , then  $\sum w_i y_i / \sum w_i \sim EDM(\mu, \phi / \sum w_i)$ , i.e. the mean of  $y$  is a weighted average with weights equal to **sample\_weight**.
- **offset** (*array-like, shape (n\_samples,)*, *optional (default=None)*) – Added to linear predictor. An offset of 3 will increase expected  $y$  by 3 if the link is linear and will multiply expected  $y$  by 3 if the link is logarithmic.
- **store\_covariance\_matrix** (*bool, optional (default=False)*) – Whether to estimate and store the covariance matrix of the parameter estimates. If **True**, the covariance matrix will be available in the **covariance\_matrix\_** attribute after fitting.
- **clusters** (*array-like, optional, default=None*) – Array with cluster membership. Clustered standard errors are computed if **clusters** is not **None**.
- **weights\_sum** (*float, optional (default=None)*) –

**Return type**

self

**get\_formatted\_diagnostics**(*full\_report=False, custom\_columns=None*)Get formatted diagnostics; can be printed with **\_report\_diagnostics**.**Parameters**

- **full\_report** (*bool, optional (default=False)*) – Print all available information. When **False** and **custom\_columns** is **None**, a restricted set of columns is printed out.
- **custom\_columns** (*iterable, optional (default=None)*) – Print only the specified columns.

**Return type**str | *DataFrame***get\_metadata\_routing**()

Get metadata routing of this object.

Please check User Guide on how the routing mechanism works.

**Returns****routing** – A **MetadataRequest** encapsulating routing information.**Return type****MetadataRequest****get\_params**(*deep=True*)

Get parameters for this estimator.

**Parameters**

**deep** (*bool, default=True*) – If **True**, will return the parameters for this estimator and contained subobjects that are estimators.

**Returns****params** – Parameter names mapped to their values.**Return type**

dict

**linear\_predictor**(*X*, *offset=None*, *alpha\_index=None*, *alpha=None*)

Compute the linear predictor,  $X * \text{coef\_} + \text{intercept\_}$ .

If *alpha\_search* is True, but *alpha\_index* and *alpha* are both None, we use the last alpha value *self.\_alphas*[-1].

**Parameters**

- **X**(*array-like*, *shape (n\_samples, n\_features)*) – Observations. *X* may be a pandas data frame with categorical types. If *X* was also a data frame with categorical types during fitting and a category wasn't observed at that point, the corresponding prediction will be *numpy.nan*.
- **offset**(*array-like*, *shape (n\_samples,)*, *optional (default=None)*) –
- **alpha\_index**(*int or list[int]*, *optional (default=None)*) – Sets the index of the alpha(s) to use in case *alpha\_search* is True. Incompatible with *alpha* (see below).
- **alpha**(*float or list[float]*, *optional (default=None)*) – Sets the alpha(s) to use in case *alpha\_search* is True. Incompatible with *alpha\_index* (see above).

**Returns**

The linear predictor.

**Return type**

array, shape (n\_samples, n\_alphas)

**property link\_instance:** [Link](#)

Return a [Link](#).

**predict**(*X*, *sample\_weight=None*, *offset=None*, *alpha\_index=None*, *alpha=None*)

Predict using GLM with feature matrix *X*.

If *alpha\_search* is True, but *alpha\_index* and *alpha* are both None, we use the last alpha value *self.\_alphas*[-1].

**Parameters**

- **X**(*array-like*, *shape (n\_samples, n\_features)*) – Observations. *X* may be a pandas data frame with categorical types. If *X* was also a data frame with categorical types during fitting and a category wasn't observed at that point, the corresponding prediction will be *numpy.nan*.
- **sample\_weight** (*array-like*, *shape (n\_samples,)*, *optional (default=None)*) – Sample weights to multiply predictions by.
- **offset**(*array-like*, *shape (n\_samples,)*, *optional (default=None)*) –
- **alpha\_index**(*int or list[int]*, *optional (default=None)*) – Sets the index of the alpha(s) to use in case *alpha\_search* is True. Incompatible with *alpha* (see below).
- **alpha**(*float or list[float]*, *optional (default=None)*) – Sets the alpha(s) to use in case *alpha\_search* is True. Incompatible with *alpha\_index* (see above).

**Returns**

Predicted values times *sample\_weight*.

**Return type**

array, shape (n\_samples, n\_alphas)

**report\_diagnostics**(*full\_report=False*, *custom\_columns=None*)

Print diagnostics to stdout.

**Parameters**

- **full\_report** (*bool, optional (default=False)*) – Print all available information. When `False` and `custom_columns` is `None`, a restricted set of columns is printed out.
- **custom\_columns** (*iterable, optional (default=None)*) – Print only the specified columns.

**Return type**

None

**score**(*X, y, sample\_weight=None, offset=None*)Compute  $D^2$ , the percentage of deviance explained.

$D^2$  is a generalization of the coefficient of determination  $R^2$ . The  $R^2$  uses the squared error and the  $D^2$ , the deviance. Note that those two are equal for `family='normal'`.

$D^2$  is defined as  $D^2 = 1 - \frac{D(y_{\text{true}}, y_{\text{pred}})}{D_{\text{null}}}$ ,  $D_{\text{null}}$  is the null deviance, i.e. the deviance of a model with intercept alone. The best possible score is one and it can be negative.

**Parameters**

- **X** (*{array-like, sparse matrix}, shape (n\_samples, n\_features)*) – Test samples.
- **y** (*array-like, shape (n\_samples,)*) – True values of target.
- **sample\_weight** (*array-like, shape (n\_samples,), optional (default=None)*) – Sample weights.
- **offset** (*array-like, shape (n\_samples,), optional (default=None)*) –

**Returns** $D^2$  of `self.predict(X)` w.r.t. `y`.**Return type**

float

**set\_params**(*\*\*params*)

Set the parameters of this estimator.

The method works on simple estimators as well as on nested objects (such as `Pipeline`). The latter have parameters of the form `<component>__<parameter>` so that it's possible to update each component of a nested object.

**Parameters****\*\*params** (*dict*) – Estimator parameters.**Returns****self** – Estimator instance.**Return type**

estimator instance

**std\_errors**(*X=None, y=None, mu=None, offset=None, sample\_weight=None, dispersion=None, robust=None, clusters=None, expected\_information=None, store\_covariance\_matrix=False*)

Calculate standard errors for generalized linear models.

See `covariance_matrix` for an in-depth explanation of how the standard errors are computed.**Parameters**

- **X** (*{array-like, sparse matrix}, shape (n\_samples, n\_features), optional*) – Training data. Can be omitted if a covariance matrix has already been computed.
- **y** (*array-like, shape (n\_samples,), optional*) – Target values. Can be omitted if a covariance matrix has already been computed.
- **mu** (*array-like, optional, default=None*) – Array with predictions. Estimated if absent.
- **offset** (*array-like, optional, default=None*) – Array with additive offsets.
- **sample\_weight** (*array-like, shape (n\_samples,), optional, default=None*) – Individual weights for each sample.
- **dispersion** (*float, optional, default=None*) – The dispersion parameter. Estimated if absent.
- **robust** (*boolean, optional, default=None*) – Whether to compute robust standard errors instead of normal ones. If not specified, the model's `robust` attribute is used.
- **clusters** (*array-like, optional, default=None*) – Array with cluster membership. Clustered standard errors are computed if `clusters` is not `None`.
- **expected\_information** (*boolean, optional, default=None*) – Whether to use the expected or observed information matrix. Only relevant when computing robust standard errors. If not specified, the model's `expected_information` attribute is used.
- **store\_covariance\_matrix** (*boolean, optional, default=False*) – Whether to store the covariance matrix in the model instance. If a covariance matrix has already been stored, it will be overwritten.

**wald\_test** (*R=None, features=None, r=None, X=None, y=None, mu=None, offset=None, sample\_weight=None, dispersion=None, robust=None, clusters=None, expected\_information=None*)

Compute the Wald test statistic and p-value for a linear hypothesis.

The left hand side of the hypothesis may be specified in the following ways:

- **R**: The restriction matrix representing the linear combination of coefficients to test.
- **features**: The name of a feature or a list of features to test.

The right hand side of the tested hypothesis is specified by **r**.

#### Parameters

- **R** (*np.ndarray, optional, default=None*) – The restriction matrix representing the linear combination of coefficients to test.
- **features** (*Union[str, list[str]], optional, default=None*) – The name of a feature or a list of features to test.
- **r** (*np.ndarray, optional, default=None*) – The vector representing the values of the linear combination. If `None`, the test is for whether the linear combinations of the coefficients are zero.
- **X** (*{array-like, sparse matrix}, shape (n\_samples, n\_features), optional*) – Training data. Can be omitted if a covariance matrix has already been computed.
- **y** (*array-like, shape (n\_samples,), optional*) – Target values. Can be omitted if a covariance matrix has already been computed.

- **mu** (*array-like, optional, default=None*) – Array with predictions. Estimated if absent.
- **offset** (*array-like, optional, default=None*) – Array with additive offsets.
- **sample\_weight** (*array-like, shape (n\_samples,), optional, default=None*) – Individual weights for each sample.
- **dispersion** (*float, optional, default=None*) – The dispersion parameter. Estimated if absent.
- **robust** (*boolean, optional, default=None*) – Whether to compute robust standard errors instead of normal ones. If not specified, the model's `robust` attribute is used.
- **clusters** (*array-like, optional, default=None*) – Array with cluster membership. Clustered standard errors are computed if `clusters` is not `None`.
- **expected\_information** (*boolean, optional, default=None*) – Whether to use the expected or observed information matrix. Only relevant when computing robust standard errors. If not specified, the model's `expected_information` attribute is used.

#### Returns

NamedTuple with test statistic, p-value, and degrees of freedom.

#### Return type

WaldTestResult

```
class glum.GeneralizedLinearRegressorCV(l1_ratio=0, P1='identity', P2='identity', fit_intercept=True,
                                       family='normal', link='auto', solver='auto', max_iter=100,
                                       gradient_tol=None, step_size_tol=None, hessian_approx=0.0,
                                       warm_start=False, n_alphas=100, alphas=None,
                                       min_alpha_ratio=None, min_alpha=None, start_params=None,
                                       selection='cyclic', random_state=None, copy_X=True,
                                       check_input=True, verbose=0, scale_predictors=False,
                                       lower_bounds=None, upper_bounds=None, A_ineq=None,
                                       b_ineq=None, force_all_finite=True, cv=None, n_jobs=None,
                                       drop_first=False, robust=True, expected_information=False)
```

Bases: `GeneralizedLinearRegressorBase`

Generalized linear model with iterative fitting along a regularization path.

The best model is selected by cross-validation.

Cross-validated regression via a Generalized Linear Model (GLM) with penalties. For more on GLMs and on these parameters, see the documentation for [GeneralizedLinearRegressor](#). CV conventions follow `sklearn.linear_model.LassoCV`.

#### Parameters

- **l1\_ratio** (*float or array of floats, optional (default=0)*) – If you pass `l1_ratio` as an array, the `fit` method will choose the best value of `l1_ratio` and store it as `self.l1_ratio`.
- **P1** (*{'identity', array-like, None}, shape (n\_features,), optional (default='identity')*) – This array controls the strength of the regularization for each coefficient independently. A high value will lead to higher regularization while a value of zero will remove the regularization on this parameter. Note that `n_features = X.shape[1]`. If `X` is a pandas DataFrame with a categorical dtype and `P1` has the same size as the number of columns, the penalty of the categorical column will be applied to all the levels of the categorical.

- **P2** (`{'identity', array-like, sparse matrix, None}`, `shape (n_features,)` or `(n_features, n_features)`, `optional (default='identity')`) – With this option, you can set the P2 matrix in the L2 penalty  $w^*P2*w$ . This gives a fine control over this penalty (Tikhonov regularization). A 2d array is directly used as the square matrix P2. A 1d array is interpreted as diagonal (square) matrix. The default 'identity' and None set the identity matrix, which gives the usual squared L2-norm. If you just want to exclude certain coefficients, pass a 1d array filled with 1 and 0 for the coefficients to be excluded. Note that P2 must be positive semi-definite. If X is a pandas DataFrame with a categorical dtype and P2 has the same size as the number of columns, the penalty of the categorical column will be applied to all the levels of the categorical. Note that if P2 is two-dimensional, its size needs to be of the same length as the expanded X matrix.
- **fit\_intercept** (`bool`, `optional (default=True)`) – Specifies if a constant (a.k.a. bias or intercept) should be added to the linear predictor ( $X * \text{coef} + \text{intercept}$ ).
- **family** (`str` or `ExponentialDispersionModel`, `optional (default='normal')`) – The distributional assumption of the GLM, i.e. the loss function to minimize. If a string, one of: 'binomial', 'gamma', 'gaussian', 'inverse.gaussian', 'normal', 'poisson', 'tweedie' or 'negative.binomial'. Note that 'tweedie' sets the power of the Tweedie distribution to 1.5; to use another value, specify it in parentheses (e.g., 'tweedie (1.5)'). The same applies for 'negative.binomial' and theta parameter.
- **link** (`{'auto', 'identity', 'log', 'logit', 'cloglog'}`, `Link` or `None`, `optional (default='auto')`) – The link function of the GLM, i.e. mapping from linear predictor ( $X * \text{coef}$ ) to expectation ( $\mu$ ). Option 'auto' sets the link depending on the chosen family as follows:
  - 'identity' for family 'normal'
  - 'log' for families 'poisson', 'gamma', 'inverse.gaussian' and 'negative.binomial'.
  - 'logit' for family 'binomial'
- **solver** (`{'auto', 'irls-cd', 'irls-ls', 'lbfgs', 'trust-constr'}`, `optional (default='auto')`) – Algorithm to use in the optimization problem:
  - 'auto': 'irls-ls' if `l1_ratio` is zero and 'irls-cd' otherwise.
  - 'irls-cd': Iteratively reweighted least squares with a coordinate descent inner solver. This can deal with L1 as well as L2 penalties. Note that in order to avoid unnecessary memory duplication of X in the fit method, X should be directly passed as a Fortran-contiguous Numpy array or sparse CSC matrix.
  - 'irls-ls': Iteratively reweighted least squares with a least squares inner solver. This algorithm cannot deal with L1 penalties.
  - 'lbfgs': Scipy's L-BFGS-B optimizer. It cannot deal with L1 penalties.
- **max\_iter** (`int`, `optional (default=100)`) – The maximal number of iterations for solver algorithms.
- **gradient\_tol** (`float`, `optional (default=None)`) – Stopping criterion. If None, solver-specific defaults will be used. The default value for most solvers is  $1e-4$ , except for 'trust-constr', which requires more conservative convergence settings and has a default value of  $1e-8$ .

For the IRLS-LS, L-BFGS and trust-constr solvers, the iteration will stop when  $\max\{|g_i|, i = 1, \dots, n\} \leq \text{tol}$ , where  $g_i$  is the  $i$ -th component of the gradient (derivative) of the objective function. For the CD solver, convergence is reached when  $\text{sum}_i(|\text{minimum}$



norm of  $g_i$ ), where  $g_i$  is the subgradient of the objective and the minimum norm of  $g_i$  is the element of the subgradient with the smallest L2 norm.

If you wish to only use a step-size tolerance, set `gradient_tol` to a very small number.

- **step\_size\_tol** (*float, optional (default=None)*) – Alternative stopping criterion. For the IRLS-LS and IRLS-CD solvers, the iteration will stop when the L2 norm of the step size is less than `step_size_tol`. This stopping criterion is disabled when `step_size_tol` is `None`.
- **hessian\_approx** (*float, optional (default=0.0)*) – The threshold below which data matrix rows will be ignored for updating the Hessian. See the algorithm documentation for the IRLS algorithm for further details.
- **warm\_start** (*bool, optional (default=False)*) – Whether to reuse the solution of the previous call to `fit` as initialization for `coef_` and `intercept_` (supersedes `start_params`). If `False` or if the attribute `coef_` does not exist (first call to `fit`), `start_params` sets the start values for `coef_` and `intercept_`.
- **n\_alphas** (*int, optional (default=100)*) – Number of alphas along the regularization path
- **alphas** (*array-like, optional (default=None)*) – List of alphas for which to compute the models. If `None`, the alphas are set automatically. Setting `None` is preferred.
- **min\_alpha\_ratio** (*float, optional (default=None)*) – Length of the path. `min_alpha_ratio=1e-6` means that `min_alpha / max_alpha = 1e-6`. If `None`, `1e-6` is used.
- **min\_alpha** (*float, optional (default=None)*) – Minimum alpha to estimate the model with. The grid will then be created over `[max_alpha, min_alpha]`.
- **start\_params** (*array-like, shape (n\_features\*,), optional (default=None)*) – Relevant only if `warm_start` is `False` or if `fit` is called for the first time (so that `self.coef_` does not exist yet). If `None`, all coefficients are set to zero and the start value for the intercept is the weighted average of `y` (If `fit_intercept` is `True`). If an array, used directly as start values; if `fit_intercept` is `True`, its first element is assumed to be the start value for the `intercept_`. Note that `n_features* = X.shape[1] + fit_intercept`, i.e. it includes the intercept.
- **selection** (*str, optional (default='cyclic')*) – For the CD solver 'cd', the coordinates (features) can be updated in either cyclic or random order. If set to 'random', a random coefficient is updated every iteration rather than looping over features sequentially in the same order, which often leads to significantly faster convergence, especially when `gradient_tol` is higher than `1e-4`.
- **random\_state** (*int or RandomState, optional (default=None)*) – The seed of the pseudo random number generator that selects a random feature to be updated for the CD solver. If an integer, `random_state` is the seed used by the random number generator; if a `RandomState` instance, `random_state` is the random number generator; if `None`, the random number generator is the `RandomState` instance used by `np.random`. Used when `selection` is 'random'.
- **copy\_X** (*bool, optional (default=None)*) – Whether to copy `X`. Since `X` is never modified by `GeneralizedLinearRegressor`, this is unlikely to be needed; this option exists mainly for compatibility with other scikit-learn estimators. If `False`, `X` will not be copied and there will be an error if you pass an `X` in the wrong format, such as providing integer `X` and float `y`. If `None`, `X` will not be copied unless it is in the wrong format.

- **check\_input** (*bool*, *optional* (*default=True*)) – Whether to bypass several checks on input: *y* values in range of *family*, *sample\_weight* non-negative, P2 positive semi-definite. Don't use this parameter unless you know what you are doing.
- **verbose** (*int*, *optional* (*default=0*)) – For the IRLS solver, any positive number will result in a pretty progress bar showing convergence. This features requires having the *tqdm* package installed. For the L-BFGS solver, set **verbose** to any positive number for verbosity.
- **scale\_predictors** (*bool*, *optional* (*default=False*)) – If *True*, estimate a scaled model where all predictors have a standard deviation of 1. This can result in better estimates if predictors are on very different scales (for example, centimeters and kilometers).

Advanced developer note: Internally, predictors are always rescaled for computational reasons, but this only affects results if **scale\_predictors** is *True*.

- **lower\_bounds** (*array-like*, *shape* (*n\_features*), *optional* (*default=None*)) – Set a lower bound for the coefficients. Setting bounds forces the use of the coordinate descent solver ('irls-cd').
- **upper\_bounds** (*array-like*, *shape*=(*n\_features*), *optional* (*default=None*)) – See **lower\_bounds**.
- **A\_ineq** (*array-like*, *shape*=(*n\_constraints*, *n\_features*), *optional* (*default=None*)) – Constraint matrix for linear inequality constraints of the form  $A_{\text{ineq}} w \leq b_{\text{ineq}}$ .
- **b\_ineq** (*array-like*, *shape*=(*n\_constraints*,), *optional* (*default=None*)) – Constraint vector for linear inequality constraints of the form  $A_{\text{ineq}} w \leq b_{\text{ineq}}$ .
- **cv** (*int*, *cross-validation generator or Iterable*, *optional* (*default=None*)) – Determines the cross-validation splitting strategy. One of:
  - *None*, to use the default 5-fold cross-validation,
  - *int*, to specify the number of folds.
  - *Iterable* yielding (train, test) splits as arrays of indices.

For integer/*None* inputs, *KFold* is used

- **n\_jobs** (*int*, *optional* (*default=None*)) – The maximum number of concurrently running jobs. The number of jobs that are needed is  $\text{len}(l1\_ratio) \times n\_folds$ . -1 is the same as the number of CPU on your machine. *None* means 1 unless in a *joblib.parallel\_backend* context.
- **drop\_first** (*bool*, *optional* (*default = False*)) – If *True*, drop the first column when encoding categorical variables.
- **force\_all\_finite** (*bool*) –
- **robust** (*bool*) –
- **expected\_information** (*bool*) –

#### **alpha\_**

The amount of regularization chosen by cross validation.

#### **Type**

float

#### **alphas\_**

Alphas used by the model.

**Type**

array, shape (n\_l1\_ratios, n\_alphas)

**l1\_ratio\_**

The compromise between L1 and L2 regularization chosen by cross validation.

**Type**

float

**coef\_**

Estimated coefficients for the linear predictor in the GLM at the optimal (l1\_ratio\_, alpha\_).

**Type**

array, shape (n\_features,)

**intercept\_**

Intercept (a.k.a. bias) added to linear predictor.

**Type**

float

**n\_iter\_**

The number of iterations run by the CD solver to reach the specified tolerance for the optimal alpha.

**Type**

int

**coef\_path\_**

Estimated coefficients for the linear predictor in the GLM at every point along the regularization path.

**Type**

array, shape (n\_folds, n\_l1\_ratios, n\_alphas, n\_features)

**deviance\_path\_**

Deviance for the test set on each fold, varying alpha.

**Type**

array, shape(n\_folds, n\_alphas)

**robust**

If true, then robust standard errors are computed by default.

**Type**

bool, optional (default = False)

**expected\_information**

If true, then the expected information matrix is computed by default. Only relevant when computing robust standard errors.

**Type**

bool, optional (default = False)

**coef\_table**(*confidence\_level=0.95, X=None, y=None, mu=None, offset=None, sample\_weight=None, dispersion=None, robust=None, clusters=None, expected\_information=None*)

Get a table of the regression coefficients.

Includes coefficient estimates, standard errors, t-values, p-values and confidence intervals.

**Parameters**

- **confidence\_level** (*float, optional, default=0.95*) – The confidence level for the confidence intervals.

- **X** (*{array-like, sparse matrix}, shape (n\_samples, n\_features), optional*) – Training data. Can be omitted if a covariance matrix has already been computed or if standard errors, etc. are not desired.
- **y** (*array-like, shape (n\_samples,), optional*) – Target values. Can be omitted if a covariance matrix has already been computed.
- **mu** (*array-like, optional, default=None*) – Array with predictions. Estimated if absent.
- **offset** (*array-like, optional, default=None*) – Array with additive offsets.
- **sample\_weight** (*array-like, shape (n\_samples,), optional, default=None*) – Individual weights for each sample.
- **dispersion** (*float, optional, default=None*) – The dispersion parameter. Estimated if absent.
- **robust** (*boolean, optional, default=None*) – Whether to compute robust standard errors instead of normal ones. If not specified, the model's `robust` attribute is used.
- **clusters** (*array-like, optional, default=None*) – Array with cluster membership. Clustered standard errors are computed if `clusters` is not `None`.
- **expected\_information** (*boolean, optional, default=None*) – Whether to use the expected or observed information matrix. Only relevant when computing robust standard errors. If not specified, the model's `expected_information` attribute is used.

#### Returns

A table of the regression results.

#### Return type

`pandas.DataFrame`

**covariance\_matrix**(*X=None, y=None, mu=None, offset=None, sample\_weight=None, dispersion=None, robust=None, clusters=None, expected\_information=None, store\_covariance\_matrix=False, skip\_checks=False*)

Calculate the covariance matrix for generalized linear models.

#### Parameters

- **X** (*{array-like, sparse matrix}, shape (n\_samples, n\_features), optional*) – Training data. Can be omitted if a covariance matrix has already been computed.
- **y** (*array-like, shape (n\_samples,), optional*) – Target values. Can be omitted if a covariance matrix has already been computed.
- **mu** (*array-like, optional, default=None*) – Array with predictions. Estimated if absent.
- **offset** (*array-like, optional, default=None*) – Array with additive offsets.
- **sample\_weight** (*array-like, shape (n\_samples,), optional, default=None*) – Individual weights for each sample.
- **dispersion** (*float, optional, default=None*) – The dispersion parameter. Estimated if absent.
- **robust** (*boolean, optional, default=None*) – Whether to compute robust standard errors instead of normal ones. If not specified, the model's `robust` attribute is used.
- **clusters** (*array-like, optional, default=None*) – Array with cluster membership. Clustered standard errors are computed if `clusters` is not `None`.

- **expected\_information** (*boolean, optional, default=None*) – Whether to use the expected or observed information matrix. Only relevant when computing robust standard errors. If not specified, the model’s `expected_information` attribute is used.
- **store\_covariance\_matrix** (*boolean, optional, default=False*) – Whether to store the covariance matrix in the model instance. If a covariance matrix has already been stored, it will be overwritten.
- **skip\_checks** (*boolean, optional, default=False*) – Whether to skip input validation. For internal use only.

## Notes

We support three types of covariance matrices:

- non-robust
- robust (HC-1)
- clustered

For maximum-likelihood estimator, the covariance matrix takes the form  $\mathcal{H}^{-1}(\theta_0)\mathcal{I}(\theta_0)\mathcal{H}^{-1}(\theta_0)$  where  $\mathcal{H}^{-1}$  is the inverse Hessian and  $\mathcal{I}$  is the Information matrix. The different types of covariance matrices use different approximation of these quantities.

The non-robust covariance matrix is computed as the inverse of the Fisher information matrix. This assumes that the information matrix equality holds.

The robust (HC-1) covariance matrix takes the form  $\mathbf{H}^1(\hat{\theta})\mathbf{G}^T(\hat{\theta})\mathbf{G}(\hat{\theta})\mathbf{H}^1(\hat{\theta})$  where  $\mathbf{H}$  is the empirical Hessian and  $\mathbf{G}$  is the gradient. We apply a finite-sample correction of  $\frac{N}{N-p}$ .

The clustered covariance matrix uses a similar approach to the robust (HC-1) covariance matrix. However, instead of using  $\mathbf{G}^T(\hat{\theta})\mathbf{G}(\hat{\theta})$  directly, we first sum over all the groups first. The finite-sample correction is affected as well, becoming  $\frac{M}{M-1} \frac{N}{N-p}$  where  $M$  is the number of groups.

## References

**property family\_instance:** [\*ExponentialDispersionModel\*](#)

Return an [\*ExponentialDispersionModel\*](#).

**fit**( $X, y$ , *sample\_weight=None, offset=None, store\_covariance\_matrix=False, clusters=None*)

Choose the best model along a ‘regularization path’ by cross-validation.

### Parameters

- **X** (*{array-like, sparse matrix}, shape (n\_samples, n\_features)*) – Training data. Note that a `float32` matrix is acceptable and will result in the entire algorithm being run in 32-bit precision. However, for problems that are poorly conditioned, this might result in poor convergence or flawed parameter estimates. If a Pandas data frame is provided, it may contain categorical columns. In that case, a separate coefficient will be estimated for each category. No category is omitted. This means that some regularization is required to fit models with an intercept or models with several categorical columns.
- **y** (*array-like, shape (n\_samples,)*) – Target values.
- **sample\_weight** (*(array-like, shape (n\_samples,)), optional (default=None)*) – Individual weights  $w_i$  for each sample. Note that, for an Exponential Dispersion Model (EDM), one has  $\text{var}(y_i) = \phi \times v(mu)/w_i$ . If  $y_i \sim EDM(\mu, \phi/w_i)$ ,

then  $\sum w_i y_i / \sum w_i \sim EDM(\mu, \phi / \sum w_i)$ , i.e. the mean of  $y$  is a weighted average with weights equal to `sample_weight`.

- **offset** (*array-like, shape (n\_samples,)*, *optional (default=None)*) – Added to linear predictor. An offset of 3 will increase expected  $y$  by 3 if the link is linear and will multiply expected  $y$  by 3 if the link is logarithmic.
- **store\_covariance\_matrix** (*bool, optional (default=False)*) – Whether to store the covariance matrix of the parameter estimates corresponding to the best model.
- **clusters** (*array-like, optional, default=None*) – Array with cluster membership. Clustered standard errors are computed if clusters is not None.

**get\_formatted\_diagnostics**(*full\_report=False, custom\_columns=None*)

Get formatted diagnostics; can be printed with `_report_diagnostics`.

#### Parameters

- **full\_report** (*bool, optional (default=False)*) – Print all available information. When False and `custom_columns` is None, a restricted set of columns is printed out.
- **custom\_columns** (*iterable, optional (default=None)*) – Print only the specified columns.

#### Return type

str | *DataFrame*

**get\_metadata\_routing()**

Get metadata routing of this object.

Please check User Guide on how the routing mechanism works.

#### Returns

**routing** – A `MetadataRequest` encapsulating routing information.

#### Return type

`MetadataRequest`

**get\_params**(*deep=True*)

Get parameters for this estimator.

#### Parameters

**deep** (*bool, default=True*) – If True, will return the parameters for this estimator and contained subobjects that are estimators.

#### Returns

**params** – Parameter names mapped to their values.

#### Return type

dict

**linear\_predictor**(*X, offset=None, alpha\_index=None, alpha=None*)

Compute the linear predictor,  $X * \text{coef\_} + \text{intercept\_}$ .

If `alpha_search` is True, but `alpha_index` and `alpha` are both None, we use the last alpha value `self._alphas[-1]`.

#### Parameters

- **X** (*array-like, shape (n\_samples, n\_features)*) – Observations. **X** may be a pandas data frame with categorical types. If **X** was also a data frame with categorical types during fitting and a category wasn't observed at that point, the corresponding prediction will be `numpy.nan`.

- **offset** (*array-like, shape (n\_samples,)*, *optional (default=None)*) –
- **alpha\_index** (*int or list[int]*, *optional (default=None)*) – Sets the index of the alpha(s) to use in case `alpha_search` is True. Incompatible with `alpha` (see below).
- **alpha** (*float or list[float]*, *optional (default=None)*) – Sets the alpha(s) to use in case `alpha_search` is True. Incompatible with `alpha_index` (see above).

**Returns**

The linear predictor.

**Return type**

array, shape (n\_samples, n\_alphas)

**property link\_instance:** [Link](#)

Return a [Link](#).

**predict** (*X, sample\_weight=None, offset=None, alpha\_index=None, alpha=None*)

Predict using GLM with feature matrix *X*.

If `alpha_search` is True, but `alpha_index` and `alpha` are both None, we use the last alpha value `self._alphas[-1]`.

**Parameters**

- **X** (*array-like, shape (n\_samples, n\_features)*) – Observations. *X* may be a pandas data frame with categorical types. If *X* was also a data frame with categorical types during fitting and a category wasn't observed at that point, the corresponding prediction will be `numpy.nan`.
- **sample\_weight** (*array-like, shape (n\_samples,)*, *optional (default=None)*) – Sample weights to multiply predictions by.
- **offset** (*array-like, shape (n\_samples,)*, *optional (default=None)*) –
- **alpha\_index** (*int or list[int]*, *optional (default=None)*) – Sets the index of the alpha(s) to use in case `alpha_search` is True. Incompatible with `alpha` (see below).
- **alpha** (*float or list[float]*, *optional (default=None)*) – Sets the alpha(s) to use in case `alpha_search` is True. Incompatible with `alpha_index` (see above).

**Returns**

Predicted values times `sample_weight`.

**Return type**

array, shape (n\_samples, n\_alphas)

**report\_diagnostics** (*full\_report=False, custom\_columns=None*)

Print diagnostics to stdout.

**Parameters**

- **full\_report** (*bool, optional (default=False)*) – Print all available information. When False and `custom_columns` is None, a restricted set of columns is printed out.
- **custom\_columns** (*iterable, optional (default=None)*) – Print only the specified columns.

**Return type**

None

**score**(*X*, *y*, *sample\_weight=None*, *offset=None*)

Compute  $D^2$ , the percentage of deviance explained.

$D^2$  is a generalization of the coefficient of determination  $R^2$ . The  $R^2$  uses the squared error and the  $D^2$ , the deviance. Note that those two are equal for `family='normal'`.

$D^2$  is defined as  $D^2 = 1 - \frac{D(y_{\text{true}}, y_{\text{pred}})}{D_{\text{null}}}$ ,  $D_{\text{null}}$  is the null deviance, i.e. the deviance of a model with intercept alone. The best possible score is one and it can be negative.

**Parameters**

- **X** (*{array-like, sparse matrix}*, *shape (n\_samples, n\_features)*) – Test samples.
- **y** (*array-like*, *shape (n\_samples,)*) – True values of target.
- **sample\_weight** (*array-like*, *shape (n\_samples,)*, *optional (default=None)*) – Sample weights.
- **offset** (*array-like*, *shape (n\_samples,)*, *optional (default=None)*) –

**Returns**

$D^2$  of `self.predict(X)` w.r.t. *y*.

**Return type**

float

**set\_params**(*\*\*params*)

Set the parameters of this estimator.

The method works on simple estimators as well as on nested objects (such as `Pipeline`). The latter have parameters of the form `<component>__<parameter>` so that it's possible to update each component of a nested object.

**Parameters**

**\*\*params** (*dict*) – Estimator parameters.

**Returns**

**self** – Estimator instance.

**Return type**

estimator instance

**std\_errors**(*X=None*, *y=None*, *mu=None*, *offset=None*, *sample\_weight=None*, *dispersion=None*, *robust=None*, *clusters=None*, *expected\_information=None*, *store\_covariance\_matrix=False*)

Calculate standard errors for generalized linear models.

See `covariance_matrix` for an in-depth explanation of how the standard errors are computed.

**Parameters**

- **X** (*{array-like, sparse matrix}*, *shape (n\_samples, n\_features)*, *optional*) – Training data. Can be omitted if a covariance matrix has already been computed.
- **y** (*array-like*, *shape (n\_samples,)*, *optional*) – Target values. Can be omitted if a covariance matrix has already been computed.
- **mu** (*array-like*, *optional*, *default=None*) – Array with predictions. Estimated if absent.
- **offset** (*array-like*, *optional*, *default=None*) – Array with additive offsets.



- **sample\_weight** (array-like, shape (n\_samples,)), optional, default=None) – Individual weights for each sample.
- **dispersion** (float, optional, default=None) – The dispersion parameter. Estimated if absent.
- **robust** (boolean, optional, default=None) – Whether to compute robust standard errors instead of normal ones. If not specified, the model's `robust` attribute is used.
- **clusters** (array-like, optional, default=None) – Array with cluster membership. Clustered standard errors are computed if `clusters` is not None.
- **expected\_information** (boolean, optional, default=None) – Whether to use the expected or observed information matrix. Only relevant when computing robust standard errors. If not specified, the model's `expected_information` attribute is used.
- **store\_covariance\_matrix** (boolean, optional, default=False) – Whether to store the covariance matrix in the model instance. If a covariance matrix has already been stored, it will be overwritten.

**wald\_test**(R=None, features=None, r=None, X=None, y=None, mu=None, offset=None, sample\_weight=None, dispersion=None, robust=None, clusters=None, expected\_information=None)

Compute the Wald test statistic and p-value for a linear hypothesis.

The left hand side of the hypothesis may be specified in the following ways:

- **R**: The restriction matrix representing the linear combination of coefficients to test.
- **features**: The name of a feature or a list of features to test.

The right hand side of the tested hypothesis is specified by **r**.

#### Parameters

- **R** (np.ndarray, optional, default=None) – The restriction matrix representing the linear combination of coefficients to test.
- **features** (Union[str, list[str]], optional, default=None) – The name of a feature or a list of features to test.
- **r** (np.ndarray, optional, default=None) – The vector representing the values of the linear combination. If None, the test is for whether the linear combinations of the coefficients are zero.
- **X** ({array-like, sparse matrix}, shape (n\_samples, n\_features), optional) – Training data. Can be omitted if a covariance matrix has already been computed.
- **y** (array-like, shape (n\_samples,)), optional) – Target values. Can be omitted if a covariance matrix has already been computed.
- **mu** (array-like, optional, default=None) – Array with predictions. Estimated if absent.
- **offset** (array-like, optional, default=None) – Array with additive offsets.
- **sample\_weight** (array-like, shape (n\_samples,)), optional, default=None) – Individual weights for each sample.
- **dispersion** (float, optional, default=None) – The dispersion parameter. Estimated if absent.

- **robust** (*boolean, optional, default=None*) – Whether to compute robust standard errors instead of normal ones. If not specified, the model's `robust` attribute is used.
- **clusters** (*array-like, optional, default=None*) – Array with cluster membership. Clustered standard errors are computed if `clusters` is not `None`.
- **expected\_information** (*boolean, optional, default=None*) – Whether to use the expected or observed information matrix. Only relevant when computing robust standard errors. If not specified, the model's `expected_information` attribute is used.

**Returns**

NamedTuple with test statistic, p-value, and degrees of freedom.

**Return type**

WaldTestResult

**class** `glum.IdentityLink`

Bases: [`Link`](#)

The identity link function  $g(x) = x$ .

**derivative**(*mu*)

Get the derivative of the identity link, a vector of ones.

See superclass documentation.

**Parameters**

**mu** (*array-like*) –

**inverse**(*lin\_pred*)

Compute the inverse link function  $h(\text{lin\_pred})$ .

Gives the inverse relationship between linear predictor and the mean  $\mu = E(Y)$ , i.e.  $h(\text{linear predictor}) = \mu$ .

**Parameters**

**lin\_pred** (*array-like, shape (n\_samples,)*) – Usually the (fitted) linear predictor.

**inverse\_derivative**(*lin\_pred*)

Compute the derivative of the inverse link function  $h'(\text{lin\_pred})$ .

**Parameters**

**lin\_pred** (*array-like, shape (n\_samples,)*) – Usually the (fitted) linear predictor.

**inverse\_derivative2**(*lin\_pred*)

Compute second derivative of the inverse link function  $h''(\text{lin\_pred})$ .

**Parameters**

**lin\_pred** (*array-like, shape (n\_samples,)*) – Usually the (fitted) linear predictor.

**link**(*mu*)

Return  $\mu$  (identity link).

See superclass documentation.

**Parameters**

**mu** (*array-like*) –

**class** `glum.InverseGaussianDistribution`

Bases: [`TweedieDistribution`](#)

Class for the scaled Inverse Gaussian distribution.

**deviance**(*y*, *mu*, *sample\_weight=None*)

Compute the deviance.

**Parameters**

- *y* (array-like, shape (n\_samples,)) – Target values.
- *mu* (array-like, shape (n\_samples,)) – Predicted mean.
- *sample\_weight* (array-like, shape (n\_samples,)), optional (default=1) – Sample weights.

**Return type**

float

**deviance\_derivative**(*y*, *mu*, *sample\_weight=1*)

Compute the derivative of the deviance with respect to *mu*.

**Parameters**

- *y* (array-like, shape (n\_samples,)) – Target values.
- *mu* (array-like, shape (n\_samples,)) – Predicted mean.
- *sample\_weight* (array-like, shape (n\_samples,)) (default=1) – Weights or exposure to which variance is inverse proportional.

**Return type**

array-like, shape (n\_samples,)

**dispersion**(*y*, *mu*, *sample\_weight=None*, *ddof=1*, *method='pearson'*)

Estimate the dispersion parameter  $\phi$ .

**Parameters**

- *y* (array-like, shape (n\_samples,)) – Target values.
- *mu* (array-like, shape (n\_samples,)) – Predicted mean.
- *sample\_weight* (array-like, shape (n\_samples,)), optional (default=None) – Weights or exposure to which variance is inversely proportional.
- *ddof* (int, optional (default=1)) – Degrees of freedom consumed by the model for *mu*.
- {'pearson' (*method* =) – Whether to base the estimate on the Pearson residuals or the deviance.
- 'deviance' – Whether to base the estimate on the Pearson residuals or the deviance.
- (default='pearson') (optional) – Whether to base the estimate on the Pearson residuals or the deviance.

**Return type**

float

**eta\_mu\_deviance**(*link*, *factor*, *cur\_eta*, *X\_dot\_d*, *y*, *sample\_weight*)

Compute eta, mu and the deviance.

Compute: \* the linear predictor, eta, as *cur\_eta* + *factor* \* *X\_dot\_d*; \* the link-function-transformed prediction, mu; \* the deviance.

**Returns**

- *numpy.ndarray*, shape (*X.shape*[0],) – The linear predictor, eta.

- *numpy.ndarray*, *shape* (*X.shape*[0],) – The link-function-transformed prediction,  $\mu$ .
- *float* – The deviance.

**Parameters**

- **link** (*Link*) –
- **factor** (*float*) –
- **cur\_eta** (*ndarray*) –
- **X\_dot\_d** (*ndarray*) –
- **y** (*ndarray*) –
- **sample\_weight** (*ndarray*) –

**in\_y\_range(x)**

Return True if *x* is in the valid range of the EDM.

**Parameters**

**x** (*array-like*, *shape* (*n\_samples*,)) – Target values.

**Return type**

*np.ndarray*

**property include\_lower\_bound: bool**

Return whether *lower\_bound* is allowed as a value of *y*.

**log\_likelihood(y, mu, sample\_weight=None, dispersion=None)**

Compute the log likelihood.

For  $1 < \text{power} < 2$ , we use the series approximation by Dunn and Smyth (2005) to compute the normalization term.

**Parameters**

- **y** (*array-like*, *shape* (*n\_samples*,)) – Target values.
- **mu** (*array-like*, *shape* (*n\_samples*,)) – Predicted mean.
- **sample\_weight** (*array-like*, *shape* (*n\_samples*,), *optional* (*default=1*)) – Sample weights.
- **dispersion** (*float*, *optional* (*default=None*)) – Dispersion parameter  $\phi$ . Estimated if None.

**Return type**

*float*

**property lower\_bound: float**

Return the lowest value of *y* allowed.

**property power: float**

Return the Tweedie power parameter.

**rowwise\_gradient\_hessian(link, coef, dispersion, X, y, sample\_weight, eta, mu, offset=None)**

Compute the gradient and negative Hessian of the log likelihood row-wise.

**Returns**

- *numpy.ndarray*, *shape* (*X.shape*[0],) – The gradient of the log likelihood, row-wise.
- *numpy.ndarray*, *shape* (*X.shape*[0],) – The negative Hessian of the log likelihood, row-wise.

**Parameters**

- **link** ([Link](#)) –
- **coef** (*ndarray*) –
- **X** (*MatrixBase* | *StandardizedMatrix*) –
- **y** (*ndarray*) –
- **sample\_weight** (*ndarray*) –
- **eta** (*ndarray*) –
- **mu** (*ndarray*) –
- **offset** (*ndarray* | *None*) –

**unit\_deviance**(*y*, *mu*)

Get the deviance of each observation.

**unit\_deviance\_derivative**(*y*, *mu*)

Compute the derivative of the unit deviance with respect to *mu*.

The derivative of the unit deviance is given by  $-2 \times (y - \mu) / v(\mu)$ , where  $v(\mu)$  is the unit variance.

**Parameters**

- **y** (*array-like*, *shape* (*n\_samples*,)) – Target values.
- **mu** (*array-like*, *shape* (*n\_samples*,)) – Predicted mean.

**Return type**

*array-like*, *shape* (*n\_samples*,)

**unit\_variance**(*mu*)

Compute the unit variance of a Tweedie distribution  $v(\mu) = \mu^{\text{power}}$ .

**Parameters**

- **mu** (*array-like*, *shape* (*n\_samples*,)) – Predicted mean.

**Return type**

*numpy.ndarray*, *shape* (*n\_samples*,)

**unit\_variance\_derivative**(*mu*)

Compute the derivative of the unit variance of a Tweedie distribution.

Equation:  $v(\mu) = p \times \mu^{(p-1)}$ .

**Parameters**

- **mu** (*array-like*, *shape* (*n\_samples*,)) – Predicted mean.

**Return type**

*numpy.ndarray*, *shape* (*n\_samples*,)

**variance**(*mu*, *dispersion=1*, *sample\_weight=1*)

Compute the variance function.

The variance of  $Y_i \sim \text{EDM}(\mu_i, \phi/s_i)$  is  $\text{var}(Y_i) = (\phi/s_i) * v(\mu_i)$ , with unit variance  $v(\mu)$  and weights  $s_i$ .

**Parameters**

- **mu** (*array-like*, *shape* (*n\_samples*,)) – Predicted mean.
- **dispersion** (*float*, *optional* (*default=1*)) – Dispersion parameter  $\phi$ .

- **sample\_weight** (*array-like, shape (n\_samples,)*, *optional (default=1)*)  
– Weights or exposure to which variance is inverse proportional.

**Return type**

*array-like, shape (n\_samples,)*

**variance\_derivative**(*mu, dispersion=1, sample\_weight=1*)

Compute the derivative of the variance with respect to  $\mu$ .

The derivative of the variance is equal to  $(\phi/s_i) * v'(\mu_i)$ , where  $v(\mu)$  is the unit variance and  $s_i$  are weights.

**Parameters**

- **mu** (*array-like, shape (n\_samples,)*) – Predicted mean.
- **dispersion** (*float, optional (default=1)*) – Dispersion parameter  $\phi$ .
- **sample\_weight** (*array-like, shape (n\_samples,)*, *optional (default=1)*)  
– Weights or exposure to which variance is inverse proportional.

**Return type**

*array-like, shape (n\_samples,)*

**class** glum.Link

Bases: object

Abstract base class for Link functions.

**abstract derivative**(*mu*)

Compute the derivative of the link  $g'(\mu)$ .

**Parameters**

**mu** (*array-like, shape (n\_samples,)*) – Usually the (predicted) mean.

**abstract inverse**(*lin\_pred*)

Compute the inverse link function  $h(\text{lin\_pred})$ .

Gives the inverse relationship between linear predictor,  $\text{lin\_pred} = X * w$ , and the mean,  $\mu = E(Y)$ , i.e.  $h(\text{lin\_pred}) = \mu$ .

**Parameters**

**lin\_pred** (*array-like, shape (n\_samples,)*) – Usually the (fitted) linear predictor.

**abstract inverse\_derivative**(*lin\_pred*)

Compute the derivative of the inverse link function  $h'(\text{lin\_pred})$ .

**Parameters**

**lin\_pred** (*array-like, shape (n\_samples,)*) – Usually the (fitted) linear predictor.

**abstract inverse\_derivative2**(*lin\_pred*)

Compute second derivative of the inverse link function  $h''(\text{lin\_pred})$ .

**Parameters**

**lin\_pred** (*array-like, shape (n\_samples,)*) – Usually the (fitted) linear predictor.

**abstract link**(*mu*)

Compute the link function  $g(\mu)$ .

The link function links the mean,  $\mu = E(Y)$ , to the linear predictor  $X * w$ , i.e.  $g(\mu)$  is equal to the linear predictor.

**Parameters**

**mu** (*array-like, shape (n\_samples,)*) – Usually the (predicted) mean.

**class** glum.LogLinkBases: [Link](#)The log link function `log(x)`.**derivative**(*mu*)

Get the derivative of the log link, one over mu.

**Parameters****mu** (*array-like*) –**Return type**

numpy.ndarray

**inverse**(*lin\_pred*)

Get the inverse of the log link, the exponential function.

See superclass documentation.

**Parameters****lin\_pred** (*array-like*) –**Return type**

numpy.ndarray

**inverse\_derivative**(*lin\_pred*)Compute the derivative of the inverse link function  $h'(lin\_pred)$ .**Parameters****lin\_pred** (*array-like*, *shape (n\_samples,)*) – Usually the (fitted) linear predictor.**inverse\_derivative2**(*lin\_pred*)Compute second derivative of the inverse link function  $h''(lin\_pred)$ .**Parameters****lin\_pred** (*array-like*, *shape (n\_samples,)*) – Usually the (fitted) linear predictor.**link**(*mu*)

Get the log of mu.

See superclass documentation.

**Parameters****mu** (*array-like*) –**Return type**

numpy.ndarray

**class** glum.LogitLinkBases: [Link](#)The logit link function `logit(x)`.**derivative**(*mu*)

Get the derivative of the logit link.

See superclass documentation.

**Parameters****mu** (*array-like*) –**Return type**

array-like

**inverse**(*lin\_pred*)

Get the inverse of the logit link.

See superclass documentation.

Note: since passing a very large value might result in an output of one, this function bounds the output to be between  $[50*\text{eps}, 1 - 50*\text{eps}]$ , where *eps* is floating point epsilon.

**Parameters****lin\_pred** (*array-like*) –**Return type**

array-like

**inverse\_derivative**(*lin\_pred*)

Compute the derivative of the inverse link function  $h'(\text{lin\_pred})$ .

**Parameters****lin\_pred** (*array, shape (n\_samples,)*) – Usually the (fitted) linear predictor.**inverse\_derivative2**(*lin\_pred*)

Compute second derivative of the inverse link function  $h''(\text{lin\_pred})$ .

**Parameters****lin\_pred** (*array, shape (n\_samples,)*) – Usually the (fitted) linear predictor.**link**(*mu*)

Get the logit function of *mu*.

See superclass documentation.

**Parameters****mu** (*array-like*) –**Return type**

numpy.ndarray

**class** glum.NegativeBinomialDistribution(*theta=1.0*)

Bases: [ExponentialDispersionModel](#)

A class for the Negative Binomial distribution.

A Negative Binomial distribution with mean  $\mu = E(Y)$  is uniquely defined by its mean-variance relationship  $\text{var}(Y) \propto \mu + \theta * \mu^2$ .

**Parameters****theta** (*float, optional (default=1.0)*) – The dispersion parameter from *unit\_variance*  $v(\mu) = \mu + \theta * \mu^2$ . For  $\theta \leq 0$ , no distribution exists.

## References

**For the log-likelihood and deviance:**

- M. L. Zwilling Negative Binomial Regression, The Mathematica Journal 2013. <https://www.mathematica-journal.com/2013/06/27/negative-binomial-regression/>

**deviance**(*y, mu, sample\_weight=None*)

Compute the deviance.

**Parameters**



- **y** (*array-like*, *shape* (*n\_samples*,)) – Target values.
- **mu** (*array-like*, *shape* (*n\_samples*,)) – Predicted mean.
- **sample\_weight** (*array-like*, *shape* (*n\_samples*,), *optional* (*default=1*)) – Sample weights.

**Return type**

float

**deviance\_derivative**(*y*, *mu*, *sample\_weight=1*)

Compute the derivative of the deviance with respect to mu.

**Parameters**

- **y** (*array-like*, *shape* (*n\_samples*,)) – Target values.
- **mu** (*array-like*, *shape* (*n\_samples*,)) – Predicted mean.
- **sample\_weight** (*array-like*, *shape* (*n\_samples*,) (*default=1*)) – Weights or exposure to which variance is inverse proportional.

**Return type**array-like, shape (*n\_samples*,)**dispersion**(*y*, *mu*, *sample\_weight=None*, *ddof=1*, *method='pearson'*)Estimate the dispersion parameter  $\phi$ .**Parameters**

- **y** (*array-like*, *shape* (*n\_samples*,)) – Target values.
- **mu** (*array-like*, *shape* (*n\_samples*,)) – Predicted mean.
- **sample\_weight** (*array-like*, *shape* (*n\_samples*,), *optional* (*default=None*)) – Weights or exposure to which variance is inversely proportional.
- **ddof** (*int*, *optional* (*default=1*)) – Degrees of freedom consumed by the model for mu.
- **{'pearson'}** (*method* =) – Whether to base the estimate on the Pearson residuals or the deviance.
- **'deviance'** – Whether to base the estimate on the Pearson residuals or the deviance.
- **(default='pearson')** (*optional*) – Whether to base the estimate on the Pearson residuals or the deviance.

**Return type**

float

**eta\_mu\_deviance**(*link*, *factor*, *cur\_eta*, *X\_dot\_d*, *y*, *sample\_weight*)

Compute eta, mu and the deviance.

Compute: \* the linear predictor, eta, as *cur\_eta* + *factor* \* *X\_dot\_d*; \* the link-function-transformed prediction, mu; \* the deviance.**Returns**

- *numpy.ndarray*, *shape* (*X.shape*[0],) – The linear predictor, eta.
- *numpy.ndarray*, *shape* (*X.shape*[0],) – The link-function-transformed prediction, mu.
- *float* – The deviance.

**Parameters**

- **link** ([Link](#)) –
- **factor** (*float*) –
- **cur\_eta** (*ndarray*) –
- **X\_dot\_d** (*ndarray*) –
- **y** (*ndarray*) –
- **sample\_weight** (*ndarray*) –

**in\_y\_range**(*x*)

Return True if *x* is in the valid range of the EDM.

**Parameters**

**x** (*array-like*, *shape* (*n\_samples*,)) – Target values.

**Return type**

*np.ndarray*

**log\_likelihood**(*y*, *mu*, *sample\_weight=None*, *dispersion=1*)

Compute the log likelihood.

**Parameters**

- **y** (*array-like*, *shape* (*n\_samples*,)) – Target values.
- **mu** (*array-like*, *shape* (*n\_samples*,)) – Predicted mean.
- **sample\_weight** (*array-like*, *shape* (*n\_samples*,), *optional* (*default=1*)) – Sample weights.
- **dispersion** (*float*, *optional* (*default=1.0*)) – Ignored.

**Return type**

*float*

**rowwise\_gradient\_hessian**(*link*, *coef*, *dispersion*, *X*, *y*, *sample\_weight*, *eta*, *mu*, *offset=None*)

Compute the gradient and negative Hessian of the log likelihood row-wise.

**Returns**

- *numpy.ndarray*, *shape* (*X.shape*[0],) – The gradient of the log likelihood, row-wise.
- *numpy.ndarray*, *shape* (*X.shape*[0],) – The negative Hessian of the log likelihood, row-wise.

**Parameters**

- **link** ([Link](#)) –
- **coef** (*ndarray*) –
- **X** (*MatrixBase* | *StandardizedMatrix*) –
- **y** (*ndarray*) –
- **sample\_weight** (*ndarray*) –
- **eta** (*ndarray*) –
- **mu** (*ndarray*) –
- **offset** (*ndarray* | *None*) –

**property theta: float**

Return the Negative Binomial theta parameter.

**unit\_deviance(y, mu)**

Get the deviance of each observation.

**Parameters**

- **y** (*ndarray*) –
- **mu** (*ndarray*) –

**Return type**

*ndarray*

**unit\_deviance\_derivative(y, mu)**

Compute the derivative of the unit deviance with respect to mu.

The derivative of the unit deviance is given by  $-2 \times (y - \mu)/v(\mu)$ , where  $v(\mu)$  is the unit variance.

**Parameters**

- **y** (*array-like, shape (n\_samples,)*) – Target values.
- **mu** (*array-like, shape (n\_samples,)*) – Predicted mean.

**Return type**

*array-like, shape (n\_samples,)*

**unit\_variance(mu)**

Compute the unit variance of a Negative Binomial distribution  $v(\mu) = \mu + \theta * \mu^2$ .

**Parameters**

- **mu** (*array-like, shape (n\_samples,)*) – Predicted mean.

**Return type**

*numpy.ndarray, shape (n\_samples,)*

**unit\_variance\_derivative(mu)**

Compute the derivative of the unit variance of a Negative Binomial distribution.

Equation:  $v(\mu) = 1 + 2 \times \theta \times \mu$ .

**Parameters**

- **mu** (*array-like, shape (n\_samples,)*) – Predicted mean.

**Return type**

*numpy.ndarray, shape (n\_samples,)*

**variance(mu, dispersion=1, sample\_weight=1)**

Compute the variance function.

The variance of  $Y_i \sim \text{EDM}(\mu_i, \phi/s_i)$  is  $\text{var}(Y_i) = (\phi/s_i) * v(\mu_i)$ , with unit variance  $v(\mu)$  and weights  $s_i$ .

**Parameters**

- **mu** (*array-like, shape (n\_samples,)*) – Predicted mean.
- **dispersion** (*float, optional (default=1)*) – Dispersion parameter  $\phi$ .
- **sample\_weight** (*array-like, shape (n\_samples,)*, *optional (default=1)*) – Weights or exposure to which variance is inverse proportional.

**Return type**

*array-like, shape (n\_samples,)*

**variance\_derivative**(*mu*, *dispersion=1*, *sample\_weight=1*)

Compute the derivative of the variance with respect to *mu*.

The derivative of the variance is equal to  $(\phi/s_i) * v'(\mu_i)$ , where  $v(\mu)$  is the unit variance and  $s_i$  are weights.

**Parameters**

- **mu** (*array-like*, *shape (n\_samples,)*) – Predicted mean.
- **dispersion** (*float*, *optional (default=1)*) – Dispersion parameter  $\phi$ .
- **sample\_weight** (*array-like*, *shape (n\_samples,)*, *optional (default=1)*) – Weights or exposure to which variance is inverse proportional.

**Return type**

*array-like*, *shape (n\_samples,)*

**class** `glum.NormalDistribution`

Bases: `TweedieDistribution`

Class for the Normal (a.k.a. Gaussian) distribution.

**deviance**(*y*, *mu*, *sample\_weight=None*)

Compute the deviance.

**Parameters**

- **y** (*array-like*, *shape (n\_samples,)*) – Target values.
- **mu** (*array-like*, *shape (n\_samples,)*) – Predicted mean.
- **sample\_weight** (*array-like*, *shape (n\_samples,)*, *optional (default=1)*) – Sample weights.

**Return type**

*float*

**deviance\_derivative**(*y*, *mu*, *sample\_weight=1*)

Compute the derivative of the deviance with respect to *mu*.

**Parameters**

- **y** (*array-like*, *shape (n\_samples,)*) – Target values.
- **mu** (*array-like*, *shape (n\_samples,)*) – Predicted mean.
- **sample\_weight** (*array-like*, *shape (n\_samples,)* (*default=1*)) – Weights or exposure to which variance is inverse proportional.

**Return type**

*array-like*, *shape (n\_samples,)*

**dispersion**(*y*, *mu*, *sample\_weight=None*, *ddof=1*, *method='pearson'*)

Estimate the dispersion parameter  $\phi$ .

**Parameters**

- **y** (*array-like*, *shape (n\_samples,)*) – Target values.
- **mu** (*array-like*, *shape (n\_samples,)*) – Predicted mean.
- **sample\_weight** (*array-like*, *shape (n\_samples,)*, *optional (default=None)*) – Weights or exposure to which variance is inversely proportional.
- **ddof** (*int*, *optional (default=1)*) – Degrees of freedom consumed by the model for *mu*.

- **'pearson'** (*method* =) – Whether to base the estimate on the Pearson residuals or the deviance.
- **'deviance'** – Whether to base the estimate on the Pearson residuals or the deviance.
- (**default='pearson'**) (*optional*) – Whether to base the estimate on the Pearson residuals or the deviance.

**Return type**

float

**eta\_mu\_deviance**(*link, factor, cur\_eta, X\_dot\_d, y, sample\_weight*)

Compute eta, mu and the deviance.

Compute: \* the linear predictor, eta, as `cur_eta + factor * X_dot_d`; \* the link-function-transformed prediction, mu; \* the deviance.**Returns**

- *numpy.ndarray, shape (X.shape[0],)* – The linear predictor, eta.
- *numpy.ndarray, shape (X.shape[0],)* – The link-function-transformed prediction, mu.
- *float* – The deviance.

**Parameters**

- **link** ([Link](#)) –
- **factor** (*float*) –
- **cur\_eta** (*ndarray*) –
- **X\_dot\_d** (*ndarray*) –
- **y** (*ndarray*) –
- **sample\_weight** (*ndarray*) –

**in\_y\_range**(*x*)Return True if *x* is in the valid range of the EDM.**Parameters****x** (*array-like, shape (n\_samples,)*) – Target values.**Return type**

np.ndarray

**property include\_lower\_bound: bool**Return whether `lower_bound` is allowed as a value of *y*.**log\_likelihood**(*y, mu, sample\_weight=None, dispersion=None*)

Compute the log likelihood.

For  $1 < \text{power} < 2$ , we use the series approximation by Dunn and Smyth (2005) to compute the normalization term.**Parameters**

- **y** (*array-like, shape (n\_samples,)*) – Target values.
- **mu** (*array-like, shape (n\_samples,)*) – Predicted mean.
- **sample\_weight** (*array-like, shape (n\_samples,)*, *optional (default=1)*) – Sample weights.

- **dispersion** (*float, optional (default=None)*) – Dispersion parameter  $\phi$ . Estimated if None.

**Return type**

float

**property lower\_bound: float**

Return the lowest value of y allowed.

**property power: float**

Return the Tweedie power parameter.

**rowwise\_gradient\_hessian**(*link, coef, dispersion, X, y, sample\_weight, eta, mu, offset=None*)

Compute the gradient and negative Hessian of the log likelihood row-wise.

**Returns**

- *numpy.ndarray, shape (X.shape[0],)* – The gradient of the log likelihood, row-wise.
- *numpy.ndarray, shape (X.shape[0],)* – The negative Hessian of the log likelihood, row-wise.

**Parameters**

- **link** (*Link*) –
- **coef** (*ndarray*) –
- **X** (*MatrixBase | StandardizedMatrix*) –
- **y** (*ndarray*) –
- **sample\_weight** (*ndarray*) –
- **eta** (*ndarray*) –
- **mu** (*ndarray*) –
- **offset** (*ndarray | None*) –

**unit\_deviance**(*y, mu*)

Get the deviance of each observation.

**unit\_deviance\_derivative**(*y, mu*)

Compute the derivative of the unit deviance with respect to mu.

The derivative of the unit deviance is given by  $-2 \times (y - \mu) / v(\mu)$ , where  $v(\mu)$  is the unit variance.

**Parameters**

- **y** (*array-like, shape (n\_samples,)*) – Target values.
- **mu** (*array-like, shape (n\_samples,)*) – Predicted mean.

**Return type**

array-like, shape (n\_samples,)

**unit\_variance**(*mu*)

Compute the unit variance of a Tweedie distribution  $v(\mu) = \mu^{\text{power}}$ .

**Parameters**

**mu** (*array-like, shape (n\_samples,)*) – Predicted mean.

**Return type**

numpy.ndarray, shape (n\_samples,)

**unit\_variance\_derivative(mu)**

Compute the derivative of the unit variance of a Tweedie distribution.

Equation:  $v(\mu) = p \times \mu^{(p-1)}$ .

**Parameters**

**mu** (array-like, shape (n\_samples,)) – Predicted mean.

**Return type**

numpy.ndarray, shape (n\_samples,)

**variance(mu, dispersion=1, sample\_weight=1)**

Compute the variance function.

The variance of  $Y_i \sim \text{EDM}(\mu_i, \phi/s_i)$  is  $\text{var}(Y_i) = (\phi/s_i) * v(\mu_i)$ , with unit variance  $v(\mu)$  and weights  $s_i$ .

**Parameters**

- **mu** (array-like, shape (n\_samples,)) – Predicted mean.
- **dispersion** (float, optional (default=1)) – Dispersion parameter  $\phi$ .
- **sample\_weight** (array-like, shape (n\_samples,)), optional (default=1) – Weights or exposure to which variance is inverse proportional.

**Return type**

array-like, shape (n\_samples,)

**variance\_derivative(mu, dispersion=1, sample\_weight=1)**

Compute the derivative of the variance with respect to mu.

The derivative of the variance is equal to  $(\phi/s_i) * v'(\mu_i)$ , where  $v(\mu)$  is the unit variance and  $s_i$  are weights.

**Parameters**

- **mu** (array-like, shape (n\_samples,)) – Predicted mean.
- **dispersion** (float, optional (default=1)) – Dispersion parameter  $\phi$ .
- **sample\_weight** (array-like, shape (n\_samples,)), optional (default=1) – Weights or exposure to which variance is inverse proportional.

**Return type**

array-like, shape (n\_samples,)

**class glum.PoissonDistribution**

Bases: [TweedieDistribution](#)

Class for the scaled Poisson distribution.

**deviance(y, mu, sample\_weight=None)**

Compute the deviance.

**Parameters**

- **y** (array-like, shape (n\_samples,)) – Target values.
- **mu** (array-like, shape (n\_samples,)) – Predicted mean.
- **sample\_weight** (array-like, shape (n\_samples,)), optional (default=1) – Sample weights.

**Return type**

float

**deviance\_derivative**(*y, mu, sample\_weight=1*)

Compute the derivative of the deviance with respect to  $\mu$ .

**Parameters**

- **y** (*array-like, shape (n\_samples,)*) – Target values.
- **mu** (*array-like, shape (n\_samples,)*) – Predicted mean.
- **sample\_weight** (*array-like, shape (n\_samples,) (default=1)*) – Weights or exposure to which variance is inverse proportional.

**Return type**

*array-like, shape (n\_samples,)*

**dispersion**(*y, mu, sample\_weight=None, ddof=1, method='pearson'*)

Estimate the dispersion parameter  $\phi$ .

**Parameters**

- **y** (*array-like, shape (n\_samples,)*) – Target values.
- **mu** (*array-like, shape (n\_samples,)*) – Predicted mean.
- **sample\_weight** (*array-like, shape (n\_samples,)*, *optional (default=None)*) – Weights or exposure to which variance is inversely proportional.
- **ddof** (*int, optional (default=1)*) – Degrees of freedom consumed by the model for  $\mu$ .
- **{'pearson'}** (*method =*) – Whether to base the estimate on the Pearson residuals or the deviance.
- **'deviance'** – Whether to base the estimate on the Pearson residuals or the deviance.
- **(default='pearson')** (*optional*) – Whether to base the estimate on the Pearson residuals or the deviance.

**Return type**

*float*

**eta\_mu\_deviance**(*link, factor, cur\_eta, X\_dot\_d, y, sample\_weight*)

Compute  $\eta$ ,  $\mu$  and the deviance.

Compute: \* the linear predictor,  $\eta$ , as  $\text{cur\_eta} + \text{factor} * \text{X\_dot\_d}$ ; \* the link-function-transformed prediction,  $\mu$ ; \* the deviance.

**Returns**

- *numpy.ndarray, shape (X.shape[0],)* – The linear predictor,  $\eta$ .
- *numpy.ndarray, shape (X.shape[0],)* – The link-function-transformed prediction,  $\mu$ .
- *float* – The deviance.

**Parameters**

- **link** (*Link*) –
- **factor** (*float*) –
- **cur\_eta** (*ndarray*) –
- **X\_dot\_d** (*ndarray*) –
- **y** (*ndarray*) –



- **sample\_weight** (*ndarray*) –

**in\_y\_range**(*x*)

Return True if *x* is in the valid range of the EDM.

**Parameters**

**x** (*array-like*, *shape* (*n\_samples*,)) – Target values.

**Return type**

*np.ndarray*

**property include\_lower\_bound: bool**

Return whether *lower\_bound* is allowed as a value of *y*.

**log\_likelihood**(*y*, *mu*, *sample\_weight=None*, *dispersion=None*)

Compute the log likelihood.

For  $1 < \text{power} < 2$ , we use the series approximation by Dunn and Smyth (2005) to compute the normalization term.

**Parameters**

- **y** (*array-like*, *shape* (*n\_samples*,)) – Target values.
- **mu** (*array-like*, *shape* (*n\_samples*,)) – Predicted mean.
- **sample\_weight** (*array-like*, *shape* (*n\_samples*,), *optional* (*default=1*)) – Sample weights.
- **dispersion** (*float*, *optional* (*default=None*)) – Dispersion parameter  $\phi$ . Estimated if *None*.

**Return type**

*float*

**property lower\_bound: float**

Return the lowest value of *y* allowed.

**property power: float**

Return the Tweedie power parameter.

**rowwise\_gradient\_hessian**(*link*, *coef*, *dispersion*, *X*, *y*, *sample\_weight*, *eta*, *mu*, *offset=None*)

Compute the gradient and negative Hessian of the log likelihood row-wise.

**Returns**

- *numpy.ndarray*, *shape* (*X.shape*[0],) – The gradient of the log likelihood, row-wise.
- *numpy.ndarray*, *shape* (*X.shape*[0],) – The negative Hessian of the log likelihood, row-wise.

**Parameters**

- **link** ([Link](#)) –
- **coef** (*ndarray*) –
- **X** (*MatrixBase* | *StandardizedMatrix*) –
- **y** (*ndarray*) –
- **sample\_weight** (*ndarray*) –
- **eta** (*ndarray*) –
- **mu** (*ndarray*) –

- **offset** (*ndarray* | *None*) –

**unit\_deviance**(*y*, *mu*)

Get the deviance of each observation.

**unit\_deviance\_derivative**(*y*, *mu*)

Compute the derivative of the unit deviance with respect to *mu*.

The derivative of the unit deviance is given by  $-2 \times (y - \mu)/v(\mu)$ , where  $v(\mu)$  is the unit variance.

**Parameters**

- **y** (*array-like*, *shape* (*n\_samples*,)) – Target values.
- **mu** (*array-like*, *shape* (*n\_samples*,)) – Predicted mean.

**Return type**

*array-like*, *shape* (*n\_samples*,)

**unit\_variance**(*mu*)

Compute the unit variance of a Tweedie distribution  $v(\mu) = \mu^{\text{power}}$ .

**Parameters**

**mu** (*array-like*, *shape* (*n\_samples*,)) – Predicted mean.

**Return type**

*numpy.ndarray*, *shape* (*n\_samples*,)

**unit\_variance\_derivative**(*mu*)

Compute the derivative of the unit variance of a Tweedie distribution.

Equation:  $v(\mu) = p \times \mu^{(p-1)}$ .

**Parameters**

**mu** (*array-like*, *shape* (*n\_samples*,)) – Predicted mean.

**Return type**

*numpy.ndarray*, *shape* (*n\_samples*,)

**variance**(*mu*, *dispersion=1*, *sample\_weight=1*)

Compute the variance function.

The variance of  $Y_i \sim \text{EDM}(\mu_i, \phi/s_i)$  is  $\text{var}(Y_i) = (\phi/s_i) * v(\mu_i)$ , with unit variance  $v(\mu)$  and weights  $s_i$ .

**Parameters**

- **mu** (*array-like*, *shape* (*n\_samples*,)) – Predicted mean.
- **dispersion** (*float*, *optional* (*default=1*)) – Dispersion parameter  $\phi$ .
- **sample\_weight** (*array-like*, *shape* (*n\_samples*,), *optional* (*default=1*))  
– Weights or exposure to which variance is inverse proportional.

**Return type**

*array-like*, *shape* (*n\_samples*,)

**variance\_derivative**(*mu*, *dispersion=1*, *sample\_weight=1*)

Compute the derivative of the variance with respect to *mu*.

The derivative of the variance is equal to  $(\phi/s_i) * v'(\mu_i)$ , where  $v(\mu)$  is the unit variance and  $s_i$  are weights.

**Parameters**

- **mu** (*array-like*, *shape* (*n\_samples*,)) – Predicted mean.
- **dispersion** (*float*, *optional* (*default=1*)) – Dispersion parameter  $\phi$ .

- **sample\_weight** (array-like, shape (n\_samples,)), optional (default=1)  
– Weights or exposure to which variance is inverse proportional.

**Return type**

array-like, shape (n\_samples,)

**class** glum.TweedieDistribution(power=0)Bases: *ExponentialDispersionModel*

A class for the Tweedie distribution.

A Tweedie distribution with mean  $\mu = E(Y)$  is uniquely defined by its mean-variance relationship  $\text{var}(Y) \propto \mu^{\text{power}}$ .

Special cases are:

Power	Distribution
0	Normal
1	Poisson
(1, 2)	Compound Poisson
2	Gamma
3	Inverse Gaussian

**Parameters**

**power** (float, optional (default=0)) – The variance power of the *unit\_variance*  $v(\mu) = \mu^{\text{power}}$ . For  $0 < \text{power} < 1$ , no distribution exists.

**deviance**(y, mu, sample\_weight=None)

Compute the deviance.

**Parameters**

- **y** (array-like, shape (n\_samples,)) – Target values.
- **mu** (array-like, shape (n\_samples,)) – Predicted mean.
- **sample\_weight** (array-like, shape (n\_samples,)), optional (default=1)  
– Sample weights.

**Return type**

float

**deviance\_derivative**(y, mu, sample\_weight=1)

Compute the derivative of the deviance with respect to mu.

**Parameters**

- **y** (array-like, shape (n\_samples,)) – Target values.
- **mu** (array-like, shape (n\_samples,)) – Predicted mean.
- **sample\_weight** (array-like, shape (n\_samples,)) (default=1) – Weights or exposure to which variance is inverse proportional.

**Return type**

array-like, shape (n\_samples,)

**dispersion**(y, mu, sample\_weight=None, ddof=1, method='pearson')Estimate the dispersion parameter  $\phi$ .**Parameters**

- **y** (*array-like*, *shape* (*n\_samples*,)) – Target values.
- **mu** (*array-like*, *shape* (*n\_samples*,)) – Predicted mean.
- **sample\_weight** (*array-like*, *shape* (*n\_samples*,), *optional* (*default=None*)) – Weights or exposure to which variance is inversely proportional.
- **ddof** (*int*, *optional* (*default=1*)) – Degrees of freedom consumed by the model for mu.
- **{'pearson'}** (*method* =) – Whether to base the estimate on the Pearson residuals or the deviance.
- **'deviance'** – Whether to base the estimate on the Pearson residuals or the deviance.
- **(default='pearson')** (*optional*) – Whether to base the estimate on the Pearson residuals or the deviance.

**Return type**

float

**eta\_mu\_deviance**(*link*, *factor*, *cur\_eta*, *X\_dot\_d*, *y*, *sample\_weight*)

Compute eta, mu and the deviance.

Compute: \* the linear predictor, eta, as *cur\_eta* + *factor* \* *X\_dot\_d*; \* the link-function-transformed prediction, mu; \* the deviance.**Returns**

- *numpy.ndarray*, *shape* (*X.shape[0]*,) – The linear predictor, eta.
- *numpy.ndarray*, *shape* (*X.shape[0]*,) – The link-function-transformed prediction, mu.
- *float* – The deviance.

**Parameters**

- **link** (*Link*) –
- **factor** (*float*) –
- **cur\_eta** (*ndarray*) –
- **X\_dot\_d** (*ndarray*) –
- **y** (*ndarray*) –
- **sample\_weight** (*ndarray*) –

**in\_y\_range**(*x*)Return True if *x* is in the valid range of the EDM.**Parameters****x** (*array-like*, *shape* (*n\_samples*,)) – Target values.**Return type**

np.ndarray

**property include\_lower\_bound: bool**Return whether *lower\_bound* is allowed as a value of *y*.**log\_likelihood**(*y*, *mu*, *sample\_weight=None*, *dispersion=None*)

Compute the log likelihood.

For  $1 < \text{power} < 2$ , we use the series approximation by Dunn and Smyth (2005) to compute the normalization term.

**Parameters**

- **y** (*array-like, shape (n\_samples,)*) – Target values.
- **mu** (*array-like, shape (n\_samples,)*) – Predicted mean.
- **sample\_weight** (*array-like, shape (n\_samples,)*, *optional (default=1)*) – Sample weights.
- **dispersion** (*float, optional (default=None)*) – Dispersion parameter  $\phi$ . Estimated if None.

**Return type**

float

**property lower\_bound: float**

Return the lowest value of y allowed.

**property power: float**

Return the Tweedie power parameter.

**rowwise\_gradient\_hessian**(*link, coef, dispersion, X, y, sample\_weight, eta, mu, offset=None*)

Compute the gradient and negative Hessian of the log likelihood row-wise.

**Returns**

- *numpy.ndarray, shape (X.shape[0],)* – The gradient of the log likelihood, row-wise.
- *numpy.ndarray, shape (X.shape[0],)* – The negative Hessian of the log likelihood, row-wise.

**Parameters**

- **link** (*Link*) –
- **coef** (*ndarray*) –
- **X** (*MatrixBase | StandardizedMatrix*) –
- **y** (*ndarray*) –
- **sample\_weight** (*ndarray*) –
- **eta** (*ndarray*) –
- **mu** (*ndarray*) –
- **offset** (*ndarray | None*) –

**unit\_deviance**(*y, mu*)

Get the deviance of each observation.

**unit\_deviance\_derivative**(*y, mu*)

Compute the derivative of the unit deviance with respect to mu.

The derivative of the unit deviance is given by  $-2 \times (y - \mu)/v(\mu)$ , where  $v(\mu)$  is the unit variance.**Parameters**

- **y** (*array-like, shape (n\_samples,)*) – Target values.
- **mu** (*array-like, shape (n\_samples,)*) – Predicted mean.

**Return type**

array-like, shape (n\_samples,)

**unit\_variance(mu)**

Compute the unit variance of a Tweedie distribution  $v(\mu) = \mu^{\text{power}}$ .

**Parameters**

**mu** (array-like, shape (n\_samples,)) – Predicted mean.

**Return type**

numpy.ndarray, shape (n\_samples,)

**unit\_variance\_derivative(mu)**

Compute the derivative of the unit variance of a Tweedie distribution.

Equation:  $v(\mu) = p \times \mu^{(p-1)}$ .

**Parameters**

**mu** (array-like, shape (n\_samples,)) – Predicted mean.

**Return type**

numpy.ndarray, shape (n\_samples,)

**variance(mu, dispersion=1, sample\_weight=1)**

Compute the variance function.

The variance of  $Y_i \sim \text{EDM}(\mu_i, \phi/s_i)$  is  $\text{var}(Y_i) = (\phi/s_i) * v(\mu_i)$ , with unit variance  $v(\mu)$  and weights  $s_i$ .

**Parameters**

- **mu** (array-like, shape (n\_samples,)) – Predicted mean.
- **dispersion** (float, optional (default=1)) – Dispersion parameter  $\phi$ .
- **sample\_weight** (array-like, shape (n\_samples,)), optional (default=1) – Weights or exposure to which variance is inverse proportional.

**Return type**

array-like, shape (n\_samples,)

**variance\_derivative(mu, dispersion=1, sample\_weight=1)**

Compute the derivative of the variance with respect to mu.

The derivative of the variance is equal to  $(\phi/s_i) * v'(\mu_i)$ , where  $v(\mu)$  is the unit variance and  $s_i$  are weights.

**Parameters**

- **mu** (array-like, shape (n\_samples,)) – Predicted mean.
- **dispersion** (float, optional (default=1)) – Dispersion parameter  $\phi$ .
- **sample\_weight** (array-like, shape (n\_samples,)), optional (default=1) – Weights or exposure to which variance is inverse proportional.

**Return type**

array-like, shape (n\_samples,)

**class glum.TweedieLink(p)**

Bases: [Link](#)

The Tweedie link function  $x^{(1-p)}$  if  $p \neq 1$  and  $\log(x)$  if  $p=1$ .

**derivative(mu)**

Get the derivative of the Tweedie link.

See superclass documentation.

**Parameters****mu** (*array-like*) –**inverse**(\*\**kwargs*)Compute the inverse link function  $h(\text{lin\_pred})$ .

Gives the inverse relationship between linear predictor,  $\text{lin\_pred} = X * w$ , and the mean,  $\mu = E(Y)$ , i.e.  $h(\text{lin\_pred}) = \mu$ .

**Parameters****lin\_pred** (*array-like*, *shape* (*n\_samples*,)) – Usually the (fitted) linear predictor.**inverse\_derivative**(\*\**kwargs*)Compute the derivative of the inverse link function  $h'(\text{lin\_pred})$ .**Parameters****lin\_pred** (*array-like*, *shape* (*n\_samples*,)) – Usually the (fitted) linear predictor.**inverse\_derivative2**(\*\**kwargs*)Compute second derivative of the inverse link function  $h''(\text{lin\_pred})$ .**Parameters****lin\_pred** (*array-like*, *shape* (*n\_samples*,)) – Usually the (fitted) linear predictor.**link**(*mu*)

Get the Tweedie canonical link.

See superclass documentation.

**Parameters****mu** (*array-like*) –**glum.get\_link**(*link*, *family*)

For the Tweedie distribution, this code follows actuarial best practices regarding link functions. Note that these links are sometimes not canonical:

- identity for normal ( $p=0$ );
- no convention for  $p < 0$ , so let's leave it as identity;
- log otherwise.

**Parameters**

- **link** (*str* / *Link*) –
- **family** (*ExponentialDispersionModel*) –

**Return type***Link*





## CHANGELOG

### 9.1 2.7.0 - 2024-02-19

**Bug fix:**

- Added cython compiler directive `legacy_implicit_noexcept = True` to fix performance regression with cython 3.

**Other changes:**

- Require Python  $\geq 3.9$  in line with [NEP 29](#)
- Build and test with Python 3.12 in CI.
- Added line search stopping criterion for tiny loss improvements based on gradient information.
- Added warnings about breaking changes in future versions.

### 9.2 2.6.0 - 2023-09-05

**New features:**

- Added the complementary log-log (`cloglog`) link function.
- Added the option to store the covariance matrix after estimating it. In this case, the covariance matrix does not have to be recomputed when calling inference methods.
- Added methods for performing Wald tests based on a restriction matrix, feature names or term names.
- Added a method for creating a coefficient table with confidence intervals and p-values.

**Bug fix:**

- Fixed `covariance_matrix()` mutating feature names when called with a data frame. See [here](#).

**Other changes:**

- When computing the covariance matrix, check whether the design matrix is ill-conditioned for all types of input. Furthermore, do it in a more efficient way.
- Pin `tabmat < 4.0.0` (the new release will bring breaking changes).

## 9.3 2.5.2 - 2023-06-02

### Bug fix

- Fix the `glm_benchmarks_analyze` command line tool. See [here](#).
- Fixed a bug in *GeneralizedLinearRegressor* when fit on a data set with a constant column and `warm_start=True`. See [here](#).

### Other changes:

- Remove dev dependency on `dask_ml`.
- We now pin `llvm-openmp=11` when creating the wheel for macOS in line with what scikit-learn does.

## 9.4 2.5.1 - 2023-05-19

### Bug fix:

- We fixed a bug in the computation of `log_likelihood()`. Previously, this method just returned `None`.

## 9.5 2.5.0 - 2023-04-28

### New feature:

- Added Negative Binomial distribution by setting the 'family' parameter of *GeneralizedLinearRegressor* and *GeneralizedLinearRegressorCV* to 'negative.binomial'.

## 9.6 2.4.1 - 2023-03-14

### Bug fixes:

- Fixed an issue with `_score_matrix()` which failed when called with a tabmat matrix input.

### Other changes:

- Removes unused scikit-learn cython imports.

## 9.7 2.4.0 - 2023-01-31

### Other changes:

- *LogitLink* has been made public.
- Apple Silicon wheels are now uploaded to PyPI.

## 9.8 2.3.0 - 2023-01-06

### Bug fixes:

- A data frame with dense and sparse columns was transformed to a dense matrix instead of a split matrix by `_set_up_and_check_fit_args()`. Fixed by calling `tabmat.from_pandas` on any data frame.

### New features:

- The following classes and functions have been made public: *BinomialDistribution*, *ExponentialDispersionModel*, *GammaDistribution*, *GeneralizedHyperbolicSecant*, *InverseGaussianDistribution*, *NormalDistribution*, *PoissonDistribution*, *IdentityLink*, *Link*, *LogLink*, *TweedieLink*, `get_family()` and `get_link()`.
- The distribution and link classes now feature a more lenient equality check instead of the default identity check, so that, e.g., `TweedieDistribution(1) == TweedieDistribution(1)` now returns `True`.

## 9.9 2.2.1 - 2022-11-25

### Other changes:

- Fixing pypi upload issue. Version 2.2.0 will not be available through the standard distribution channels.

## 9.10 2.2.0 - 2022-11-25

### New features:

- Add an argument to *GeneralizedLinearRegressorBase* to drop the first category in a Categorical column using [implementation in tabmat](<https://github.com/Quantco/tabmat/pull/168>)
- One may now request the Tweedie loss by setting the 'family' parameter of *GeneralizedLinearRegressor* and *GeneralizedLinearRegressorCV* to 'tweedie'.

### Bug fixes:

- Setting bounds for constant columns was not working (bounds were internally modified to 0). A similar issue was preventing inequalities from working with constant columns. This is now fixed.

### Other changes:

- No more builds for 32-bit systems with python  $\geq 3.8$ . This is due to scipy not supporting it anymore.

## 9.11 2.1.2 - 2022-07-01

### Other changes:

- Next attempt to build wheel for PyPI without `--march=native`.

## 9.12 2.1.1 - 2022-07-01

### Other changes:

- We are now building the wheel for PyPI without `--march=native` to make it more portable across architectures.

## 9.13 2.1.0 - 2022-06-27

### New features:

- Added `aic()`, `aicc()` and `bic()` attributes to the *GeneralizedLinearRegressor*. These attributes provide the information criteria based on the training data and the effective degrees of freedom of the maximum likelihood estimate for the model's parameters.
- `std_errors()` and `covariance_matrix()` of *GeneralizedLinearRegressor* now accept data frames with categorical data.

### Bug fixes:

- The `score()` method of *GeneralizedLinearRegressor* and *GeneralizedLinearRegressorCV* now accepts offsets.
- Fixed the calculation of the information matrix for the Binomial distribution with logit link, which affected nonrobust standard errors.

### Other:

- The CI now runs daily unit tests against the nightly builds of numpy, pandas and scikit-learn.
- The minimally required version of tabmat is now 3.1.0.

## 9.14 2.0.3 - 2021-11-05

### Other:

- We are now specifying the run time dependencies in `setup.py`, so that missing dependencies are automatically installed from PyPI when installing `glum` via `pip`.

## 9.15 2.0.2 - 2021-11-03

### Bug fix:

- Fixed the sign of the log likelihood of the Gaussian distribution (not used for fitting coefficients).
- Fixed the wide benchmarks which had duplicated columns (categorical and numerical).

### Other:

- The CI now builds the wheels and upload to pypi with every new release.
- Renamed functions checking for `qc.matrix` compliance to refer to `tabmat`.

## 9.16 2.0.1 - 2021-10-11

### Bug fix:

- Fixed pyproject.toml. We now support installing through pip and pep517.

## 9.17 2.0.0 - 2021-10-08

### Breaking changes:

- Renamed the package to glum!! Hurray! Celebration.
- *GeneralizedLinearRegressor* and *GeneralizedLinearRegressorCV* lose the `fit_dispersion` parameter. Please use the `dispersion()` method of the appropriate family instance instead.
- All functions now use `sample_weight` as a keyword instead of `weights`, in line with scikit-learn.
- All functions now use `dispersion` as a keyword instead of `phi`.
- Several methods *GeneralizedLinearRegressor* and *GeneralizedLinearRegressorCV* that should have been private have had an underscore prefixed on their names: `tear_down_from_fit()`, `_set_up_for_fit()`, `_set_up_and_check_fit_args()`, `_get_start_coef()`, `_solve()` and `_solve_regularization_path()`.
- *glum.GeneralizedLinearRegressor.report\_diagnostics()* and *glum.GeneralizedLinearRegressor.get\_formatted\_diagnostics()* are now public.

### New features:

- P1 and P2 now accepts 1d array with the same number of elements as the unexpanded design matrix. In this case, the penalty associated with a categorical feature will be expanded to as many elements as there are levels, all with the same value.
- *ExponentialDispersionModel* gains a `dispersion()` method.
- *BinomialDistribution* and *TweedieDistribution* gain a `log_likelihood()` method.
- The `fit()` method of *GeneralizedLinearRegressor* and *GeneralizedLinearRegressorCV* now saves the column types of pandas data frames.
- *GeneralizedLinearRegressor* and *GeneralizedLinearRegressorCV* gain two properties: `family_instance` and `link_instance`.
- `std_errors()` and `covariance_matrix()` have been added and support non-robust, robust (HC-1), and clustered covariance matrices.
- *GeneralizedLinearRegressor* and *GeneralizedLinearRegressorCV* now accept `family='gaussian'` as an alternative to `family='normal'`.

### Bug fix:

- The `score()` method of *GeneralizedLinearRegressor* and *GeneralizedLinearRegressorCV* now accepts data frames.
- Upgraded the code to use tabmat 3.0.0.

### Other:

- A major overhaul of the documentation. Everything is better!
- The methods of the link classes will now return scalars when given scalar inputs. Under certain circumstances, they'd return zero-dimensional arrays.

- There is a new benchmark available `glm_benchmarks_run` based on the Boston housing dataset. See [here](#).
- `glm_benchmarks_analyze` now includes `offset` in the index. See [here](#).
- `glmnet_python` was removed from the benchmarks suite.
- The innermost coordinate descent was optimized. This speeds up coordinate descent dominated problems like LASSO by about 1.5-2x. See [here](#).

## 9.18 1.5.1 - 2021-07-22

### Bug fix:

- Have the `linear_predictor()` and `predict()` methods of *GeneralizedLinearRegressor* and *GeneralizedLinearRegressorCV* honor the `offset` when `alpha` is `None`.

## 9.19 1.5.0 - 2021-07-15

### New features:

- The `linear_predictor()` and `predict()` methods of *GeneralizedLinearRegressor* and *GeneralizedLinearRegressorCV* gain an `alpha` parameter (in complement to `alpha_index`). Moreover, they are now able to predict for multiple penalties.

### Other:

- Methods of *Link* now consistently return NumPy arrays, whereas they used to preserve pandas series in special cases.
- Don't list `sparse_dot_mkl` as a runtime requirement from the conda recipe.
- The minimal numpy pin should be dependent on the numpy version in host and not fixed to 1.16.

## 9.20 1.4.3 - 2021-06-25

### Bug fix:

- `copy_X = False` will now raise a value error when `X` has dtype `int32` or `int64`. Previously, it would only raise for dtype `int64`.

## 9.21 1.4.2 - 2021-06-15

### Tutorials and documentation improvements:

- Adding tutorials to the documentation.
- Additional documentation improvements.

### Bug fix:

- Verbose progress bar now working again.

### Other:

- Small improvement in documentation for the `alpha_index` argument to *predict()*.

- Pinned pre-commit hooks versions.

## 9.22 1.4.1 - 2021-05-01

We now have Windows builds!

## 9.23 1.4.0 - 2021-04-13

### Deprecations:

- Fusing the `alpha` and `alphas` arguments for *GeneralizedLinearRegressor*. `alpha` now also accepts array like inputs. `alphas` is now deprecated but can still be used for backward compatibility. The `alphas` argument will be removed with the next major version.

### Bug fix:

- We removed entry points to functions in `glum_benchmarks` from the conda package.

## 9.24 1.3.1 - 2021-04-12

### Bug fix:

- `glum._distribution.unit_variance_derivative()` is evaluating a proper numexpr expression again (regression in 1.3.0).

## 9.25 1.3.0 - 2021-04-12

### New features:

- We added a new solver based on `scipy.optimize.minimize(method='trust-constr')`.
- We added support for linear inequality constraints of type `A_ineq.dot(coef_) <= b_ineq`.

## 9.26 1.2.0 - 2021-02-04

We removed `glum_benchmarks` from the conda package.

## 9.27 1.1.1 - 2021-01-11

Maintenance release to get a fresh build for OSX.

## 9.28 1.1.0 - 2020-11-23

### New feature:

- Direct support for pandas categorical types in `fit` and `predict`. These will be converted into a `CategoricalMatrix`.

## 9.29 1.0.1 - 2020-11-12

This is a maintenance release to be compatible with `tabmat`  $\geq 1.0.0$ .

## 9.30 1.0.0 - 2020-11-11

### Other:

- Renamed `alpha_level` attribute of *GeneralizedLinearRegressor* and *GeneralizedLinearRegressorCV* to `alpha_index`.
- Clarified behavior of `scale_predictors`.

## 9.31 0.0.15 - 2020-11-11

### Other:

- Pin `tabmat`  $< 1.0.0$  as we are expecting a breaking change with version 1.0.0.

## 9.32 0.0.14 - 2020-08-06

### New features:

- Add Tweedie Link.
- Allow infinite bounds.

### Bug fixes:

- Unstandardize regularization path.
- No copying in `predict`.

### Other:

- Various memory and performance improvements.
- Update pre-commit hooks.



### **9.33 0.0.13 - 2020-07-23**

See git history.

### **9.34 0.0.12 - 2020-07-07**

See git history.

### **9.35 0.0.11 - 2020-07-02**

See git history.

### **9.36 0.0.10 - 2020-06-30**

See git history.

### **9.37 0.0.9 - 2020-06-26**

See git history.

### **9.38 0.0.8 - 2020-06-24**

See git history.

### **9.39 0.0.7 - 2020-06-17**

See git history.

### **9.40 0.0.6 - 2020-06-16**

See git history.

## **9.41 0.0.5 - 2020-06-10**

See git history.

## **9.42 0.0.4 - 2020-06-08**

See git history.

## **9.43 0.0.3 - 2020-06-08**

See git history.

genindex

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